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**THESIS**

ESTIMATION OF RETENTION RATES OF P-CODED  
NAVAL OFFICERS CONDITIONED ON THEIR  
YEAR OF GRADUATION AND SUBSPECIALTY CODE.

by

10 Heinz Dieter Mueller

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Thesis advisor: R. A. Weitzman

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20 (continued)

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It is found that the model yields reliable estimates of the retention rates that can provide a meaningful substitute for actually observed rates especially within prediction models. The author finally recommends an approach that extends his model into a prediction model for the retention-rates of P-coded naval officers.

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Estimation of Retention Rates of P-Coded Naval Officers  
Conditioned on Their  
Year of Graduation and Subspecialty Code

by

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## ABSTRACT

This thesis analyzes retention behavior of P-coded naval officers and derives a model to estimate retention rates for this group of officers conditioned on their respective year of graduation and the Subspecialty-code obtained following completion of postgraduate education. The model is based on Bayes' estimation technique.

The estimates resulting from applying the model to the data obtained by observing the behavior of 3,981 naval officers who were graduated between 1970 and 1975 and who obtained one of 41 selected Subspecialty-codes are analyzed with respect to common trends and differences in the behavior as well as with respect to the usefulness of the Bayes' estimation technique underlying the model.

It is found that the model yields reliable estimates of the retention rates that can provide a meaningful substitute for actually observed rates especially within prediction models. The author finally recommends an approach that extends his model into a prediction model for the retention-rates of P-coded naval officers.

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## I. INTRODUCTION

The estimation of retention rates for P-coded naval officers has been a major problem for personnel planners. The procedure currently used by the Navy overestimates the actual need.<sup>1</sup>

One of the reasons for this overestimation is the fact that the samples available are too small to provide a basis for reliable estimation by traditional procedures.

In a Memorandum of 21 September 1979, Weitzman<sup>2</sup> suggests two approaches to a solution to this problem:

- a. Aggregation of subsamples to increase the reliability of traditional estimators;
- b. Development of new estimation models.

One of the basic devices to cope with the problem of too small sample sizes in this context is the Bayesian estimation technique. It could provide the basis for a new estimation model.

It is the purpose of this study to derive a Bayesian estimator which will yield reliable estimates of retention rates of P-coded naval officers conditioned on their respective year of graduation and their respective Subspeciality-code obtained after graduation.

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<sup>1</sup>Weitzman, R.A., Memorandum ND4(54Wz)/bd 21 Sept 1979: Naval Officer Subspecialty Analysis.

<sup>2</sup>See Weitzman, R. A., Memorandum NC4

The data base for this study consisted of 3,981 naval officers who were graduated in the years 1970 to 1975 and who attained one of 41 selected Subspecialty-codes which are related to the different curricula of the Naval Postgraduate School. The retention behavior of these officers was observed from the year of graduation up to the year 1979, which was the last year that corresponding data were obtained from the Officer Master File and Attrition File. The observed retention rates are analyzed mainly with respect to

- differences in the behavior of the graduates over the six subsequent graduation years,
- trends within the respective graduation-year groups over years  $k$  after graduation, and
- differences between the ten Subspecialty-code groups into which the 41 Subspecialty-codes were grouped.

Based on the Bayesian estimator for this study and the specific phenomena observed, a model is formulated that allows use of the estimator to obtain an estimate of the retention rate that can provide a more reliable substitute than the actually observed rate, especially within prediction models.

The estimates are calculated and their usefulness as a reliable basis for prediction models is analyzed. The analysis and discussion of the obtained results are concluded by correlating the estimates with their respective rates. This procedure led to final conclusions about the actual retention behavior as well as about the usefulness of the Bayesian estimation technique within this context.

Finally, recommendations are given with respect to possible extensions of the derived model into prediction models.

## II. BAYESIAN ANALYSIS

### A. IDEA OF BAYESIAN ANALYSIS

Bayesian Analysis is an attempt to incorporate into the process of statistical inference all information about the underlying state of nature of a random phenomenon. It tolerates explicitly the use of subjective judgement where a priori verifiable information is not available.<sup>3</sup>

Suppose there exists a set of mutually exclusive events, say officer A remains in service or he leaves service in a certain year k and a priori there exists no certainty about his behavior. Then the Bayesian estimation technique allows for assigning prior probabilities to each of these events on the basis of whatever evidence is known or subjectively assumed in advance. Then, if additional facts become available, for example retention rates for a group of officers to whom officer A may belong, the initial probabilities are revised by means of Bayes' Theorem. As a result of this revision process, posterior probabilities are obtained. They do not completely supersede the prior information, but they contain it.

Whether this prior information is still useful after additional evidence has been obtained depends on the phenomenon under consideration. In theory as well as in practice

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<sup>3</sup>Morgan, B. W., An Introduction to Bayesian Statistical Decision Processes, p. 1-14, and Preface, Prentice-Hall, 1968.

the above described distinctive feature of Bayesian analysis is the subject of controversial opinions. Examples of bizarre results of Bayesian analysis can be found in almost all standard statistics books.<sup>4</sup>

## B. BASIC STRUCTURE OF BAYESIAN ANALYSIS<sup>5</sup>

Suppose that it is possible to summarize a priori information about the relative likelihood of where in a specified parameter space  $\Omega$  the unknown value of parameter  $\theta$  lies by constructing a probability distribution for  $\theta$  on  $\Omega$ .

Assuming a continuous case, the p.d.f.  $\xi(\theta)$  of this distribution is then called the prior p.d.f. of  $\theta$ .

Suppose then that a random sample  $X_1, \dots, X_n$  (Vector notation:  $\tilde{X}$ ) is drawn from a distribution with p.d.f.  $f(x|\theta)$ , where the value of  $\theta$  is unknown and the prior distribution of  $\theta$  has the p.d.f.  $\xi(\theta)$ . The joint p.d.f. of the random variables  $X_1, \dots, X_n$  is in vector notation  $f(\tilde{X}|\theta)$  with marginal joint p.d.f.

$$(1) \quad f(\tilde{X}) = \int_{\Omega} f(\tilde{X}|\theta) \xi(\theta) d\theta .$$

Then after  $X_1 = x_1, \dots, X_n = x_n$  has been observed, the conditional p.d.f. of  $\theta$  given these values

---

<sup>4</sup>Example in Wonnacott, T. H., and Wonnacott, R. J., Introductory Statistics, Third Edition, p. 591-593, John Wiley & Sons, 1977.

<sup>5</sup>The argumentation in this chapter is based on: DeGroot, M. H., Probability and Statistics, Menlo Park, 1975. Winkler, R. L., Introduction to Bayesian Inference and Decision, Holt, Rinehart and Winston, Inc., 1972. Wonnacott, Introductory Statistics, 1977.

$$(2) \quad \xi(\theta|\tilde{X}) = \frac{f(\tilde{X}|\theta) \xi(\theta)}{\int_{\Omega} f(\tilde{X}|\theta) \xi(\theta) d\theta} \quad \text{for } \theta \text{ in } \Omega$$

is denoted as the posterior p.d.f., where  $f(\tilde{X}|\theta)$  represents the likelihood function and  $\xi(\theta)$  the prior distribution.

Equation (2) is exactly the Bayes Theorem for continuous random variables. As Winkler<sup>6</sup> states, this equation provides conceptually a convenient way to revise prior information when additional evidence by means of sample information is obtained. However, except for relatively simple mathematical functions, it might prove to be impossible to carry out the required integration.

These potential difficulties led to resorting to the concept of conjugate prior distributions: in essence families of distributions for which the likelihood function  $\xi(\theta|\tilde{X})$  is uniquely determined once a data-generating model is specified.

Conjugate families of distributions corresponding to some likelihood functions that are important for practical purposes have been developed. In the context of this study, in which sampling from a stationary and independent Bernoulli process is the underlying data-generating model, the conjugate family is the family of Beta-distributions.

Suppose the retention behavior of a randomly chosen group of  $n$  officers has been observed.

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<sup>6</sup>Winkler, R. L., p. 147.

$$\text{Let } X_1 = \begin{cases} 1 & \text{if officer 1 remained in service in year } k \\ 0 & \text{otherwise} \end{cases}$$

Then  $X_1, \dots, X_n$  form random sample from a Bernoulli distribution for which the value of  $\theta (0 < \theta < 1)$  is unknown.

Assuming that the prior distribution of  $\theta$  is a Beta-distribution with parameters  $\alpha$  and  $\beta$  ( $\alpha > 0, \beta > 0$ ), then the posterior distribution of  $\theta$  given that  $X_1 = x_1, \dots, X_n = x_n$  is a Beta-distribution with parameters

$$\alpha + \sum_{i=1}^n x_i \quad \text{and} \quad \beta + n - \sum_{i=1}^n x_i .$$

### C. DEFINITION OF THE BAYES' ESTIMATOR FOR THIS STUDY

#### 1. Bayes' Estimator

Based on the observed values of the random vector  $\tilde{X}$ , the value of  $\theta$  can be estimated. Thus the estimator of  $\theta$  is a real valued function of  $\tilde{X}$  denoted by  $\delta(X)$ .

To determine the goodness of the estimator in terms of the closeness of the estimate to the true value of the parameter  $\theta$ , the quadratic loss function is used as suggested by most statisticians:<sup>7</sup>

$$(3) \quad L(\theta, a) = (\theta - a)^2 .$$

Hence the estimate should be chosen such that  $E[(\theta - a)^2 | \tilde{X}]$  is minimal, where the expected loss is

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<sup>7</sup>See DeGroot, p. 276, and Wonnacott, p. 573.

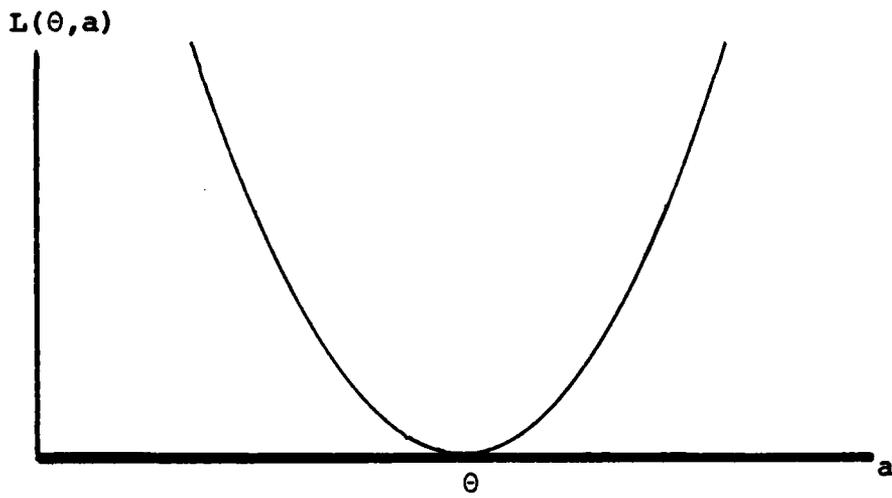


Figure I: Quadratic Loss Function for a Given  $\theta$

$$(4) \quad E[(\theta - a)^2 | \tilde{X}] = \int_{\Omega} (\theta - a)^2 \xi(\theta | \tilde{X}) d\theta,$$

and  $\xi(\theta | \tilde{X})$  denotes the posterior p.d.f., defined with equation (2).

