Detecting intrinsic slow variables in stochastic dynamical systems by anisotropic diffusion maps

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Nonlinear independent component analysis is combined with diffusion-map data analysis techniques to detect good observables in high-dimensional dynamic data. These detections are achieved by integrating local principal component analysis of simulation bursts by using eigenvectors of a Markov matrix describing anisotropic diffusion. The widely applicable procedure, a crucial step in model reduction approaches, is illustrated on stochastic chemical reaction network simulations.

Slow manifold | dimensionality reduction | chemical reactions

Evolution of dynamical systems often occurs on two or more time scales. A simple deterministic example is given by the coupled system of ordinary differential equations (ODEs)

\[
\begin{align*}
\frac{du}{dt} &= \alpha(u, v), \\
\frac{dv}{dt} &= \tau^{-1}\beta(u, v),
\end{align*}
\]

with the small parameter \(0 < \tau \ll 1\), where \(\alpha(u, v)\) and \(\beta(u, v)\) are \(O(1)\). For any given initial condition \((u(0), v(0))\), already at \(t = O(\tau)\) the system approaches a new value \((u(t), v(t))\), where \(v\) satisfies the asymptotic relation \(\beta(u(t), v) = 0\). Although the system is fully described by two coordinates, the relation \(\beta(u(t), v) = 0\) defines a slow one-dimensional manifold which approximates the slow dynamics for \(t \gg \tau\). In this example, it is clear that \(v\) is the fast variable whereas \(u\) is the slow one. Projecting onto the slow manifold is rather easy: The fast foliation is simply “vertical”, i.e. \(u = \) const. However, when we observe the system in terms of the variables \(x = x(u, v)\) and \(y = y(u, v)\) which are unknown nonlinear functions of \(u\) and \(v\), then the “observables” \(x\) and \(y\) have both fast and slow dynamics. Projecting onto the slow manifold becomes nontrivial, because the transformation from \((x, y)\) to \((u, v)\) is unknown. Detecting the existence of an intrinsic slow manifold under these conditions and projecting onto it are important in any model reduction technique. Knowledge of a good parametrization of such a slow manifold is a crucial component of the equation-free framework for modeling and computation of complex/multiscale systems (1–3).

Principal component analysis (PCA, also known as POD) (4–6) has traditionally been used for data and model reduction in contexts ranging from meteorology (7) and transitional flows (8) to protein folding (9, 10); in these contexts the PCA procedure is used to detect good global reduced coordinates that best capture the data variability. In recent years, diffusion maps (11–17) have been used in a similar spirit to detect low-dimensional, nonlinear manifolds underlying high-dimensional datasets.

In this paper, we integrate ensembles of local PCA analyses in the diffusion-map framework to enable the detection of slow variables in high-dimensional data arising from dynamic model simulations. The proposed algorithm is built along the lines of the nonlinear independent component analysis method recently introduced in ref. 18. The approach takes into account the time dependence of the data, whereas in the diffusion-map approach the time labeling of the data points is not included. We demonstrate our algorithm for stochastic simulators arising in the context of chemical/biochemical reaction modeling.

Multiscale Chemical Reactions: A Toy Example

Consider the reversible chemical reaction [a dimerization, which is a part of several biochemical mechanisms (19, 20)] involving two molecular species \(X\) and \(Y\),

\[
X + X \rightleftharpoons Y,
\]

where \(k_1\) and \(k_2\) are the forward and backward rate constants. The probability that an additional molecule of type \(Y\) is produced from two \(X\) molecules (respectively, two molecules of \(X\) produced from one molecule of \(Y\)) in an infinitesimally small time interval \([t, t+dt]\) is \(k_1X(t)(X(t)−1)dt\) (respectively, \(k_2Y(t)dt\)), where \(X(t)\) and \(Y(t)\) are the number of molecules of type \(X\) and \(Y\) at time \(t\) (21). The chemical reaction in Eq. 3 satisfies the stoichiometric conservation law

\[
X(t) + 2Y(t) = \text{const},
\]

so that the state vector \([X(t), Y(t)]\) is restricted to a line in the phase plane. We now couple the chemical reaction in Eq. 3 with a slow production of \(X\) molecules from an external source

\[
\emptyset \xrightarrow{k_3} X,
\]

where in Eq. 5 means that the probability of the external production of an additional molecule of type \(X\) in an infinitesimally small time interval \([t, t+dt]\) is \(k_3dt\); the rate constants and the initial state are chosen in such a way that the production process in Eq. 5 is much slower than the dimerization reactions in Eq. 3. This is the case, for example, for the following choice of parameters:

\[
X(0) = 100,\ Y(0) = 100,\ k_1 = 1,\ k_2 = 100,\ k_3 = 50.
\]

