STATISTICAL ANALYSIS OF HIT/MISS DATA
(PREPRINT)

Jeremy S. Knopp
Materials State Awareness & Supportability Branch
Structural Materials Division

Li Zeng
University of Texas at Arlington

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Jeremy S. Knopp (RXCA)

Li Zeng (University of Texas at Arlington)

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**14. Abstract**

There has been renewed interest in the analysis of Hit/Miss or Bernoulli data in the context of Nondestructive Evaluation (NDE). Many contributions have been made with a focus on confidence bound estimation on this type of data. Some concern regarding the proper calculation and use of a90 and a90/95 estimates has been raised. In particular, the behavior of a Probability of Detection (POD) curve for large flaw sizes above stated a90 estimates has been of concern. This paper will give a brief historical overview of the analysis of hit/miss data, and propose a remedy regarding recent concerns. The end result will be a procedure that helps avoid pitfalls in the analysis of hit/miss data.

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Statistical Analysis of Hit/Miss Data

Jeremy S. Knopp¹ and Li Zeng²

¹Air Force Research Laboratory (AFRL/RXLP), Wright-Patterson AFB, OH 45433, USA
²University of Texas at Arlington, Arlington, TX, 76019, USA

ABSTRACT

There has been renewed interest in the analysis of Hit/Miss or Bernoulli data in the context of Nondestructive Evaluation (NDE). Many contributions have been made with a focus on confidence bound estimation on this type of data. Some concern regarding the proper calculation and use of a₉₀ and a₉₀/₉₅ estimates has been raised. In particular, the behavior of a Probability of Detection (POD) curve for large flaw sizes above stated a₉₀ estimates has been of concern. This paper will give a brief historical overview of the analysis of hit/miss data, and propose a remedy regarding recent concerns. The end result will be a procedure that helps avoid pitfalls in the analysis of hit/miss data.

KEYWORDS

Probability of Detection, hit/miss, Bernoulli data, Bayesian methods

INTRODUCTION

Studies of the reliability of Nondestructive Inspections (NDI) began with the United States Air Force (USAF) in the 1960’s (Packman et al., 1968) and continued during the 1970’s with major efforts conducted by the National Aeronautics and Space Administration (NASA) (Rummel et al., 1974) and the USAF (Lewis et al., 1978). Initial efforts to explain the data from these inspection studies used
binomial statistics (Rummel, 1982). In the 1980’s, Berens and Hovey developed parametric methods to analyze both hit/miss and signal response data (Berens et al, 1981) and (Berens, 1989). This provided the foundation for guidance for POD studies that has been adopted by the Department of Defense (DoD), and has been published as a 2nd edition of a handbook (MIL-HDBK-1823A, 2009). Other agencies and industries have also made use of this guidance (Gandossi et al., 2010) and (Drury et al., 2006). It should also be noted that efforts in medical statistics are quite similar to the methods currently used for NDE (Collett, 2002). In medical statistics, the term effective dose is equivalent to $a_{50}$, and lethal dose is equivalent to $a_{90}$.

**CURRENT METHODS OF ANALYSIS FOR HIT/MISS DATA**

The current accepted methods used to analyze hit/miss data are similar to the ones published in an overview paper on the topic (Berens, 1989). In that work, the confidence bounds were global, that is they were calculated to apply to the entire POD curve. Later it was decided that this approach was overly conservative, and it was sufficient to apply confidence bounds for each flaw size locally (Berens, 2000). The likelihood ratio method was also introduced which provided more accurate confidence bounds for hit/miss data (Annis et al., 2007) and (Harding et al., 2003). These two developments were incorporated in the second edition of MIL-HDBK-1823. Recently, concerns have been raised about the current approach, and binomial methods have once again been proposed to mitigate concerns about the behavior of POD for flaw sizes greater than the established $a_{90}$ estimates generated by parametric approaches (Generazio, 2011). Further work on nonparametric methods has been done by Spencer with the assumption that the POD is a continuous non-decreasing function of flaw size (Spencer, 2011).