The Bayes' estimator of  $\theta$  is now the function  $\delta^*(\tilde{X})$  which yields for every possible value  $\tilde{x}$  of the random vector  $\tilde{X}$  an estimate for which the loss according to equation (4) is minimal. For the conjugate family of Beta-distributions

$$(5) \quad \delta^*(X) = \frac{\alpha + \sum_{i=1}^n x_i}{\alpha + \beta + n}, \quad \text{i.e.}$$

$\delta^*(X)$  is the mean of the posterior Beta-distribution.

Given that the optimal estimate  $\delta^*(X)$  is found, then it is easily possible to use this revised information as new prior information for a second application of Bayes' Theorem.<sup>8</sup>

In this case the posterior distribution after the first application of Bayes' Theorem becomes the prior distribution for the next stage and can again be revised by newly obtained information on the basis of a second sample.

The nature of Bayesian estimation implies that the obtained optimal estimate is a compromise between two or more sets of sample information.<sup>9</sup> In case of considerable differences in the sizes of obtainable samples, larger samples

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<sup>8</sup>See Winkler, p. 160-162.

<sup>9</sup>Wonnacott, p. 546.

gain more weight than do smaller samples. This might lead to the effect that, independent of the quality of a set of sample information, the quantity in terms of the size of the sample used in the revision process determines the value of the estimate.

However, as long as with decreasing sample size the quality of the obtainable information decreases, this effect is desirable. Looking at P-coded officers who belong to a certain subspecialty code, it might for example be that within a certain period the behavior of only three officers is observable. If two of those officers happened to die in a car accident, the retention rate for this year, conditioned on the Subspecialty-code under consideration, would be  $1/3$ . This mathematically "true" rate would undoubtedly be completely meaningless as an entry in a prediction model for retention behavior.

## 2. The Special Problem of a Diffuse Prior Distribution

Consider a situation in which a prior distribution has to be assessed without sufficient a priori information relative to some "overwhelming" sample information. Winkler denotes this state as being diffuse relative to the sample information.<sup>10</sup>

Generally this diffuse state is most appropriately represented by a diffuse conjugate prior distribution,

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<sup>10</sup>Winkler, p. 198.

specifically the Uniform (0,1)-distribution which is a Beta-distribution with parameters  $\alpha = 1$  and  $\beta = 1$ .<sup>11</sup> Thus equal probabilities are assigned to all possible values of  $\theta$  within  $\Omega$ , and according to Morgan<sup>12</sup> the maximum possible error will be minimized. However, there exist situations in which the applicability of the Uniform (0,1)-distribution as a diffuse prior distribution is doubtful.<sup>13</sup>

Suppose that, of a second sample of three officers, one remained in service in year  $k$ , i.e. the observed actual retention rate for year  $k$  is  $1/3$ . Suppose further that prior information is represented by the Uniform (0,1)-distribution. Then the Bayes' estimate will be

$$\delta^* = \frac{1 + 1}{1 + 1 + 3} = \frac{2}{5} .$$

This shows that in cases where extremely small samples are involved the Uniform (0,1)-distribution which was supposed to have no influence on the posterior distribution does not act as a diffuse prior distribution. For practical purposes in an inferential or decision-making situation, Winkler suggests therefore to use as parameters of the diffuse prior distribution  $\alpha = 0$  and  $\beta = 0$ .<sup>14</sup> From a theoretical standpoint

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<sup>11</sup>See Winkler, p. 198-201 and Morgan, p. 46.

<sup>12</sup>Morgan, p. 47.

<sup>13</sup>See Winkler, p. 198-204.

<sup>14</sup>Winkler, p. 202.

this implies the involvement of an improper Beta-distribution where the total area under the p.d.f. does not equal to one. However, looking at the parameters of the prior distribution as equivalent to an a priori information status, this approach seems to be the only choice, as Winkler states it.<sup>15</sup>

The effect of this approach on the value of the estimate is the same as the effect of a large sample size in the revision process relative to a small prior base. In both cases the posterior distribution will solely depend on the sample information.

### 3. The Bayes' Estimator Used in This Study

To estimate retention rates of officers conditioned on their respective Subspecialty codes the idea of the successive application of Bayes' Theorem will be used.

It is assumed that a priori the individual decision of an officer to remain in service in year k after he was graduated and thus obtained his P-code is appropriately represented by a diffuse prior Beta-distribution.

As the first sample information in the revision process of this prior information, the retention behavior of the group m of officers will be measured who belong to the same Subspecialty-code group.<sup>16</sup> The resulting posterior distribution will then be used as the prior distribution

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<sup>15</sup>Winkler, p. 202

<sup>16</sup>A list of the observed Subspecialty-code groups and the Subspecialty codes belonging to them is contained in Table 1 and Appendix A-1.

for one more application of Bayes' Theorem. As the second sample information, the retention behavior of a subgrouping  $n$  of officers belonging to the same Subspecialty-code within the above used Subspecialty-code group will be measured and used to revise the prior distribution obtained in the first step.

In addition to the assumption stated above, Winkler's suggestion concerning the diffuse prior distribution will be followed, i.e., in calculating the final estimate, the parameters of the original prior distribution will be treated as if they were zero.

Let

$y$  denote the number of officers belonging to  $n$  who left service up to and including year  $k$  after graduation and let

$z$  denote the number of officers belonging to  $m$  who left service up to and including year  $k$  after graduation.

Then as Bayes' estimate of the rate of officers who left service up to and including year  $k$  after graduation and belong to both of the above described groupings  $n$  and  $m$

$$(6) \quad \delta^* = \frac{z + y}{m + n}$$

will be used.<sup>17</sup>

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<sup>17</sup>This is one of the estimators that Professor R. A. Weitzman (NPS) has suggested for use in pattern analysis to cope with the problem of empty or near-empty cells (personal communication).

### III. DATA BASE

#### A. DESCRIPTION OF THE DATA

Data for this study have been provided by the Department of the Navy.<sup>18</sup> They were extracted from both the Officer Master File (OMF) and the Attrition File (AF). In order to conduct the study with a data base as broad as possible, records of a total of 6,372 American Navy officers were established. The officers were selected according to three criteria:

- a. They have been assigned a "P"-code.<sup>19</sup>
- b. They were graduated from a graduate-level program between 1960 and 1975, inclusive.
- c. They have obtained one of the 41 Subspecialty-codes (SSC) listed in Table 1.<sup>20</sup>

To be able to extract retention rates conditioned on the SSC's selected, the records obtained from above mentioned files have to represent the complete group of naval officers satisfying the three criteria. First preliminary summary statistics about the losses between 1960 and 1969 revealed that the number of lost officers seemed to be zero. It was

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<sup>18</sup> Department of the Navy - NMPC, 8 April 1980.

<sup>19</sup> Suffix "P" means "Master level"; in the context of this study a P-coded officer has reached at least the Master level.

<sup>20</sup> The SSC's are related to curricula at NPS (OPNAVNOTE 1520, 25 June 79).

SSC group	Subspecialty codes (SSC)						
xx2x	xx21	xx22	xx23	xx24	xx25	xx26	xx27
xx3x	xx31	xx32	xx33	xx34	xx38		
xx4x	xx42	xx44	xx48	xx49			
xx5x	xx51	xx52	xx54	xx55	xx56		
xx6x	xx61	xx62	xx63	xx67			
xx7x	xx71	xx72					
xx8x	xx81	xx82					
xx9x	xx91	xx95					
11xx	1101	1102	1103				
13xx	1301	1302	1304	1305	1306	1307	1308

Table 1: List of the 41 SSC's Selected for This Study

found that the data available for this period in the form needed contain only officers who were still in service past 1969. Therefore the final data base consists of those officers fulfilling criteria a. and c. who were graduated in the years 1970 to 1975. Their total numbers for the graduation years indicated are shown at the top of the next page.

Data obtainable for each individual and sufficient to conduct the study were structured in records as shown in Appendix A-2. Table 2 contains a listing of the data. In order to be able to calculate the retention rates for each year k after graduation, ten entries for the vector X--which was

Graduation year	Total number of officers
1970	562
1971	715
1972	768
1973	669
1974	649
1975	618
-----	
Total data base	3,981
=====	

introduced in Chapter II--were added. They contain for the 10 calendar years under consideration (1970 - 1979) the value

$$x = \begin{cases} 1 & \text{if the officer was graduated and still serves} \\ & \text{in this year,} \\ 0 & \text{otherwise.} \end{cases}$$

These modified records served as the basis for any further calculations. A detailed listing of the number of officers who were graduated in the years 1970 to 1975, categorized according to their respective Subspecialty-code (SSC) is contained in Appendix B.

**B. RESTRICTIONS AND ASSUMPTIONS IMPOSED BY THE NATURE OF THE DATA**

The nature of the data imposed two major restrictions on the conduct of the study:

- a. The OMF contains the date of graduation only in the

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Entry-No.	Content
(1)	Social Security Number
(2)	Month and year of birth
(3)	Month and year of entry in the system
(4)	Year of graduation from a graduate-level program
(5)	Minimum Service Requirement (month and year)
(6)	Month and year of loss
(7)	First year eligible to retire (year)
(8)	Subspecialty-code (SSC)

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Table 2: Content of an Individual Record as Obtained From  
Department of the Navy - NMPC

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year in which this event occurred. Thus the basic time unit for this study is the calendar year.

- b. The Minimum Service Requirement for each officer as obtainable from the OMF is not kept updated. Thus no analysis is possible relating the loss of a person to that date.

The loss date as extractable from OMF/AF and as contained in the records under entry-No. six<sup>21</sup> shows

- the actual loss date when an individual was lost prior to 1980,
- no entry or an expected future loss date when an individual was not lost prior to 1980.

As "first year eligible to retire" (RY), the OMF/AF and the records structured for this study show the calculated first possible retirement year based on the 20-year limit. This date is kept on the OMF/AF even if an individual served longer.

With regard to the loss date, three assumptions had to be made:

- a. A person lost in the year recorded as RY was lost due to reaching his retirement age.
- b. A person lost after the year recorded as RY was lost due to reaching his retirement age.

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<sup>21</sup>See Table 2.

- c. A person lost during the year prior to the recorded RY was lost due to reaching his retirement age if the difference between the year an individual joined the forces and his actual loss year was found to be 20.<sup>22</sup>

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<sup>22</sup>See Table 2: Difference between entries six and three.

#### IV. ANALYSIS OF THE RAW DATA

##### A. OBSERVED SAMPLE SIZES AND DURATION OF OBSERVATIONS

As the purpose of this study is to condition the retention behavior of the observed group of officers on their respective SSC's, a first check should be devoted to the number of officers in each SSC. The listing in Appendix B reveals extreme differences in sample sizes. They range from zero in 16 cases<sup>23</sup> and one in 19 cases up to 96 for SSC xx42 for graduation year 1973. Even if the SSC's are grouped into SSC groups,<sup>24</sup> there remain considerable differences, as Table 3 shows.

As Table 3 already suggests, sample sizes vary also within the SSC's over the six subsequent graduation years observed. It was found that, as extremes, sample sizes varied for

- SSC xx42 between 45 and 96 officers,
- SSC xx44 between 0 and 20 officers, and
- SSC xx82 between 0 and 48 officers.

The given variation in group sizes and especially the high number of cases with sample sizes of zero or one officers will have an effect on the applicability of standard

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<sup>23</sup>Sample size 0: SSC xx44 and 1102 in 1970; SSC 1103 and 1308 in 1971; SSC xx44 and 1103 in 1972; SSC xx21 in 1973; SSC xx44, 1102, and 1307 in 1974; SSC xx21, xx25, xx27, and xx67 in 1975.