The average time to produce an additional \(X\) molecule is \(k_3^{-1} = 0.02\), whereas the average times for the forward and backward dimerization are \((k_1X(0)(X(0)−1))^{-1} \approx 10^{-4}\) and \((k_2Y(0))^{-1} \approx 10^{-3}\). This finding implies that both \(X\) and \(Y\) are fast variables; yet their linear combination \(X + 2Y\) is a slow variable. The conservation law in Eq. 4 no longer holds since production was added. Instead, \(X + 2Y\) is slowly growing. To confirm this fact, we simulate the time evolution of the pair \([X(t), Y(t)]\) by using the Gillespie stochastic simulation algorithm (SSA) (21). In Fig. 1, we plot the time evolution of \(X, Y, \) and \(X + 2Y\).

This finding naturally leads to the following question: How does one detect the slow variable \(X + 2Y\) from data? A priori knowledge that we seek a linear combination of the original variables lends


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itself to fitting the coefficients of such a combination. Such fitting is, however, not possible for the general nonlinear case.

**Short Simulation Bursts**

It is convenient to analyze our approach in the diffusion limit, for which the simulation is well approximated by a stochastic differential equation (SDE). The chemical Langevin equation for the time evolution of $X$ and $Y$, which is formally derived from the corresponding chemical master equation, is given in the Itô form by refs. 22–24

$$\frac{dx}{dt} = (2k_2 y - 2k_1 x(x - 1) + k_3) dt - 2\sqrt{k_1 x(x - 1)} dw_1 + 2\sqrt{k_2 y} dw_2 + \sqrt{k_3} dw_3,$$

$$\frac{dy}{dt} = (k_1 x(x - 1) - k_3 y) dt + \sqrt{k_1 x(x - 1)} dw_1 - \sqrt{k_2 y} dw_2,$$

where $w_i (i = 1, 2, 3)$ are standard independent Brownian motions. The approximation in Eqs. 7 and 8 is also characterized by a time scale separation and possesses the slow variable $x + 2y$; multiplying Eq. 8 by two and adding it to Eq. 7 gives

$$d(x + 2y) = k_3 dt + \sqrt{k_3} dw_3.$$

Eq. 9 shows that the approximated stochastic dynamics of $x + 2y$ are decoupled from the individual dynamics of $x$ and $y$, as expected from Eqs. 3 and 4.

The Euler–Maruyama method for Eqs. 7 and 8 suggests that in a short time step $\Delta t$, the state vector $[x(t), y(t)]$ propagates to the random step vector $[x(t + \Delta t), y(t + \Delta t)]$

$$x(t + \Delta t) \approx x(t) + (2k_2 y(t) - 2k_1 x(t)(x(t) - 1) + k_3) \Delta t,$$

$$y(t + \Delta t) \approx y(t) + (k_1 x(t)(x(t) - 1) - k_3 y(t)) \Delta t + \sqrt{k_1 x(t)(x(t) - 1) + k_3 y(t)} Z_1,$$

where $Z_1, Z_2 \sim \mathcal{N}(0, \Delta t)$ are independent, normally distributed random variables with zero mean and variance $\Delta t$, and $Z_1$ and $Z_2$ correspond to the $dw_1$ and $dw_2$ terms, respectively, in Eqs. 7 and 8), which means that if we were to run many simulations for a short time step $\Delta t$, all starting at $[x(t), y(t)]$, the trajectories would end up at random locations forming a “point” cloud in the phase plane. The point cloud has a bivariate normal distribution, whose center is located at $\mu = [\mu_x, \mu_y]^T$, given by

$$\mu_x = x(t) + (2k_2 y(t) - 2k_1 x(t)(x(t) - 1) + k_3) \Delta t,$$

$$\mu_y = y(t) + (k_1 x(t)(x(t) - 1) - k_3 y(t)) \Delta t,$$

and whose two-by-two covariance matrix $\Sigma$ is

$$\Sigma = BB^T,$$

where

$$B = \sqrt{\Delta t} \left[ \begin{array}{cc} -2\sqrt{k_1 x(t)(x(t) - 1) + k_3 y(t)} & \sqrt{k_1 x(t)(x(t) - 1) + k_3 y(t)} \\ \sqrt{k_1 x(t)(x(t) - 1) + k_3 y(t)} & 0 \end{array} \right].$$