In the authors’ opinion, there is a valid concern raised in (Generazio, 2011), and that is that one could analyze hit/miss data that doesn’t meet the requirements of MIL-HDBK-1823A and still generate an estimate for $a_{90/95}$ that could be misused. Another concern in (Generazio, 2011) is the assumption of
a monotonically increasing POD function. Proper 3-point calibration (Rummel, 2005) should mitigate, at least in part, concerns about monotonicity.

A REMEDY FOR CONCERNS ABOUT HIT/MISS ANALYSIS

It is indeed possible to generate estimates of $a_{90/95}$ for hit/miss data that can be misleading. In this paper, we propose a Bayesian approach that involves 3 and 4 parameter models that will provide a remedy for this problem. Generally, specific statistical expertise should be consulted to avoid errors in POD studies.

First, the idea of using more than 2 parameters for the POD model has been proposed in the past (Moore et al., 2001). Additional parameters are added for 2 reasons: 1) the evidence doesn’t support the POD curve converging to one for large flaw sizes, and 2) better accounting of false call rates such that the POD curve doesn’t converge to 0 for small flaw sizes. The difficulty with conventional methods is in the estimation of the parameters. The Bayesian methods proposed in this work facilitate easy estimation of the model parameters, a simple way to compute confidence bounds, and a systematic method for model selection based on the Bayes factor.

A detailed description of Bayesian methods is beyond the scope of this paper, but an overview of how it can be applied in an NDE context is given in (Thompson, 2010), and all the mathematical details can be found in (Christensen, 2010). A very simple statement of a problem posed in Bayesian terms is shown in Figure 1.

The posterior is a probability density function that literally reads as the probability of model parameters $\theta$, given the data. The likelihood is also a probability density function that is read as the probability of the data given assumptions about the model parameters. The prior can be noninformative as will be the case for this work, or it can represent expert opinion or information from previous experiments. The normalizing constant or marginal likelihood in the denominator is actually integration
over the numerator. Most people are familiar with Bayes rule, but there is a computational challenge in calculating the normalizing constant or evidence as it is commonly called in the literature. Calculating the evidence requires a high dimensional integration that is typically performed via sampling methods such as Markov Chain Monte Carlo (MCMC).

**COMPARISON OF BAYESIAN APPROACH WITH STANDARD METHODS**

Before applying 3 and 4 parameter models to difficult data sets in (Generazio, 2011), Bayesian methods are illustrated on data from Berens seminal paper on the topic (Berens, 1989). For hit/miss analysis the data is assumed to follow a Bernoulli distribution. Either logit or probit models can be used.

\[ y_i \sim \text{Bernoulli}(p_i) \]  
\[ p_i = \frac{\exp(b_0 + b_1 \ln(a_i))}{1 + \exp(b_0 + b_1 \ln(a_i))} \quad \text{(logit)} \]  
\[ p_i = \Phi(b_0 + b_1 \ln(a_i)) \quad \text{(probit)} \]

The traditional methods for estimating parameters \( b_0 \) and \( b_1 \) are already established, and the Bayesian methods will be compared with standard cases. Once again, noninformative priors will be used in this process, and MCMC sampling will be used to extract information about the parameters. The data provided in (Berens, 1989) and displayed in Figure 7 of that work is used as a benchmark. Figure 2 displays the POD curves which appear identical to the ones using the traditional analysis methods. An alternative form of the logit model is given in equation 4, and the parameters \( \mu \) and \( \sigma \) are related to the parameters in equation 2, by equations 5 and 6. A comparison summary of the results is shown in Table 1.

\[ p_i = \left[ 1 + \exp\left( -\frac{\pi}{\sqrt{3}} \left( \frac{\ln a - \mu}{\sigma} \right) \right) \right]^{-1} \]  
\[ \mu = \frac{-b_0}{b_1} \]  
\[ \sigma = \frac{1}{b_1} \]
\[
\sigma = \frac{\pi}{b_1 \sqrt{3}} \quad \quad (6)
\]