<sup>24</sup>See Table 1.

SSC group	1970	1971	1972	1973	1974	1975
xx2x	33	39	32	34	40	13
xx3x	106	115	107	100	101	114
xx4x	92	139	163	155	107	131
xx5x	93	89	117	96	97	87
xx6x	38	43	53	42	46	49
xx7x	29	42	49	57	49	44
xx8x	6	39	63	36	41	34
xx9x	93	119	83	78	75	56
11xx	32	43	47	36	57	51
13xx	40	47	54	35	36	39
Total	562	715	768	669	649	618

Table 3: Number of Officers Within the SSC Groups Who Were Graduated in the Years Indicated

statistical procedures in analyzing the data and in trying to apply estimation techniques to them.

The problem is intensified by the limited amount of years  $k$  after graduation for which observations were possible.

As retention rates in the first few years after graduation can be expected to be high

- the amount of variation over the observed time period can be expected to be extremely small and thus loss rates can be expected to stay close to zero.
- The influence of chance on the variation in cases where sample sizes are as small as described can be expected to have made the observed loss rates unreliable.

## B. OVERALL LOSSES AND TRENDS

### 1. Presentation of Losses Encountered

Prior to analyzing losses by subspecialties, overall losses for the total group of officers observed will be examined to get a feel for the actual magnitude of losses encountered and to find out about possible trends over time.

Table 4 shows in part a. the accumulated losses for the indicated years  $k$  after graduation where  $k$  equals one in the year of graduation.  $N$  denotes the number of officers who were graduated in the indicated calendar year. Part b. shows the equivalent accumulated loss rates as fraction of the losses over the respective base groups  $N$ .

-----  
 Grad year N 1 2 3 4 5 6 7 8 9 10  
 -----

a. accumulated number of officers lost

1970	562	0	1	3	5	9	40	74	97	126	158
1971	715	0	0	5	9	65	98	141	183	223	-
1972	768	0	0	4	47	90	127	182	214	-	-
1973	669	0	1	16	48	102	147	184	-	-	-
1974	649	0	10	17	46	108	159	-	-	-	-
1975	618	2	14	25	54	100	-	-	-	-	-

b. accumulated losses as a fraction of the base group N

1970	562	.0	.002	.005	.009	.016	.071	.132	.173	.224	.281
1971	715	.0	.0	.007	.013	.091	.137	.197	.256	.312	-
1972	768	.0	.0	.005	.061	.117	.165	.237	.279	-	-
1973	669	.0	.002	.024	.072	.153	.220	.275	-	-	-
1974	649	.0	.015	.026	.071	.166	.245	-	-	-	-
1975	618	.003	.023	.041	.087	.162	-	-	-	-	-

-----

Table 4: Accumulated losses in the graduation-year groups 1970 to 1975 for years  $k = 1, 2, \dots, 10$  after graduation ( $k = 1$  for the year of graduation).

## 2. Trend Relative to Years After Graduation

Table 4 indicates a parallel or at least similar development of the losses relative to the years k after graduation for all graduation years i. Figure II shows the unaccumulated loss rates ( $RL_{i,k}$ ) for graduation years i = 1971, 1973, and 1975 as examples. This figure also includes a curve showing the unaccumulated loss rate development over years k after graduation regardless of the graduation year ( $RTL_k$ ). It is calculated as

$$(7) \quad RTL_k = \frac{\sum_{i=1970}^m L_{i,k}}{\sum_{i=1970}^m N_{i,k}} \quad \text{for each year } k,^{25}$$

where

$L_{i,k}$  = loss of graduation-year group<sub>k</sub> in year k,

$N_i$  = size of graduation-year group<sub>i</sub>

$$m = \begin{cases} 1975 & \text{for } k = 1, \dots, 5; \\ 1979 - k + 1 & \text{for } k = 6, \dots, 10 \end{cases}$$

All curves show a rapid increase in losses up to the 5th year after graduation. After that year, no clear trend is determinable. The peak in the 5th year is explainable by the vanishing effect of the Minimum Service Requirement (MSR).

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<sup>25</sup><sub>k</sub> = 1 for the year of graduation

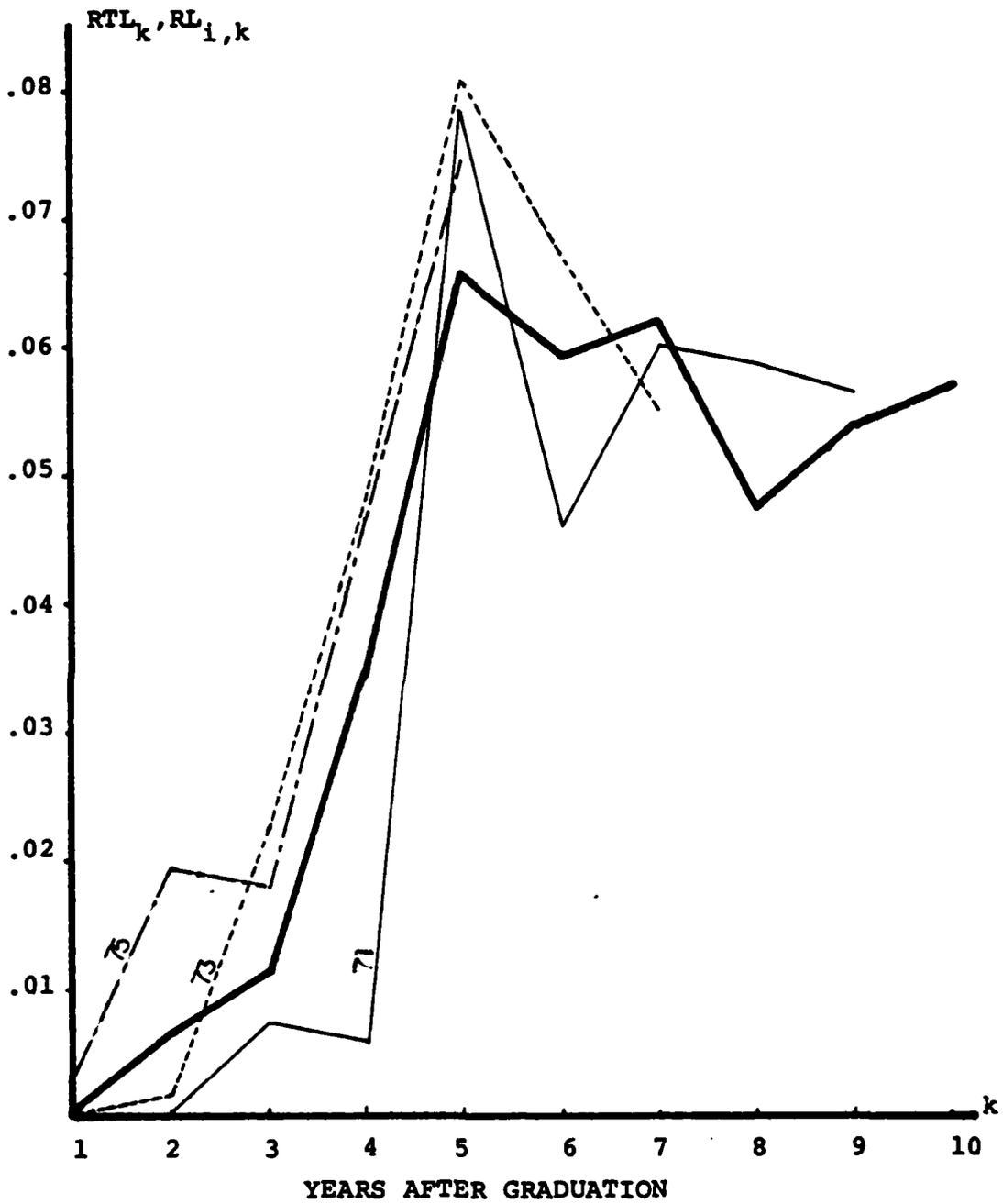


FIGURE II. Development of Unaccumulated Loss Rates Over Time ( $k = 1$  for the year of graduation). Shown are the curves for graduation years 71, 73, and 75 and as bold-faced curve the  $RTL_k$  curve according to equation (7).

Officers assigned to postgraduate education have to serve on active duty for a period of three years for the first year of education and one more year for every year thereafter upon completion of the program.<sup>26</sup> As follows from the description of the SSC's selected for this study, almost all of the officers observed can be expected to have been graduated from NPS. Most of the curricula requires MSRs of at least four years.<sup>27</sup> Thus the steep increase of losses in the fifth year could be a result of the cessation of the MSR obligation.

For further clarification of the observed development of the losses, it was found to be helpful to distinguish between

- a. observable losses due to the fact that officers reach their retirement age, denoted by  $LRO$ ,<sup>28</sup> and
- b. losses due to other reasons, denoted by  $LO$ .

The data available allowed for an extraction of the  $LRO$  and the  $LO$ . Both rates could now be calculated as respective losses over the base groups  $N_i$ . Figure III shows the unaccumulated rate for the  $LRO$ -losses, denoted by  $RLRO_k$ , and the unaccumulated rate for the  $LO$ -losses, denoted by  $RLO_k$ . All rates are calculated according to equation (7) with the appropriate loss type in the numerator.

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<sup>26</sup>OPNAVNOTE 1520, 25 June 1979, Paragraph 6

<sup>27</sup>See OPNAVNOTE 1520, Enclosure 1

<sup>28</sup>See assumptions about the date of retirement in paragraph III. B.