The shape of the point cloud is an ellipse because the level lines of the probability density function

$$p(x, y) = \frac{1}{2\pi \sqrt{\det \Sigma}} \exp \left\{ -\frac{1}{2} (x - \mu)^T \Sigma^{-1} (x - \mu) \right\}$$

are ellipses $(x = [x, y]^T)$. When there is a separation of time scales, the ellipses are thin and elongated. For example, for the set of parameters given in Eq. 6, the eigenvalues of $\Sigma$ for $[x, y] = [100, 100]$ are $\sigma_1^2 \approx 10^7 \Delta t$ and $\sigma_2^2 \approx 10^5 \Delta t$. These approximations mean that the long axis of the ellipse is two orders of magnitude longer than the short axis ($\sigma_1 / \sigma_2 \approx 10^7$). The eigenvector corresponding to $\sigma_1$ is approximately $[-2, 1]^T$, pointing in the direction of the fast dynamics on the line $x + 2y = \text{const}$. The second eigenvector is approximately $[1, 2]^T$, pointing in the direction of the slow dynamics.

The eigen-decomposition of the covariance matrix is simply the PCA of the local point cloud generated by the short simulation burst. We produce many short simulation bursts starting at different initialization points $[x, y]$. For each burst, we perform the PCA and estimate its covariance matrix $\Sigma_{(t)}$. The principal components of $\Sigma_{(t)}$ are the local directions of the rapidly changing variables at $[x, y]$, whereas components with small eigenvalues correspond to the slow variables.

We wish to piece together the locally defined components into globally consistent coordinates. The toy model in Eqs. 3–5 presents no special difficulty because the principal components of $\Sigma_{(t)}$ are approximately $[-2, 1]$ and $[1, 2]$ everywhere (independent of $[x, y]$). In general, however, the slow variable may be some complicated nonlinear function of the state variables. In such cases, it is not trivial to find a globally consistent slow coordinate.