<table>
<thead>
<tr>
<th>Parameter</th>
<th>A</th>
<th>B</th>
<th>C</th>
<th>Composite</th>
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<td>(\hat{\mu})</td>
<td>0.96/0.96</td>
<td>1.11/1.11</td>
<td>0.82/0.82</td>
<td>0.96/0.96</td>
</tr>
<tr>
<td>(\sigma)</td>
<td>0.59/0.59</td>
<td>1.04/1.05</td>
<td>0.87/0.89</td>
<td>0.88/0.85</td>
</tr>
</tbody>
</table>

Table 1. Original analysis/Bayesian methods

There are no substantial differences in the results for the 2 parameter models. The reason this Bayesian approach is introduced is because the inference needed for 3 and 4 parameter models is quite difficult using traditional methods. The sampling methods in the Bayesian approach are just as easy to implement for the 3 and 4 parameter models as they are for the 2-parameter model.

**ADDITIONAL PARAMETERS IN HIT/MISS MODELS**

The motivation for adding additional parameters to the model is to better represent false call rates for small flaw sizes, and to better represent the tail behavior of the POD curve for large flaw sizes. Equations 7 and 8 show the form of a 3-parameter model with a lower asymptote for the logit and probit models respectively. Figure 3a depicts this and the value of \(\alpha\) can be thought of as a measure of the false call rate. Equations 9 and 10 show the form of a 3-parameter model with an upper asymptote. Figure 3b depicts this 3-parameter model. The \(\beta\) term is a measure of the probability of missing flaws as the flaw size goes to infinity. The 4-parameter model has both terms in it and is described in equations 11 and 12, and depicted in Figure 3c. Proper estimation of the upper asymptote is very important for addressing pitfalls in POD analysis.

\[
p_i = \alpha + (1 - \alpha) \cdot \frac{\exp(b_0 + b_1 \log(a_i))}{1 + \exp(b_0 + b_1 \log(a_i))} \quad \quad \text{(logit)} \quad \quad (7)
\]

\[
p_i = \alpha + (1 - \alpha) \cdot \Phi(b_0 + b_1 \log(a_i)) \quad \quad \text{(probit)} \quad \quad (8)
\]
\[ p_i = \beta \cdot \frac{\exp(b_0 + b_1 \log(a_i))}{1 + \exp(b_0 + b_1 \log(a_i))} \]  
\text{ (logit) (9)}

\[ p_i = \beta \cdot \Phi(b_0 + b_1 \log(a_i)) \]  
\text{ (probit) (10)}

\[ p_i = \alpha + (\beta - \alpha) \cdot \frac{\exp(b_0 + b_1 \log(a_i))}{1 + \exp(b_0 + b_1 \log(a_i))} \]  
\text{ (logit) (11)}

\[ p_i = \alpha + (\beta - \alpha) \cdot \Phi(b_0 + b_1 \log(a_i)) \]  
\text{ (probit) (12)}

If \( \beta \) is smaller than 0.9, then that means \( a_{90} \) and \( a_{90/95} \) simply do not exist, even if the 2 parameter model provides estimates. So there is a choice between 4 models: 1) 2-parameter model, 2) 3-parameter model with lower asymptote parameter \( \alpha \), 3) 3-parameter model with upper asymptote parameter \( \beta \), and 4) 4-parameter model. Each of these can use either the logit or probit link function, so that makes a total of 8 models. The determination of which model is appropriate for a given data set is determined by looking at the ratio of marginal likelihoods a.k.a. the Bayes factor.

**AVOIDING PITFALLS IN HIT/MISS ANALYSIS**

The combination of the 3 and 4 parameter models with Bayesian methods are powerful tools that can be used to avoid pitfalls in the analysis of hit/miss data. The process is simply to calculate the marginal likelihood for each of the 8 possible models and to determine which model is most appropriate. If the 3 or 4 parameter model is selected, it is very possible that \( a_{90} \) or \( a_{90/95} \) might not exist even though inference with the 2-parameter model provides estimate. Let’s look at an example that appeared in (Generazio, 2011). There was a particular set of data that the author identified as a “Case 2 data set D8001(3)L.” It was reported that the “Logit-ML” \( a_{90/95} \) value was 12.95 mm. The latest version of MIL-HDBK-1823 was not cited in the paper, so this analysis could have been done with older software with a known deficiency, which was corrected by implementing the likelihood ratio method for confidence bound calculation (Annis et al., 2007). The value obtained for \( a_{90/95} \) using the likelihood ratio method is 22 mm and many warnings are displayed throughout the calculation (Annis, 2012).
Since there are many misses for large flaw sizes in this data set, it most likely doesn’t meet the requirements to be analyzed according to the procedures set forth in (MIL-HDBK-1823A, 2009). For this case, the data is analyzed using 8 potential models. The 2-parameter, 3-parameter with upper bound, 3-parameter with lower bound, and the 4-parameter model will all be examined for both the logit and probit models. The marginal likelihoods for each are listed in Table 1.