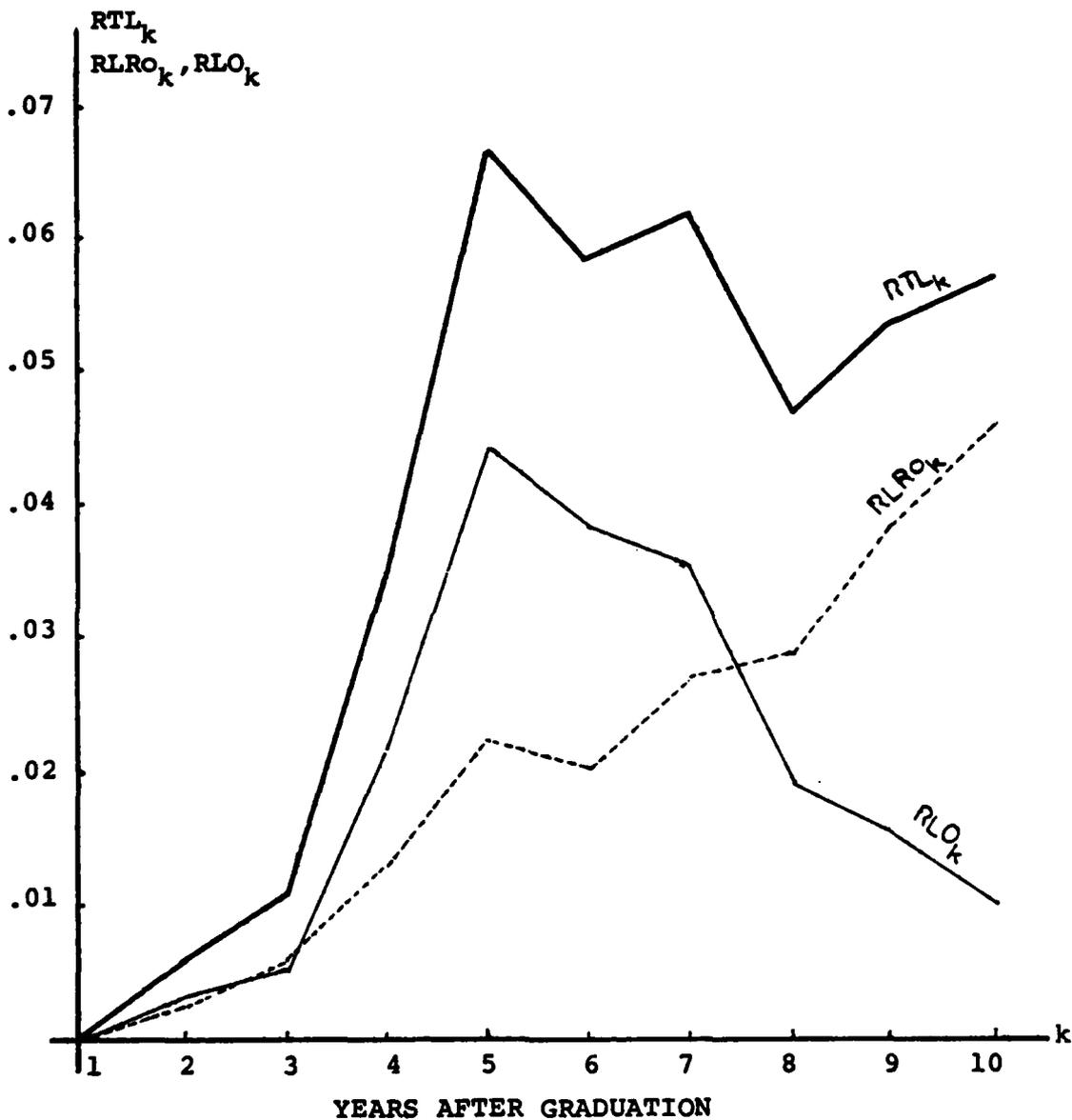


Figure III. Development of rates for total loss ( $RTL_k$ ), retirement loss ( $RLRo_k$ ) and loss due to other-than-retirement reasons ( $RLO_k$ ) over time ( $k = 1$  for the year of graduation). All rates are calculated according to equation (7) with the appropriate loss type in the numerator.

Figure III already suggests the increasing importance of retirement losses relative to total losses. This relation is emphasized in Figure IV. From the 8th year after graduation on, more than 50% of the yearly losses are accountable as retirement losses. With the 10th year, the relative importance of RLRO has reached 81%. The  $RLRO_k/RTL_k$  ratio for the year  $k = 1$  could only be calculated on the basis of two out of 3981 cases.

The trend depicted in Figures III and IV and the likelihood of a steady increase of the relative importance of RLO for  $k$  greater than ten are essentially attributable to two facts:

- a. Officers who have only a few years of service left before reaching the 20-year limit--in case this limit applies to them--assess the value of the pension obtainable high enough not to leave deliberately.
- b. With increasing distance from the graduation year, the age of the officers will naturally increase. Thus the share of officers who will have to retire will increase until it approaches 100%.

The low 34% share of retirement losses in the 5th year after graduation is a result of increased losses due to other reasons, described earlier. Three major conclusions for establishing estimation and prediction models can be drawn from the facts so far known:

- a. Both losses, the observable retirement loss (LRO) as well as the loss due to other reasons (LO), are not a linear

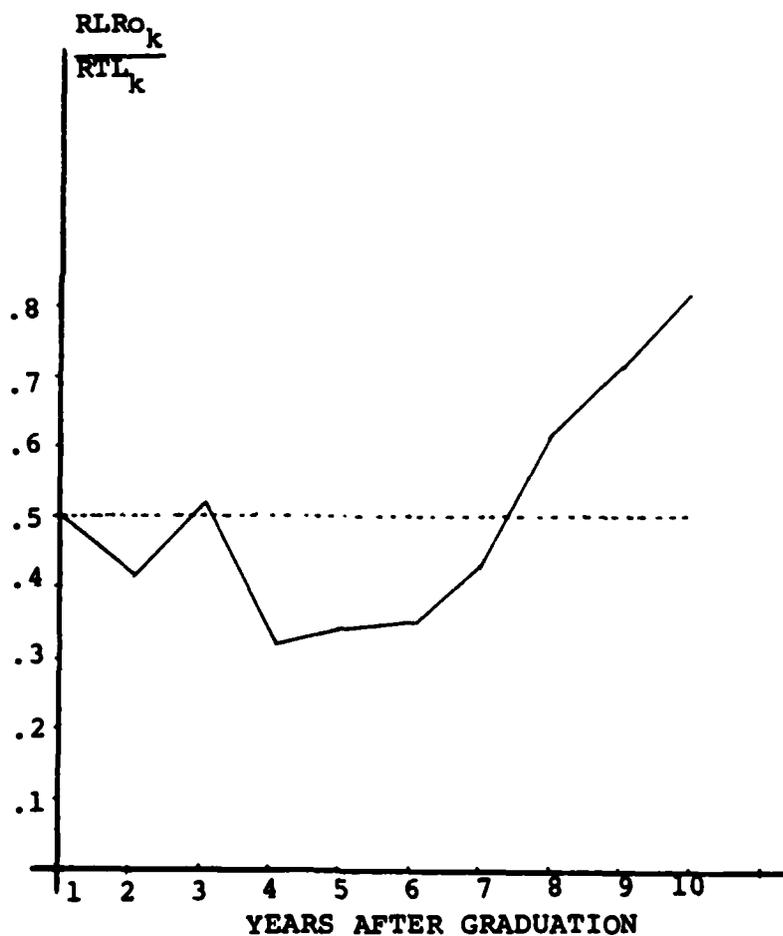


Figure IV. Development of the relative importance of retirement losses (ratio  $RLRO_k/RTL_k$ ) over time (all graduation-year groups<sup>k</sup> combined).

function of time (k) after graduation. Thus models requiring linear relationships are not applicable.

- b. The relation between total losses (L) and retirement losses (LRO) depicted in Figures III and IV suggest that estimation and prediction efforts have to be concentrated on the first few years after graduation.
- c. As losses of officers which occur because those officers retire are known deterministically, they do not have to be estimated. They are given.<sup>29</sup>

### 3. Trend Relative to the Year of Graduation

With respect to trends relative to year i of graduation, Table 4 indicates a generally increasing trend in the total accumulated losses observed.

In observing retention behavior of regular Marine Corps Officers, McAfee<sup>30</sup> found that the probability that an officer still stays in the system k years after entrance is the same independent of the entrance-year group the officer belonged to. Based on this assumption, which he verified with tests of homogeneity, he was able to construct a classic prediction model for retention rates k years after entrance. The probability he assumed to be stationary corresponds to the complement of the accumulated total loss rate of this study.

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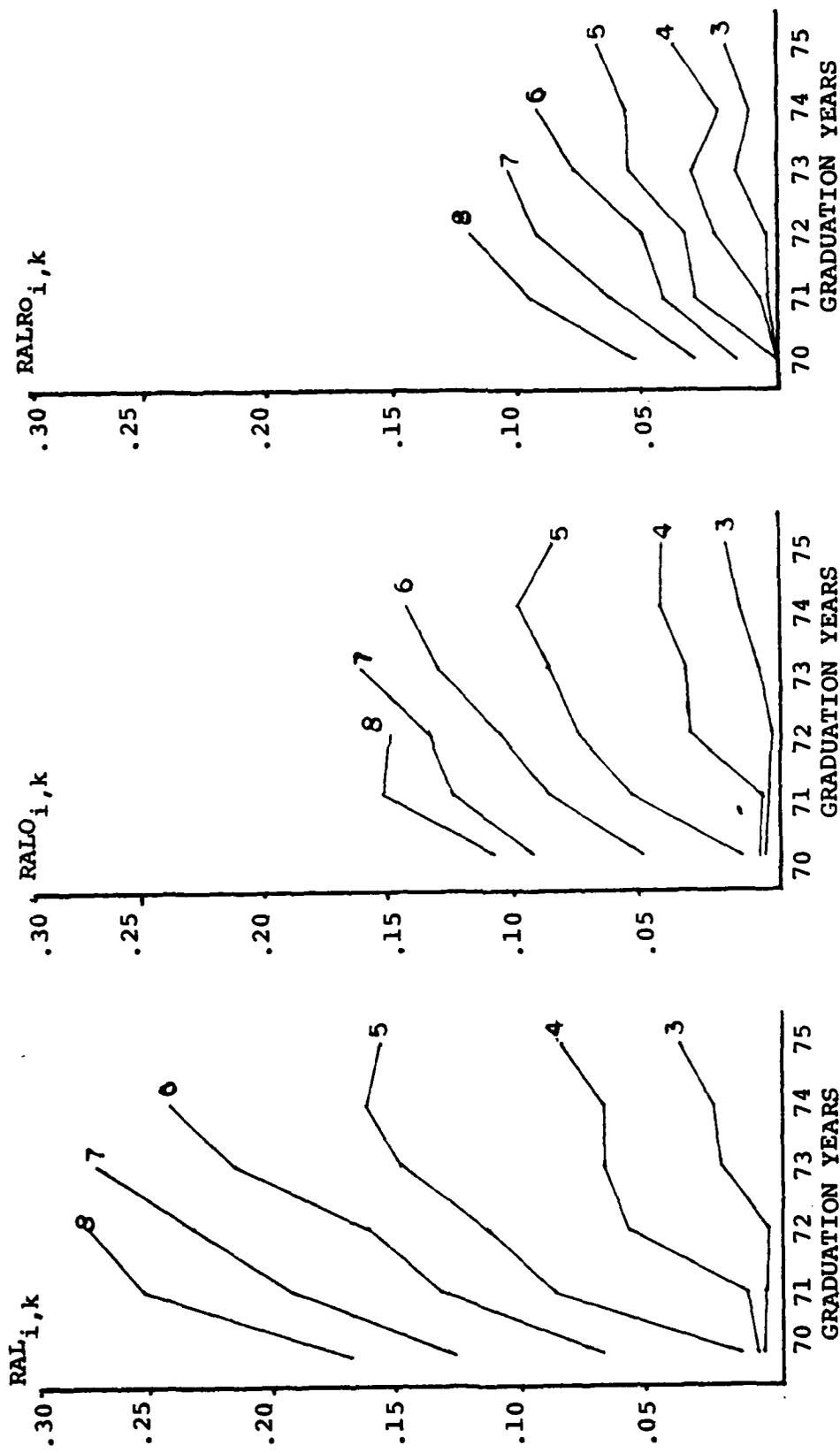
<sup>29</sup>The retirement loss observed (LRO) is not identical with the retirement loss which could be calculated once a group of officers has entered the system. Part of the group will be lost due to other reasons before those officers reach their retirement age. Implications will be shown later.

<sup>30</sup>McAfee, C. K., A Cohort Model for Predicting Retention of Regular Marine Corps Officers, MS Thesis NPS, Monterey, 1970.