**Anisotropic Diffusion Maps**

To integrate the local information into global coordinates, we use anisotropic diffusion maps (ADM), introduced in ref. 18. Suppose $u = u(x, y) = x + 2y$ (respectively, $v = v(x, y) = -2x + y$) are the slowly changing (respectively, the rapidly changing variables). Together, they define a map $f : (x, y) \rightarrow (u, v)$ from the observable state variables $x$ and $y$ to the “dynamically meaningful” coordinates $u$ and $v$. Alternatively, the inverse map $f^{-1} : (u, v) \rightarrow (x, y)$ is given by $x = x(u, v)$ and $y = y(u, v)$. The point cloud in the observable $(x, y)$ plane, generated by the short bursts, is the image under $f$ of a similar point cloud in the inaccessible $(u, v)$ plane. The slow manifold (curve) in the $(x, y)$ plane can be thought of as the image of the $u$ axis, $f(0, 0) = [x(u, 0), y(u, 0)]$. The ellipses in the $(u, v)$ plane are also thin and elongated, and they share an important property: They all have the $v$ axis as their long axis and...
words, the change in x with the (isotropic) graph Laplacian in the inaccessible space. We formulate our method in a general setting. Then we apply it to the simulated trajectories. For each data point \( x^{(i)} \), we generate an ensemble of short simulation bursts initialized at the data point, i.e. \( x^{(i)}(0) \equiv \{ x_1(0), x_2(0), \ldots, x_N(0) \} \). We collect the statistics of the simulated trajectories after a short time period \( \Delta t \). In particular, we compute the averaged position \( \mu_j^{(i)} = \frac{1}{\Delta t} \sum_{t=\Delta t}^{2\Delta t} x(t) \), \( j = 1, \ldots, M \), and the elements of the covariance matrix \( \Sigma^{(i)} = \{ \Sigma_{jk} \}, k = 1, \ldots, M \), by 
\[
\Sigma_{jk}^{(i)} = \frac{1}{\Delta t} \langle [y_j(\Delta t) x_k(\Delta t) - \mu_j^{(i)} - \mu_k^{(i)}] \rangle
\]
where the notation \( \langle \cdot \rangle \) stands for statistical averaging over many simulated trajectories. For each data point \( x^{(i)} \), we calculate \( \Sigma_j^{(i)} \), the inverse of the covariance matrix. We define a symmetric \( \Sigma \)-dependent squared distance between pairs of data points in the observable space \( \mathbb{R}^M \);
\[
d_j^2(x^{(i)}, x^{(j)}) = \frac{1}{2} (x^{(i)} - x^{(j)})^T \left( (\Sigma_j^{(i)})^{-1} + (\Sigma_j^{(j)})^{-1} \right) (x^{(i)} - x^{(j)}).
\]
Note that for the toy model in Eqs. 3–5 the distance \( d_M \) is a second order approximation of the Euclidean distance in the inaccessible \((u, v)\)-space;
\[
d_2(u^{(i)}, u^{(j)}) \approx \left\| u^{(i)} - u^{(j)} \right\|^2 + \tau^2 \left\| v^{(i)} - v^{(j)} \right\|^2.
\]
Because \( \tau \) is a small parameter, \( d_M \) is controlled by the difference in the slow coordinate. The approximation in Eq. 13 is also valid in higher dimensions, where there may be more than one slow coordinate \( u \) and several fast coordinates \( v \) and the ellipse is replaced by an ellipsoid. In such cases,
\[
d_M^2(u^{(i)}, u^{(j)}) \approx \left\| u^{(i)} - u^{(j)} \right\|^2 + \tau^2 \left\| v^{(i)} - v^{(j)} \right\|^2.
\]
Therefore, the ADM based on the “dynamic proximity” \( d_M \) approximates the Laplacian on the slow manifold. We construct an \( N \times N \) weight matrix \( W \)
\[
W_{ij} = \exp \left\{ -\frac{d_j^2(x^{(i)}, x^{(j)})}{\epsilon^2} \right\},
\]
where \( \epsilon > 0 \) is the single parameter of the method. The elements of the matrix \( W \) are all \( \leq 1 \). Nearby points (i.e., their projection on the slow manifold is close) have \( W_{ij} \) close to 1, whereas distant points have \( W_{ij} \) close to 0. Next, we define a diagonal \( N \times N \) normalization matrix \( D \) whose values are given by the row sums of \( W \)
\[
D_{ii} = \sum_{k=1}^N W_{ik}.
\]
We then compute the eigenvalues and right eigenvectors of the row stochastic matrix
\[
A = D^{-1/2} W D^{-1/2},
\]
which can be viewed as a Markov transition probability matrix for a jump process over the data points \( x^{(i)} \), \( i = 1, \ldots, N \). The discrete jump process converges in the limit of \( N \to \infty \) and \( \epsilon \to 0 \) to a continuous diffusion process over the observable data manifold. The diffusion process is anisotropic due to the metric \( d_M \), so that the diffusion coefficient changes with direction. Therefore, the eigenvectors of \( A \) are approximately the eigenvalues of the anisotropic differential diffusion generator over the observable manifold. The approximation in Eq. 14 implies that the slow manifold. These eigenvectors are functions of the slow \( u \) coordinates that do not depend on the fast \( v \) variables. Hence, the low order eigenvectors of \( A \) give an approximate parametrization of the slow manifold.

As discussed in refs. 12 and 25–27, the leading eigenvectors may be used as a basis for a low-dimensional representation of the data. To compute those eigenvectors, we use the fact that the leading eigenvectors of \( A \) are similar and thus have the same spectrum. Because \( S \) is symmetric, it has a complete set of eigenvectors \( \psi_j \), \( j = 0, \ldots, N - 1 \), with corresponding eigenvalues
\[
\lambda_0 \geq \lambda_1 \geq \ldots \geq \lambda_N - 1.
\]
The right eigenvectors of \( A \) are given by 
\[
u_{ij} = D^{1/2} \psi_j,
\]
Because \( A \) is a Markov matrix, all its eigenvalues are \( \leq 1 \), with largest eigenvalue \( \lambda_0 = 1 \) and a corresponding trivial eigenvector \( \psi_0 = [1, 1, \ldots, 1] \). We define the low-\( n \)-dimensional representation of the state vectors by the following ADM
\[
\Psi : \mathbb{X} \to \psi_{ij} = [u_1^{(i)}, u_2^{(i)}, \ldots, u_n^{(i)}],
\]
that is, the point \( x^{(i)} \) is mapped to a vector containing the \( ij \)th coordinate of each of the first \( n \) leading eigenvectors of the matrix \( A \). The variables \( u_1^{(i)}, u_2^{(i)}, \ldots, u_n^{(i)} \) (which are defined on the data points) are the candidate slow variables that we were looking for.