<table>
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<tr>
<th>Model type</th>
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<th>probit</th>
<th>Bayes factor logit/probit</th>
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<td>2-parameter</td>
<td>1.1806e-87</td>
<td>2.7022e-88</td>
<td>4.3690</td>
</tr>
<tr>
<td>3-parameter lower bound</td>
<td>2.9183e-89</td>
<td>4.2119e-89</td>
<td>0.6929</td>
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<tr>
<td>3-parameter upper bound</td>
<td>6.1396e-83</td>
<td>8.28e-83</td>
<td>0.7415</td>
</tr>
<tr>
<td>4-parameter</td>
<td>9.2848e-83</td>
<td>4.0047e-81</td>
<td>0.0232</td>
</tr>
</tbody>
</table>

Table 1. Marginal likelihoods and Bayes factors for each possible model

The ratio of evidence or marginal likelihoods for competing models is used to select the best model. This ratio is known as the Bayes factor. The marginal likelihood for an individual model doesn’t provide useful information, but the relative marginal likelihoods do. Note that all the models with upper asymptotes have far higher marginal likelihoods than models without, so all evidence suggests that an upper bound is necessary for this data set. Also note, that the Bayes factor is only greater than 1 for the 2-parameter case. Since this is the ratio logit/probit, this implies that the probit model better fits the data for all models except the 2-parameter model. The 4-parameter probit model has the highest marginal likelihood compared to all other possibilities, so it is the best choice out of all the models. It turns out
that \( a_{90} \) and \( a_{90/95} \) do not exist for this data set. The lower asymptote is 0.2657, and the upper asymptote is 0.8574. This basically indicates that the false call rate is quite high, and that the POD on average is 0.8574 as flaw size goes to infinity. Figure 4 displays the 4-parameter probit model which is most appropriate for this data set.

**SUMMARY AND CONCLUSIONS**

The proper analysis of hit/miss inspection data is still very important since this type of data is still widely collected for POD studies. A brief historical survey of the analysis of this type of data in the context of POD was given and accompanied by recent advances. Some concerns related to the statistical analysis and design of POD studies were listed and addressed. In particular, the potential to determine estimates of \( a_{90} \) or \( a_{90/95} \) with conventional 2-parameter models when the data do not support their existence was examined. An approach using more parameters in the logit or probit model to adequately describe the tail behavior of POD curves was presented. Bayesian methods can be used to determine parameter estimates for these more complicated models. Finally the Bayes factor is used to determine which model is most suitable, and should mitigate concerns about improperly estimating \( a_{90} \) and \( a_{90/95} \).

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**REFERENCES**


\[ p(\theta | y) = \frac{p(y|\theta)p(\theta)}{p(y)} \]

- **Prior** – Belief about the statistical parameters or expert opinion.
- **Evidence** – Normalizing constant or marginal likelihood that is important for model selection purposes.
- **Posterior** – Integration of experimental data and beliefs, expert opinion, and physics based models
- **Likelihood** – Probability that the data is generated from statistical model with parameters

Figure 1: Diagram of Bayesian approach used for statistical inference.
Figure 2. Data from (Berens, 1989) analyzed using Bayesian approach with noninformative priors.
Figure 3. POD model options: (a) 3-parameter with lower bound, (b) 3-parameter with upper bound, and (c) 4-parameter with lower and upper bound.
Figure 4. Probit 4-parameter model that best fits data set D8001(3)L.