The data underlying this study do not suggest stationarity for the accumulated total losses with respect to subsequent graduation-year groups. For the years after graduation  $k = 3, 4, 6, 7$  and  $8$ , accumulated total loss rates ( $RAL_{i,k}$ ) are increasing over the whole range of subsequent graduation years  $i$ . Only for  $k = 5$  do the  $RAL_{i,k}$  decrease from 1974 as graduation year  $i$  on. Accumulated loss rates due to other-than-retirement reasons ( $RALO_{i,k}$ ), as one part of the total losses, show the same trend for  $k = 3, 5, 6$ , and  $7$ . However, the increase occurs with a smaller rate of change and the decrease for  $k = 5$  from graduation year 1974 on is more obvious. For  $k = 4$ , and  $8$ ,  $RALO_{i,k}$  decrease in the last respectively observable graduation year, too. The trend for the accumulated retirement losses ( $RALRO_{i,k}$ ), as the second group of the total losses, is increasing for all years  $k$  after graduation over the range of the subsequent graduation years  $i$ .

Figure V shows the above described developments for  $RAL_{i,k}$  in part a., for  $RALO_{i,k}$  in part b., and for  $RALRO_{i,k}$  in part c. for years after graduation  $k = 3, 4, 5, 6, 7$ , and  $8$  over the six graduation years.

Prior to carrying out statistical tests to determine the significance of the differences among the six independent graduation-year groups for various years  $k$  after graduation, no certain conclusions about the stationary at least of the unaccumulated losses are possible. However, two summarizing remarks can be made:



a. Accumulated total loss rates  
 b. Accumulated loss rates due to other-than-retirement reasons  
 c. Retirement-loss rates

FIGURE V. Development of loss rates over graduation years for selected years after graduation  $k = 3, 4, 5, 6, 7, \text{ and } 8$ .

- a. Accumulated losses are generally increasing for all observed years  $k$  after graduation with increasing years of graduation.
- b. The fact that accumulated retirement losses ( $ALRo_{i,k}$ ) follow this increasing trend too might be caused by two reasons:
  - Postgraduate education starts progressively later in the career of Navy officers. Thus, with increasing graduation year, the average age of the graduates is higher.
  - Progressively fewer P-coded officers pass the 20-year limit. Thus, with increasing graduation year, more officers retire after 20 years of service.

#### 4. Hypothesis Testing for Stationarity

As stated in paragraph IV.B.2, retirement losses are assumed to be known. They do not have to be estimated.

Thus, in testing the significance of differences between the six graduation-year groups, the main concern has to be with the losses for other-than-retirement reasons. Tests will have to be performed for

- accumulated losses due to other-than-retirement reasons (ALO), and
- unaccumulated losses due to other-than-retirement reasons (LO).

For this purpose, Chi-square tests for independent samples have

been chosen. The tests were applied to the data according to the procedure suggested by Siegel.<sup>31</sup>

For each year  $k$  after graduation, there is a contingency table constructable with rows representing the graduation-year groups and two columns representing the category of officers who were lost versus the category of officers who were still in service. The requirement for each of the cells within the contingency table to obtain an expected frequency of five or more for more than 80% of the cells is fulfilled for  $k > 3$ . Tests were carried out for  $k = 4, 5, 6,$  and  $7$ .

Two hypotheses with the following  $H_0$  were tested:

- a.  $H_0$ : The graduation-year groups do not differ with respect to ALO for a given year  $k$  after graduation.
- b.  $H_0$ : The graduation-year groups do not differ with respect to LO for a given year  $k$  after graduation.

As was to be expected for the accumulated loss case (ALO),  $H_0$  was rejected for all  $k$  at a level of significance less than 0.00001. However, leaving off graduation-year groups 1970 and 1971 led to an acceptance of  $H_0$  at a level of significance bigger than 0.1. That means for the latter case that if  $H_0$  were rejected the error probability would be bigger than 0.1. For the unaccumulated case (LO),  $H_0$  was rejected for  $k = 4$  and  $k = 5$  at a level of significance of less than 0.0001. However

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<sup>31</sup> Siegel, S., Nonparametric Statistics for the Behavioral Sciences, McGraw-Hill, New York, 1956.

$H_0$  was accepted for  $k = 6$  and  $k = 7$  at a significance level bigger than 0.1.

The results of the Chi-square tests seem to indicate that the differences between the graduation-year groups with respect to ALO and LO are not significant at the significance levels stated when only graduation years after 1971 are considered.

However, especially when for the ALO case the development for  $k = 6$  is considered, as shown in Figure V, then the result of the Chi-square test to accept the independence hypothesis should not be valued too highly. The reason for the discrepancy between reality and test results appears to be twofold: On the one hand, the total number of cases is inflated<sup>32</sup> in comparison to the actual frequencies in the loss category of the contingency table. According to Siegel, this invalidates the test. On the other hand, the test is not sensitive to trends. This can be easily recognized from the construction of the statistic.<sup>33</sup>

Using as a line fitting procedure the linear regression for  $RALO_{i,k}$  for  $k = 6$  and using the five graduation years  $i = 1970, \dots, 1974$   $X = 1, 2, \dots, 5$  as the carriers the fitted line equation is

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<sup>32</sup>See Siegel, S., p. 109.

<sup>33</sup>See Siegel, S., p. 175.

$$RALO_{i,6} = 0.0390 + 0.0234x \quad \text{with } R^2 = 0.968.^{34}$$

The positive slope of the line shows the increasing trend already recognized in Figure V. The following example illuminates the magnitude of the losses: Given that 600 officers graduated each year, then the losses due to other-than-retirement reasons up to the 6th year after graduation were:

- 37 officers of the first graduation-year group,
- 80 officers of the 4th graduation-year group, and  
extrapolating
- 164 officers of the 10th graduation-year group.

This trend is certainly significant.

The findings with respect to the development of losses relative to the year of graduation allow essentially two major conclusions:

- a. Any model construction for estimation and prediction of losses of P-coded officers conditioned on the graduation-year group the officer belongs to that is based on the general assumption that losses occur independently from the graduation year does not comply with reality.
- b. The trends found for accumulated losses up to and including a year k after graduation over sequential graduation years are quantifiable. Using simple linear regression with graduation years as the independent

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<sup>34</sup>Values for  $RALO_{i,6}$  for graduation years  $i = 1970, \dots, 1974$  are contained in Tables 3(D) of Appendix D.

variable seems to be a starting point to predict accumulated losses in future years  $k$  after graduation for officers belonging to a specified graduation-year group.

V. ESTIMATION OF LOSSES CONDITIONED ON GRADUATION-YEARS AND  
SUBSPECIALTY CODES

A. DERIVATION OF THE MODEL

In Chapter II the Bayes Estimator that will be used in estimating loss rates was derived. In the previous chapter it was shown that it is possible to exclude retirement losses from the application of the Bayesian estimation technique.

As it is desired to estimate losses conditioned on the SSC an officer has obtained, it was necessary to extract the accumulated observable retirement losses (ALRO), the accumulated losses due to other reasons (ALO), and the accumulated total losses (AL) that occur up to and including the k-th year after graduation

- for each graduation-year group and within those
- for each of the 41 SSC's.

Thus the numeric losses  $AL_{i,j,k}$ ,  $ALO_{i,j,k}$ , and  $ALRO_{i,j,k}$  were obtained, where the subscripts

- $i = 70, \dots, 75$  stand for the graduation-year observed
- $j = 21, \dots, 1308$  stand for the 41 SSC's included, and
- $k = 1, 2, \dots, k_{79}$  stand for the years k after graduation with  $k_{79}$  being the k reached in calendar year 79.

In the following derivation of the model the subscripts will be omitted and it is understood that the losses refer to a certain graduation-year i, a certain year k after graduation, and a certain SSC j. Assuming now that ALO and ALRO are

mutually exclusive, then

$$(8) \quad AL = ALO + ALRo \quad \text{for each } i, j, k.$$

As was already mentioned in paragraph 4.2.2. ALRo is not identical with the retirement loss (ALR) obtainable by calculating from the personal data of each individual the year k after graduation when he would reach his retirement age. The relation between ALR and ALRo can be described by

$$(9) \quad ALRo = ALR - (ALO \cap ALR)$$

where  $(ALO \cap ALR)$  is the retirement loss that would have been observed in addition to ALRo for year k after graduation had some officers not been lost due to other-than-retirement reasons prior to reaching their retirement age in year k after graduation. Thus equation (8) becomes

$$(10) \quad AL = ALO + ALR - (ALO \cap ALR).$$

Calculating for each  $i, j, k$  total accumulated loss rates ( $r$ ) from the left side of equation (8) yields

$$(11) \quad r = \frac{AL}{N}$$

where  $N$  is the number of graduates belonging to graduation year  $i$  and to SSC  $j$ .

Calculating  $r$  from the right side of equation (8) yields

$$(12) \quad r = \frac{ALO + ALRo}{N}$$

which can be algebraically manipulated and rewritten to

$$(13) \quad r = \frac{\frac{ALO}{N} \times N + \frac{ALRo}{(N - ALO)} \times (N - ALO)}{N}$$

Equation (13) gives a basis for application of the Bayes Estimation Technique to the losses due to other-than-retirement reasons (ALO) and for including external given retirement losses.<sup>35</sup>

Let

$$r_{\text{ALO}} = \frac{\text{ALO}}{N} \quad \text{and}$$

$$r_{\text{ALR}} = \frac{\text{ALR}_0}{(N - \text{ALO})}$$

then equation (13) can be rewritten as

$$(14) \quad r = r_{\text{ALO}} + r_{\text{ALR}} (1 - r_{\text{ALO}})$$

Assuming independence between retirement losses and losses due to other-than-retirement reasons, equation (14) is the probability statement corresponding to equation (10), where

$r = P$  (an officer belonging to  $i$  and  $j$  is lost by year  $k$ ) and  $r_{\text{ALO}}$  and  $r_{\text{ALR}}$  are the corresponding probabilities for the loss reasons described by ALO and ALR.

An alternative way to reach the result shown in equation (14) by making use of conditional probabilities is shown in Appendix C. Now, instead of the actual observed  $r_{\text{ALO}}$  the Bayesian estimate  $b_{\text{ALO}}$  will be used. Thus equation (14) can be rewritten to obtain the equation for the estimate (b) for

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<sup>35</sup> For this study ALR were not given externally, but they are calculable by applying  $r_{\text{ALR}}$  to  $N$ .

the accumulated total losses:

$$(15) \quad b = b_{\text{ALO}} + r_{\text{ALR}} (1 - b_{\text{ALO}})$$

where the estimates and rates have to be calculated for each graduation year  $i$  and within it for each SSC  $j$  to get the loss estimate for each observable year  $k$  after graduation.

#### B. DERIVATION OF THE ESTIMATOR FOR THE MODEL

According to the definition of paragraph II.C. the estimate  $b_{\text{ALO}}$  will now be derived and combined with the retirement loss  $r_{\text{ALR}}$  as stated in equation (15) of the previous chapter.

Let

- $m_{i,h}$  denote the number of officers belonging to Subspecialty-code group  $h$ <sup>36</sup> having graduated in graduation year  $i$ ;
- $n_{i,j}$  denote the number of officers belonging to SSC  $j$ , where  $j$  is one of the SSC's belonging to SSC group  $h$  who were graduated in the same year  $i$ ;
- $z_{i,h,k}$  denote the number of officers belonging to  $m_{i,h}$  who have been lost up to and including year  $k$  after graduation for other than retirement reasons; and let
- $y_{i,j,k}$  denote the number of officers belonging to  $n_{i,j}$  who have been lost up to and including year  $k$

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<sup>36</sup>There are 10 SSC groups  $h$  as specified in paragraph IV.A.

after graduation for other than retirement reasons.