**Application of ADM to the Toy Example**
We use \( N = 2000 \) data points \( x^{(i)} = [x_1^{(i)}, x_2^{(i)}] = [X^{(i)}, Y^{(i)}], i = 1, \ldots, 2000 \), uniformly sampled from the illustrative trajectory of Fig. 1 (in fact, the trajectory in Fig. 1 is visualized using these 2000 data points). For every data point \( x^{(i)} = [X^{(i)}, Y^{(i)}], i = 1, \ldots, 2000 \), we run \( 10^7 \) replicas of stochastic simulations initialized at the data point for time \( \Delta t = 10^{-4} \). We estimate \( \mu_j^{(i)} \) and \( \sigma_j^{(i)} i = 1, \ldots, 2000 \), \( j = 1, 2, k = 1, 2 \) by Eqs. 10 and 11 as an average over \( 10^7 \) realizations. For each data point \( x^{(i)} = [X^{(i)}, Y^{(i)}], \)
we also calculate the inverse covariance matrix and the symmetric \( \Sigma \)-dependent squared distance \( d^2_L(x^{(i)}, x^{(j)}) \) by Eq. 12. We construct a \( 2000 \times 2000 \) weight matrix \( W \) by Eq. 15 for \( \varepsilon = 0.1 \) and a matrix \( A \) by Eq. 16. We compute the leading eigenvectors \( u_1 \) of \( A \) by Eq. 18. In Fig. 2, we plot our dataset where the points are colored according to the first nontrivial eigenvector \( u_1 \). We see that the eigenvector \( u_1 \) gives a good description of slow dynamics of this system. The slow dynamics are given by function \( X + 2Y \) as can be seen in the Right frame of Fig. 1. The plot of \( u_1 \) vs. \( X + 2Y \) is shown in Center frame of Fig. 2. We again confirm that we obtained a good slow description of the system. Finally, plotting the eigenvector \( u_1 \) vs. \( X \) confirms that \( X \) is not a good slow variable (Right frame of Fig. 2).

Here we used a simulation burst of \( 10^7 \) trajectories. The number of simulation “bursts” needed to construct a distance metric based on the covariance in a high-dimensional system depends on the dimensionality and the desired degree of approximation. The central limit theorem suggests that the estimated covariance matrix entries converge with the square-root number of simulated trajectories. However, the convergence of the eigenvalues and eigenvectors (principal components) of the covariance matrix depends on the dimensionality \( M \) (see, e.g. ref. 28) as crossings of eigenvalues may occur.

**Oscillating “Half-Moons”**

Next, consider the system of stochastic differential equations

\[
\begin{align*}
\mathrm{d}t &= a_1 \mathrm{d}t + a_2 \mathrm{d}w_1, \\
\mathrm{d}v &= a_3 (1 - v) \mathrm{d}t + a_4 \mathrm{d}w_2,
\end{align*}
\]

where \( a_i, i = 1, 2, 3, 4 \), are constants and \( \hat{w}_i, i = 1, 2 \) are independent \( \delta \)-correlated white noises (Wiener processes). We consider Eqs. 20 and 21 together with the following nonlinear transformation of variables

\[
x = v \cos(u + v - 1), \quad y = v \sin(u + v - 1).
\]

We will assume that the observables \( x \) and \( y \) are the actual observables, whereas \( u \) and \( v \) are unknown. We choose the values of parameters as \( a_1 = a_2 = 10^{-3}, a_3 = a_4 = 10^{-1} \). The illustrative trajectory that starts at \( x(0), y(0) \) is plotted in the Left frame of Fig. 3. The trajectory is colored according to time. We run simulations for a longer time \( 8 \times 10^4 \), which accounts for about 12–13 rotations, and record 2000 data points at equidistant time intervals of length \( 8 \times 10^3 \). This dataset is plotted in the Center frame of Fig. 3. Again, points are colored according to time. We clearly see that there is no correlation between time and the slow variable (which is \( u \) \( \text{MOD} \) \( 2\pi \)) because of oscillations.