The Bayesian estimate for the loss rate  $r_{\text{ALO};i,j,k}$ , where

$$r_{\text{ALO};i,j,k} = \frac{y_{i,j,k}}{n_{i,j}}$$

for each  $i, j, k$  as specified in paragraph V.A will be calculated according to equation (6) in paragraph II.C as

$$(16) \quad b_{\text{ALO};i,j,k} = \frac{z_{i,j,k} + y_{i,j,k}}{m_{i,h} n_{i,j}}$$

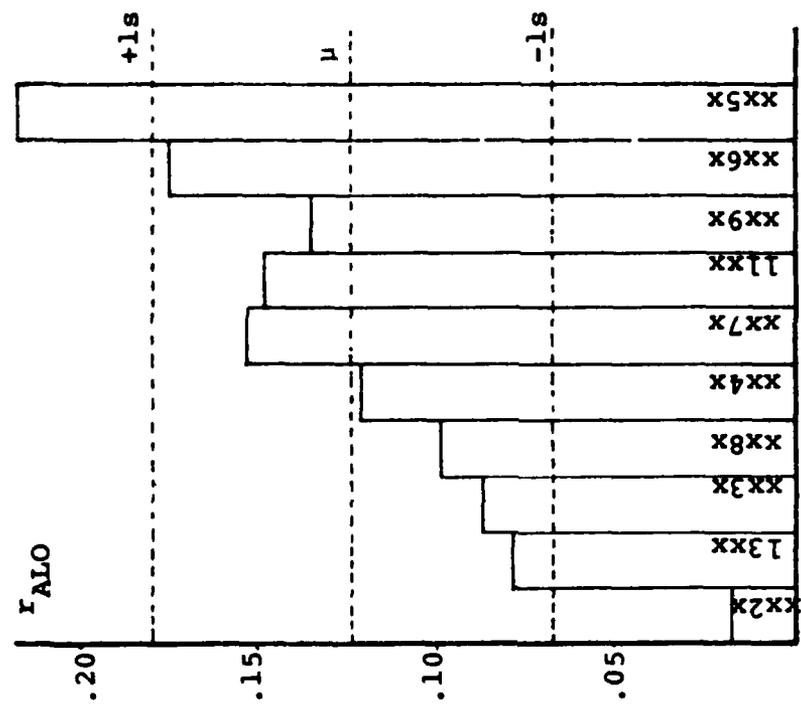
for each  $i, j, k$  as specified in paragraph V.A and each  $h$  as specified above. The calculation of the estimate  $b_{\text{ALO}}$  is possible as long as  $m_{i,h}$  is not empty.<sup>37</sup>

The construction of the Bayesian estimate, as explained in Chapter II, would have allowed for including the total group of graduates of a graduation year  $i$  as a first refining sample into the estimator. However, this would have the undesirable effect that this inflated sample would almost completely cover up basic differences in the retention behavior between officers belonging to different SSC groups and to different SSC's. The existence of differences among the SSC groups is shown in Figure VI.

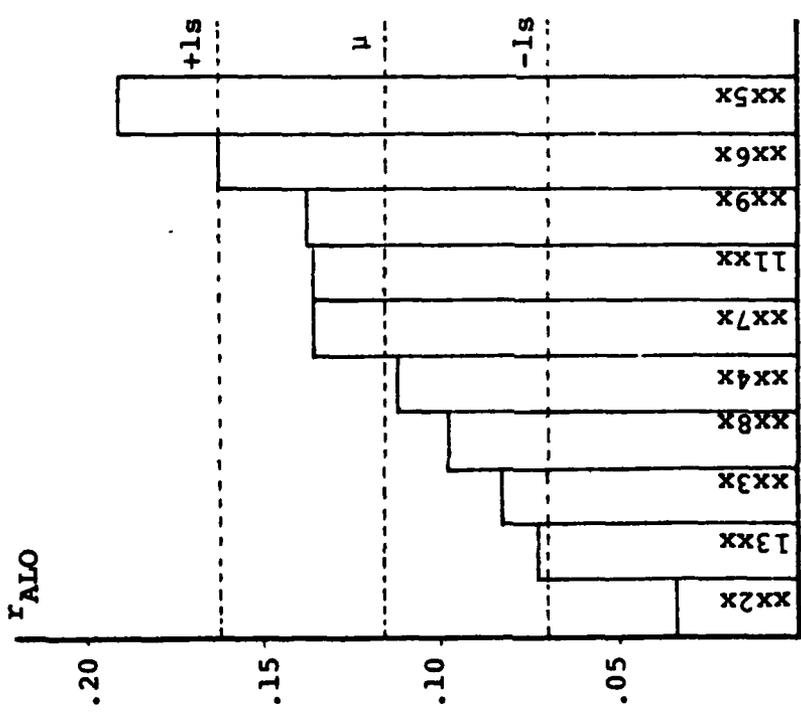
The first histogram shows the accumulated losses due to other-than-retirement reasons (ALO) up to year  $k = 5$  as average over graduation years 1972 to 1975. The SSC groups

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<sup>37</sup>For the six graduation years ( $i = 1970, \dots, 1975$ ) and for all SSC groups ( $h = 1, \dots, 10$ )  $m_{i,h}$  was found not to be empty.



SSC GROUPS  
 $r_{ALO}$  for  $k = 6$  as average over graduation years 72 to 74



SSC GROUPS  
 $r_{ALO}$  for  $k = 5$  as average over graduation years 72 to 75

Figure VI: Accumulated losses due to other-than-retirement reasons within the Subspecialty-code groups up to and including the indicated year  $k$  after graduation as average over the indicated graduation years.

have been ordered in terms of ascending loss rates. The second histogram shows the same losses up to year  $k = 6$  as averages over graduation years 1972 to 1974.<sup>38</sup> The order of the SSC groups remained the same as in the first histogram. In addition, the histograms show the mean losses and lines bounding the region of  $\pm 1$  standard deviation  $s$ . Besides minor changes in the order of those groups with losses near the mean ( $\mu$ ), the ranking in terms of ascending loss rates stays stable. Groups xx2x and xx5x lie outside the region of  $\pm 1 s$  and groups l3xx and x6xx lie close to the border lines of  $\pm 1 s$ .

In order to test the significance of the differences among the SSC groups, again Chi-square tests were applied as introduced in paragraph IV.B.3.

The hypotheses

$H_0$ : The SSC groups do not differ with respect to their ALO rates for  $k = 5$ ;

$H_0$ : The SSC groups do not differ with respect to their ALO rates for  $k = 6$ ;

were both rejected at a significance level of less than 0.00001.

For the calculation of the retirement losses  $r_{ALR;i,j,k}$ <sup>39</sup> which is the second entity needed to estimate the overall loss-rate estimate  $b_{i,j,k}$  two assumptions have to be made:

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<sup>38</sup> Last year of observation was year 1979. There were no data available for  $k = 6$  for graduation year 1975 (corresponds to calendar year 1980).

<sup>39</sup> Subscripts have been added to  $r_{ALR}$  as specified above.

Let

$u_{i,j,k}$  denote the number of officers belonging to SSC  $j$  of graduation year  $i$  who retired up to and including year  $k$  after graduation.<sup>40</sup>

Then, as was shown in paragraph V.A.

$$(17) \quad r_{ALR;i,j,k} = \frac{u_{i,j,k}}{n_{i,j} - y_{i,j,k}}$$

for any  $i, j, k$  as specified in chapter V.A.

Now, if  $n_{i,j}$  equals  $y_{i,j,k}$  for any  $k$  than  $u_{i,j,k}$  will be zero and thus  $r_{ALR;i,j,k}$  has to be assumed to be zero for the SSC  $j$  involved. This means that if everyone belonging to  $n_{i,j}$  was lost up to year  $k$  after graduation due to other-than-retirement reasons, nobody of that group could have retired.

The second assumption is necessary for the case that nobody belonging to a certain SSC  $j$  graduated in a certain graduation year  $i$ . In this case  $n_{i,j}$  is zero for that SSC  $j$  in the specific year  $i$  and  $u_{i,j,k}$  and  $y_{i,j,k}$  are zero for all  $k$  and for the specific  $j$  and  $i$  under consideration. However, the Bayesian estimator allows for calculating the estimate  $b_{ALO}$  as long as at least  $m_{i,h}$  is not zero too, as was shown earlier. For calculating the retirement rate, it was assumed that if  $n_{i,j}$  were empty it would have the same retirement rate as the SSC group  $m_{i,h}$  it belongs to.<sup>41</sup>

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<sup>40</sup> $u_{i,j,k}$  corresponds to ALRo for the graduation year  $i$ , the SSC  $j$  and the year  $k$  after graduation under consideration.

<sup>41</sup>This method was mentioned by Prof. R. Weitzman in his lectures at NPS and is known as "Incomplete Tree Method."

## C. PRESENTATION AND DISCUSSION OF RESULTS

### 1. Presentation of Results

Appendix D shows in six sections--one section for each graduation year--loss rates and estimates of the rates conditioned on the graduation year and the Subspecialty code an officer has obtained.

The graduation-year sections contain the following tables:

- Table 1(D): Accumulated total loss rates ( $r_{i,j,k}$ ) and their respective estimates ( $b_{i,j,k}$ ) for each SSC  $j$  and each year  $k$  after graduation for the graduation year  $i$  of that section.
- Table 2(D): Accumulated loss rates for other-than-retirement losses ( $r_{ALO;i,j,k}$ ) and their estimates ( $b_{ALO;i,j,k}$ ) for the same  $i$ ,  $j$ , and  $k$  as in Table 1(D).
- Table 3(D):  $r_{i,h,k}$ ,  $r_{ALO;i,j,k}$  and  $r_{ALR;i,h,k}$  for the ten SSC groups  $h$ ;  $RAL_{i,k}$ ,  $RALO_{i,k}$ , and  $RALR_{i,k}$  for the total group of graduates in that graduation year  $i$  and for all years  $k$  after graduation.

### 2. Discussion of Results for Selected Subspecialty Codes

As it seemed impossible to exhaustively discuss the results for all 41 Subspecialty codes involved, three of the ten SSC groups with their Subspecialty codes were selected according to their special appearance in Figure VI of Chapter

V.B. These are:

- SSC group xx2x which shows low loss rates and extremely small sample sizes of the SSC involved,
- SSC group xx5x which shows the highest loss rates and also high sample sizes of the SSC belonging to it, and
- SSC group xx4x which shows loss rates closest to the mean loss rate of the ten SSC groups and which also shows high sample sizes of the SSC involved.

a. Subspecialty Code Group xx2x

SSC group xx2x seems to be one of the groups that caused the rejection of the equivalence tests in paragraph V.B. It is the group with the lowest loss rates and with the lowest sample sizes of the SSC's belonging to the group.

Table 5 shows the magnitude of the actual accumulated losses due to other-than-retirement reasons (ALO) for graduation years 1972 to 1975 up to and including the 5th year, the 6th year, and the 7th year after graduation (k equals one in the year of graduation).

A total of four officers out of 119 graduates were lost up to and including the 5th year after graduation and only two out of 106 graduates were lost up to and including year six after graduation.

These extremely small losses do not allow for recognition of any trends over the graduation years or differences among the Subspecialty codes. Table 6 shows the group sizes of all SSC's in group xx2x, and Table 7 shows the accumulated

Graduation Year	Size of group xx2x	Officers lost (ALO) up to and including			from SSC	SSC size
		k = 5	6	7		
1972	32	1	1	1	xx21	3
1973	34	0	0	0	----	-
1974	40	1	1	-	xx22	10
1975	13	2	-	-	xx24	7

Table 5: Accumulated losses due to other-than-retirement reasons (ALO) in SSC group xx2x for k = 5, 6, 7. (Last year of observation is 1979. k = 1 in the year of graduation.)