To apply the ADM, we run \( 10^6 \) replicas of stochastic simulations initialized at each data point \( x^{(i)} = [x^{(i)}, y^{(i)}] \) for a time step \( \Delta t = 0.1 \) and estimate \( \mu_k^i \) and \( \sigma_k^i \) for \( i = 1, \cdots, 2000, j = 1, 2, k = 1, 2 \) by Eqs. 10 and 11 as an average over \( 10^6 \) realizations. For each data point \( x^{(i)} = [x^{(i)}, y^{(i)}] \), we also calculate the inverse covariance matrix and the symmetric \( \Sigma \)-dependent squared distance \( d^2_L(x^{(i)}, x^{(j)}) \) by Eq. 12. Next, we have to choose the value of parameter \( \varepsilon \). To do that, we construct the \( \varepsilon \)-dependent \( 2000 \times 2000 \) weight matrix \( W = W(\varepsilon) \) by Eq. 15 for several values of \( \varepsilon \). Then we compute

\[
L(\varepsilon) = \sum_{j=1}^{N} \sum_{i=1}^{N} W_{ij}(\varepsilon).
\]

The function \( L(\varepsilon) \) is plotted in the Right frame of Fig. 3 (it is a log–log plot) (29). It clearly has two constant asymptotes when \( \varepsilon \to 0 \) and \( \varepsilon \to \infty \); as we expect, these asymptotes are smoothly connected, by an approximately straight line of slope \( d \) in a log–log plot, where \( d \) is the dimension of the slow manifold. Thus, the log–log plot of \( L(\varepsilon) \) suggests to choose \( \varepsilon \) where the log–log graph of \( L(\varepsilon) \) appears linear. We choose \( \varepsilon = 6 \). We form \( A \) by Eq. 16 and compute its few leading eigenvectors \( u_1 \) by Eq. 18. The first nontrivial eigenvector \( u_1 \) then describes the slow dynamics of the

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**Fig. 2.** Toy example. (Left) The dataset with each point colored according to \( u_1 \). (Center) Vector \( u_1 \) as a function of \( X + 2Y \). (Right) Vector \( u_1 \) as a function of \( X \).

**Fig. 3.** Oscillating half moons. The short illustrative trajectory of Eqs. 20–22 which starts at \( x(0), y(0) \) is [1, 0]. (Left) The trajectory is colored according to time. The representative dataset sampled at equal time steps from a longer stochastic simulation. (Center) The points are colored according to time. (Right) Plot of \( L(\varepsilon) \) given by Eq. 23.
The reactions in Eq. 16094 degradation of are assumed to occur on an intermediate time scale (i.e. slower than the fast time scale, but faster than the slow time scale). We consider the following set of chemical reactions

\[
X \xrightarrow{k_1} X + Z, \quad Y + Z \xrightarrow{k_2} Y, \quad \emptyset \xrightarrow{k_3} Y, \quad Y \xrightarrow{k_5} \emptyset, \quad \emptyset \xrightarrow{k_5} X. \tag{24}
\]

The first two reactions in Eq. 24 are production and degradation of \(Z\) (catalyzed by \(X\) and \(Y\), respectively). The production and degradation of \(Z\) is assumed to be happening on a fast time scale. The reactions in Eq. 25 are production and degradation of \(Y\). They are assumed to occur on an intermediate time scale (i.e. slower than the fast time scale, but faster than the slow time scale). The reaction in Eq. 26 is production of \(X\), which is assumed to be slow.

We choose the values of the rate constants as

\[
k_1 = 1000, \quad k_2 = 1, \quad k_3 = 40, \quad k_4 = 1, \quad k_5 = 1. \tag{27}
\]

This choice of rate constants guarantees that the reactions in Eq. 24 are the fastest, the reactions in Eq. 25 happen on a faster time scale, and the reaction in Eq. 26 is the slowest. The model in Eqs. 24–26 is approximated by the ODE system for the \(O(1)\) variables \(x = X/100, y = Y/40\) and \(z = Z/2500\) as follows: \(dx/dt = k_5/100, dy/dt = k_3/40 - ky, dz/dt = 100k_1x/2500 - 40ky2z\). By using the parameter values in Eq. 27, we obtain \(dx/dt = x/100, dy/dt = 1 - y, dz/dt = 40(x - y2)\). The quasiequilibrium approximation in the \(z\) equation (fastest) is \(z = x/y\), which gives rise to the “half-moon shaped” profile (hyperbola + noise) dynamics in the \(Y-Z\) plane. The variable \(y\) changes on a faster time scale than \(x\). Roughly speaking, the fluctuations in \(y\) lead to the dynamics in \(z\) according to the formula \(z = x/y\), where \(x\) changes very slowly, as illustrated in Fig. 5. We initialize the system at \([X(0), Y(0), Z(0)] = [100, 40, 2500]\) and simulate the time evolution using the Gillespie stochastic simulation algorithm. Fig. 5 shows the time evolution of \(X\) (Upper Left frame), \(Y\) (Upper Center frame), and \(Z\) (Upper Right frame). The same trajectory plotted in the \(Y-Z\) plane is given in the Lower frames of Fig. 5. We plot the

\[
X \rightarrow X + Z, \quad Y + Z \rightarrow Y, \quad \emptyset \rightarrow \emptyset, \quad \emptyset \rightarrow X. \tag{26}
\]

Inherently Nonlinear Chemical Reactions

We consider the following set of chemical reactions

We confirm that we recovered the slow dynamics correctly. Plotting \(u1\) against the observable \(x\) confirms that the latter is not a good observable (Center frame of Fig. 4). The slow variable is given as a nonlinear transformation of \(x\) and \(y\) which can be computed by inverting Eq. 22 locally. It is basically a function of \(u \text{ MOD } 2\pi\). The eigenvector \(u1\) is plotted against the slow variable \(u \text{ MOD } 2\pi\) in the Right frame of Fig. 4. We again confirm that we recovered the slow dynamics correctly.

Fig. 4. Oscillating half moons. (Left) The dataset with each point colored according to \(u1\). (Center) Vector \(u1\) as a function of \(x\). (Right) Vector \(u1\) as a function of \(u \text{ MOD } 2\pi\).
2000 data points lying on this trajectory colored by time (Lower Left frame). In the Lower Center frame of Fig. 5, we provide the similar Y-Z plot where the data points are colored according to the value of X.

The set of 2000 data points (plotted in the Lower Right frame of the Fig. 5) is the input of the diffusion-map approach. To emphasize the strength of our approach, the data points are ordered randomly in the inputting dataset. In our model, the slow variable X is a nondecreasing function of time t (see Fig. 5 Upper Left frame). Consequently, the dataset recorded from the stochastic simulations is ordered according to the slow variable. In more complicated chemical examples [e.g. problems with oscillations (30)], or the oscillating half-moons from the previous example, there is no obvious relation between the “dynamic proximity” of data points and the order in which they are recorded. Our approach works in more complicated situations because the ADM is independent of the order of the inputting data points.

We use short bursts of time \( \Delta t = 5 \times 10^{-4} \) (which correspond to approximately 100 Gillespie SSA time steps) of stochastic simulations initialized at the \( N = 2000 \) data points from Fig. 5 (Lower Right frame). For every data point \( X(t) = [X^1(t), X^2(t), X^3(t)] \), \( i = 1, \ldots, N \), we run 106 replicates of stochastic simulations initialized at the data point to estimate the covariance matrix \( \Sigma(t) \). We use \( \epsilon = 1 \). In the Right frame of Fig. 6, we plot our dataset [given in Fig. 5 (Lower Right frame)] and we color the data points according to the first nontrivial eigenvector \( u_1 \). We see that the eigenvector \( u_1 \) gives a good description of slow dynamics of

the system in Eqs. 24–26. The slow dynamics can be described by the variable X, as can be seen in the Upper Left frame of Fig. 5. The plot of \( u_1 \) vs. X is shown in the Center frame of Fig. 6. We again confirm that we obtained a good description of the slow dynamics of the system. Finally, plotting the eigenvector \( u_1 \) vs. Y confirms that Y is not a slow good variable (Right frame of Fig. 6).

**Summary**

Finding a reduced model for dynamical systems with a large number of degrees of freedom is of great importance in many fields. Dimensional reduction methods often use similarity measures between different states of the dynamical system to reveal its low-dimensional structure. Those methods are limited when the similarity measure does not take into account the time-labeling of the states. We encode the time dependence into an anisotropic similarity measure by using short bursts of local simulations. The resulting leading eigenvectors of the anisotropic diffusion map approximate the eigenfunctions of the Laplacian over the manifold corresponding to the dynamically meaningful slowly varying coordinates. We demonstrated the usefulness of the ADM in analyzing dynamical systems by its successful recovery of meaningful coordinates in the particular case of multiscale chemical reactions.

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