Grad year	SSC						
	xx21	xx22	xx23	xx24	xx25	xx26	xx27
1972	3	9	3	5	6	5	1
1973	0	11	4	11	2	4	2
1974	3	10	3	15	1	7	1
1975	0	4	1	7	0	1	0

Table 6: Sample sizes of Subspecialty codes in group xx2x for graduation years 1972 to 1975.

SSC

## Graduation years

		1972		1973		1974		1975	
		ALO	AL	ALO	AL	ALO	AL	ALO	AL
xx21	r	.333	.333	----	----	.0	.0	----	----
	b	.057	.057	.0	.059	.023	.023	.154	.308
xx22	r	.0	.0	.0	.0	.100	.200	.0	.0
	b	.024	.024	.0	.0	.040	.147	.118	.118
xx23	r	.0	.0	.0	.250	.0	.0	.0	.0
	b	.029	.029	.0	.250	.023	.023	.143	.143
xx24	r	.0	.0	.0	.0	.0	.0	.286	.571
	b	.027	.027	.0	.0	.018	.018	.200	.520
xx25	r	.0	.0	.0	.0	.0	1.0	----	----
	b	.026	.026	.0	.0	.024	1.0	.154	.308
xx26	r	.0	.200	.0	.0	.0	.0	.0	.0
	b	.027	.222	.0	.0	.021	.021	.143	.143
xx27	r	.0	.0	.0	.500	.0	.0	----	----
	b	.030	.030	.0	.500	.024	.024	.154	.308

Table 7: Accumulated loss rates due to other-than-retirement reasons ( $r_{ALO}$ ) and their estimates ( $b_{ALO}$ ) and total accumulated loss rates ( $r_{AL}$ ) and their estimates ( $b_{AL}$ ) for  $k = 5$ .

loss rates due to other-than-retirement reasons ( $r_{ALO}$ ) and their estimates ( $b_{ALO}$ ) as well as the accumulated total loss rates ( $r$ ) and their estimates ( $b$ ) for the fifth year after graduation.

As was desired, the estimate rows do not show empty cells. Instead, they show group averages according to the construction of the Bayesian estimator. The smoothing effect of the estimation technique is due to the fact that the estimate is a compromise between the information gained from the group behavior and the information gained from the behavior within a specific SSC.<sup>42</sup> The effect of chance influences on extremely small samples like the SSC's in this group is clearly visible in the case of SSC xx24. Up to the 5th year after graduation there is no loss in graduation-year groups 1972 to 1974. Graduation-year group 1975, however, not only happens to be smaller than the two previous ones, but this group loses two officers for other-than-retirement reasons and two more retire. Thus, more than 50% of this group are lost within the first five years after graduation. Looking at the loss rates only, this is an alarming increase. As the total SSC group xx2x consists of only 13 officers for that graduation year, the Bayesian estimate fails to remove this obvious chance effect significantly. However, as part

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<sup>42</sup>See paragraph II.C.1.

of the chance influences are caused by the retirement losses, chance effects are less dramatic for the  $b_{ALO}$  values. This supports the idea already mentioned in the context of analyzing general trends in paragraph IV.B.2 to base prediction models on the  $b_{ALO}$  values and insert known future retirement losses in the way described by the model derived in paragraph V.A.

b. Subspecialty-code group xx5x

SSC group xx5x is the group with the highest loss rates for  $k = 5$  as well as for  $k = 6$  in Figure VI. Together with groups xx3x and xx4x, it is also a group with high group sizes  $m_{i,h}$  relative to the other groups. Its losses--as shown in Figure VI--exceed the region bounded by  $\pm 1$  standard deviation significantly.

Table 8 shows the sizes  $n_{i,j}$  of the SSC's in this group for graduation years 1972 to 1975.

SSC					
Grad year	xx51	xx52	xx54	xx55	xx56
1972	11	19	54	27	6
1973	11	15	38	26	6
1974	9	18	28	31	11
1975	12	17	33	18	7

Table 8: Sample sizes of Subspecialty codes in group xx5x for graduation years 1972 to 1975.

Within this group, no extreme low SSC sizes are depicted. However the SSC's vary significantly in size. Table 9 shows again the accumulated loss rates  $r_{ALO}$  and  $r$  as well as their estimates  $b_{ALO}$  and  $b$ .

SSC	Graduation-years								
	1972		1973		1974		1975		
	ALO	AL	ALO	AL	ALO	AL	ALO	AL	
xx51	r	.0	.0	.0	.0	.222	.222	.083	.083
	b	.102	.102	.112	.112	.142	.142	.091	.091
xx52	r	.105	.105	.267	.267	.056	.056	.118	.118
	b	.110	.110	.144	.144	.122	.122	.096	.096
xx54	r	.204	.222	.184	.237	.143	.286	.121	.303
	b	.140	.160	.142	.197	.136	.280	.100	.286
xx55	r	.0	.037	.038	.192	.194	.290	.056	.222
	b	.090	.124	.107	.250	.148	.251	.086	.247
xx56	r	.0	.333	.0	.0	.0	.091	.0	.143
	b	.106	.404	.118	.118	.120	.200	.085	.216

Table 9: Accumulated loss rates due to other-than-retirement reasons ( $r_{ALO}$ ) and their estimates ( $b_{ALO}$ ) and total accumulated loss rates ( $r_{AL}$ ) and their estimates ( $b_{AL}$ ) for  $k = 5$ .

Again the "smoothing" effect of the estimation technique can be seen. And again the idea is supported that

prediction models should be based on the  $b_{ALO}$  values as carriers rather than on the  $b$  values: For SSC xx54, which is the SSC with the highest sizes  $n_{i,j}$ , the  $b$  values indicate a steady increase of losses from graduation year 1972 to 1975. However, looking at the  $b_{ALO}$  values, the rates  $r_{ALO}$  and their estimates  $b_{ALO}$  decrease. Thus, only the relative number of officers who had to retire increased for the subsequent graduation year groups.<sup>43</sup> For SSC xx56, retirement losses are the only cause of variation over the subsequent graduation years.

c. Subspecialty-code group xx4x

SSC-group xx4x is the group for which the accumulated loss rates for  $k = 5$  and  $k = 6$  shown in Figure VI lie closest to the mean rates of the ten Subspecialty-code groups. It is also the group with the highest group sizes  $m_{i,h}$  between graduation years 1971 and 1975. Table 10 shows the sizes  $n_{i,j}$  of the SSC in this group for graduation years 1972 to 1975.

SSC xx4x not only shows the most extreme differences in the SSC sizes but also contains the SSC's with the most extreme differences in terms of curricula at the NPS as Table 11 shows. However, these differences do not cause correspondingly great differences among the loss rates of the various SSC's within the SSC group xx4x, as Table 12 shows.

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<sup>43</sup>The fact that retirement losses show an increasing trend with increasing graduation-year was already shown in paragraph IV.B.2.

SSC				
Grad year	xx42	xx44	xx48	xx49
1972	85	0	34	44
1973	96	0	23	36
1974	54	0	25	28
1975	57	20	30	24

Table 10: Sample sizes of Subspecialty codes in group xx4x for graduation years 1972 to 1975.

SSC	Subspecialty Title	Curriculum
xx42	Operations Research/Systems Analysis	360
xx44	Anti-Submarine Warfare Systems Technology	525
xx48	Meteorology	372
xx49	Oceanography	440

Table 11: Subspecialty codes with their title and Curricula at NPS  
Source: OPNAVNOTE 1520, 25 June 1979

SSC	Graduation years								
	1972		1973		1974		1975		
	ALO	AL	ALO	AL	ALO	AL	ALO	AL	
xx42	r	.071	.094	.063	.135	.148	.167	.140	.228
	b	.089	.112	.064	.137	.118	.137	.101	.193
xx44	r	----	----	----	----	----	----	.0	.0
	b	.098	.123	.065	.129	.103	.121	.073	.073
xx48	r	.176	.206	.043	.130	.080	.120	.067	.100
	b	.112	.143	.062	.147	.098	.138	.081	.114
xx49	r	.091	.114	.083	.111	.036	.036	.042	.083
	b	.097	.119	.068	.096	.089	.089	.077	.118

Table 12: Accumulated loss rates due to other-than-retirement reasons ( $r_{ALO}$ ) and their estimates ( $b_{ALO}$ ) and total accumulated loss rates ( $r_{AL}$ ) and their estimates ( $b_{AL}$ ) for  $k = 5$ .

Figures VII and VIII show the differences in the variation of the loss rates and their estimates between group xx5x and group xx4x for  $k = 5$  and graduation years 1972 to 1975. Figure VII shows for each graduation year the mean of the  $r_{ALO}$  for the SSC's in group xx4x and group xx5x and the region bounded by  $\pm 1$  standard-deviation. Figure VIII shows the same for the respective  $b_{ALO}$ . In both cases it can be seen that--except for graduation year 1975--the standard deviations are bigger for group xx5x relative to group xx4x.

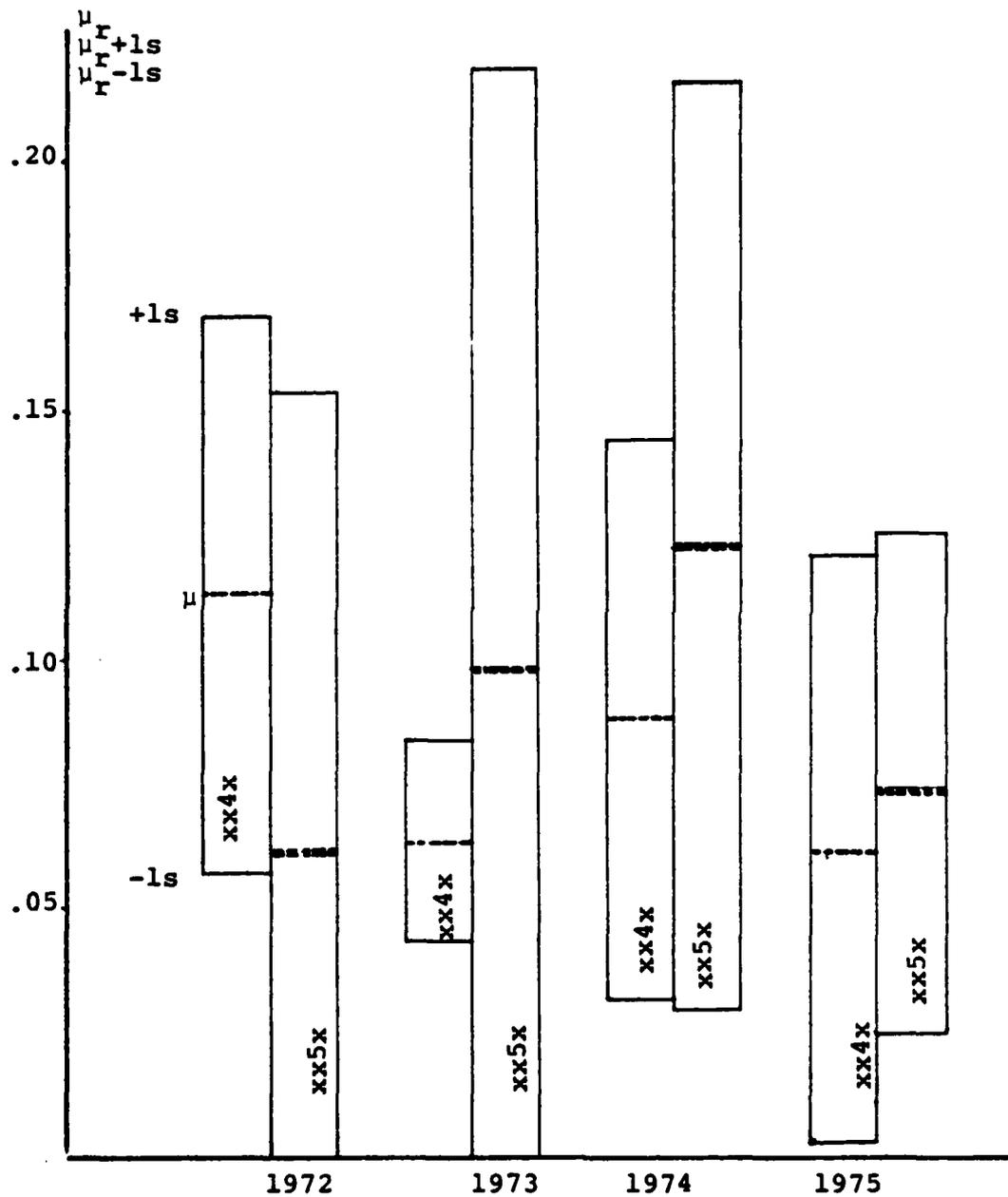


Figure VII: Means ( $\mu$ ) of the  $r_{ALO}$  of the SSC's in SSC groups xx4x and xx5x and the region bounded by +1 standard deviation (s) for  $k = 5$  years after graduation. Shown are graduation-year groups 1972 to 1975.

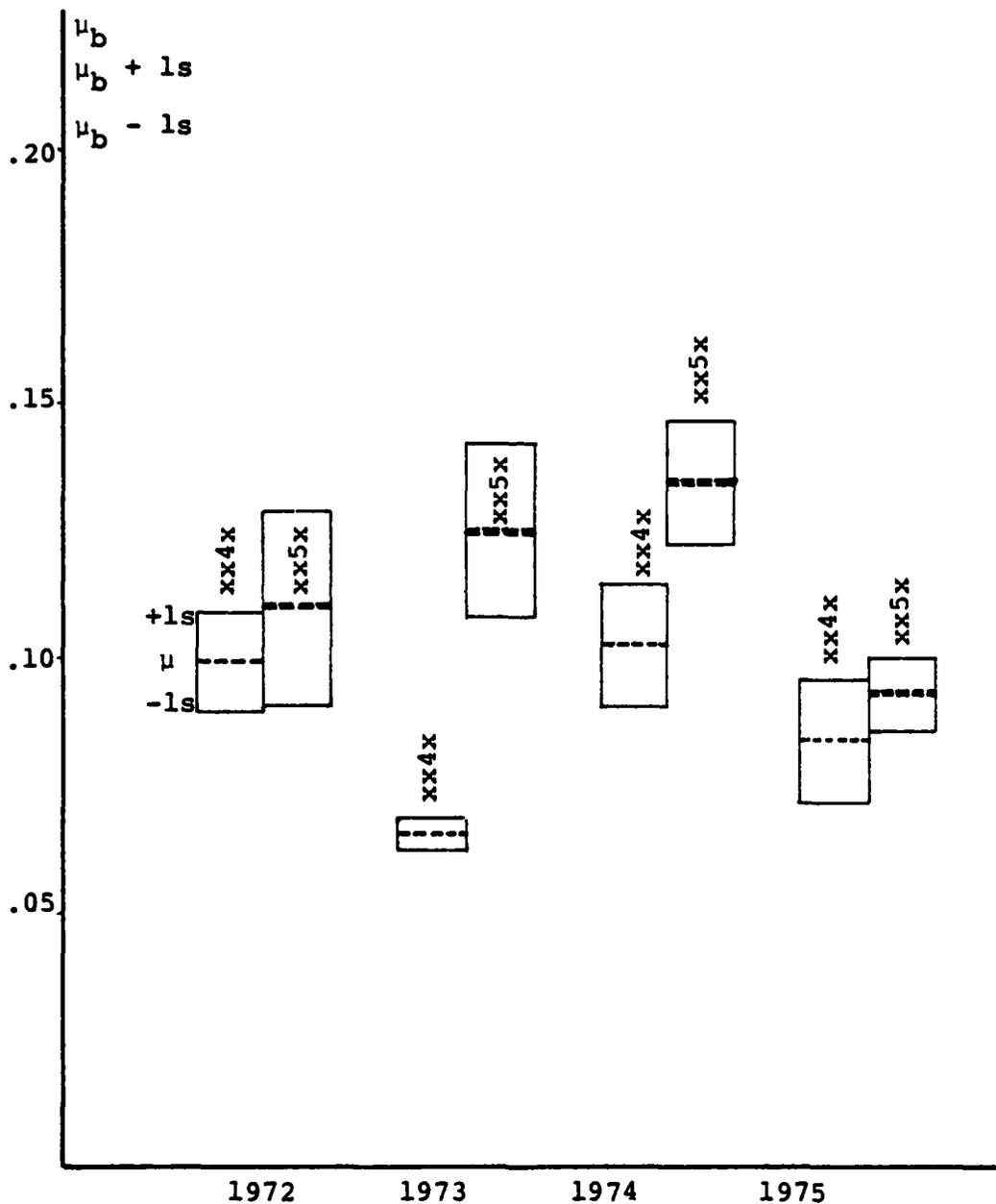


Figure VIII: Means ( $\mu$ ) of the  $b_{ALO}$  of the SSC's in SSC groups xx4x and xx5x and the regions bounded by  $\pm$  standard deviation (s) for  $k = 5$  years after graduation. Shown are graduation-year groups 1972 to 1975.

A comparison of the figures reveals also the effect of the Bayesian estimation technique on the differences between the losses for each SSC within the SSC group they belong to. The size of the standard deviations depicted have decreased considerably from the  $r_{ALO}$  figure to the  $b_{ALO}$  figure.

### 3. Correlation Between Rates and Their Estimates

The estimates calculated according to equation (15) of paragraph V.A. in conjunction with equation (16) of paragraph V.B. are supposed to provide for valid estimates of the loss rates conditioned on the graduation year and the SSC. Provided there exists a linear relationship between the estimates  $b$  and the actual rates  $r_\ell$  the Pearson correlation coefficient ( $r_{br}$ ) should be an indication for the closeness of the estimates to the actual rates.

The coefficient is calculated according to the following equation:

$$(18) \quad r_{br} = \frac{\sum_{\ell=1}^L b_\ell r_\ell - \sum_{\ell=1}^L b_\ell \sum_{\ell=1}^L r_\ell}{\sqrt{\sum_{\ell=1}^L b_\ell^2 - \left(\sum_{\ell=1}^L b_\ell\right)^2} \sqrt{\sum_{\ell=1}^L r_\ell^2 - \left(\sum_{\ell=1}^L r_\ell\right)^2}}$$

where  $-1 < r_{br} < +1$ ,

$L$  = total number of estimates =  
total number of rates.

The graphs included in Appendix E depict the relation between rates and estimates for each graduation year. They do not reveal any strong pattern of deviation from linearity.

However, they show for all graduation years that for rates close to zero the estimate  $b$  tends to deviate more from the actual rate  $r$  than for higher loss rates.

Table 13 summarizes the findings for  $r_{br}$  for graduation years 1970 to 1975. In addition it shows the amount of variation of the loss rates that is explained by the estimator, denoted by  $r_{br}^2$ . The total number of estimates-- $L$  as shown in the table--is the product of the number of SSC's  $j$  and the number of years  $k$  after graduation for the specific graduation year for the specific graduation year, where SSC's of size zero are excluded for that graduation year.

In analyzing the correlation between the loss rates and their estimates it is possible to answer the question how close the behavior of officers belonging to the various SSC's follows the trend of the respective SSC groups.

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	Graduation years					
	1970	1971	1972	1973	1974	1975
$r_{br}$	.85	.95	.93	.91	.76	.87
$r_{br}^2$	.72	.91	.86	.84	.58	.76
$L$	380	360	304	273	228	185

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Table 13: Correlation-coefficients  $r_{br}$  for accumulated total loss-rates  $r$  and their estimates  $b$  and the respective  $r_{br}^2$ .  $L$  is the number of coefficients calculated for the resp. graduation year.

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According to the structure of the estimator,  $r_{br}$  would equal 1.0 only if for that graduation year for each of the ten SSC groups the loss rates for all SSC's belonging to the SSC group were identical. The more  $r_{br}$  deviates from 1.0 the more differences between the various SSC's and their respective SSC groups are encountered. Thus, according to the coefficients in Table 13, this variation is highest for graduation year 1974 and extremely low for graduation year 1971.

As was mentioned before, chance influences should have more impact in smaller groups than for bigger ones. Accordingly the coefficients  $r_{br}$  should be smaller when only small SSC sizes are considered for calculating  $r_{br}$ .

Table 13 compares the  $r_{br}$  which were obtained for SSC sizes

- bigger than 20 versus smaller than 20,
- bigger than 10 versus smaller than 10, and
- bigger than 5 versus smaller than 5.

The values of  $r_{br}$  show generally an increasing trend with increasing SSC sizes used for calculated them. However, for graduation year 1970, 1971, and 1973 the values  $r_{br}$  increase again when SSC sizes less than five are used for calculating them. This unexpected increase becomes explainable when only accumulated loss rates due to other than retirement reasons  $r_{ALO}$  are correlated with their estimates  $b_{ALO}$ . Now, not only the expected trend is true for all graduation years but the differences between the coefficients are more obvious, as Table 15 shows.

SSC sizes selected				
Grad year		a) GE 20 b) LT 20	a) GE 10 b) LT 10	a) GE 5 b) LT 5
1970	a)	.99 (100)	.99 (140)	.82 (230)
	b)	.83 (280)	.81 (240)	.92 (150)
1971	a)	.99 ( 99)	.93 (162)	.92 (225)
	b)	.95 (261)	.96 (198)	.98 (135)
1972	a)	.97 (112)	.95 (160)	.96 (216)
	b)	.92 (192)	.92 (144)	.90 ( 88)
1973	a)	.99 ( 84)	.97 (133)	.90 (175)
	b)	.90 (189)	.89 (140)	.92 ( 98)
1974	a)	.99 ( 72)	.96 (114)	.94 (168)
	b)	.73 (156)	.71 (114)	.68 ( 60)
1975	a)	.97 ( 65)	.94 (105)	.92 (130)
	b)	.86 (120)	.86 ( 80)	.85 ( 55)

Table 14: Comparison of correlation coefficients  $r_{br}$  for accumulated total loss rates  $r$  and their estimates  $b$  under consideration of different SSC sizes. (Number of cases in parentheses)

Grad-year	SSC sizes selected		
	a) GE 20 b) LT 20	a) GE 10 b) LT 10	a) GE 5 b) LT 10
1970	a) .98 b) .64	.98 .56	.75 .48
1971	a) .98 b) .60	.89 .45	.74 .26
1972	a) .92 b) .44	.89 .06	.83 .06
1973	a) .98 b) .51	.94 .23	.74 .25
1974	a) .98 b) .42	.90 .40	.82 .50
1975	a) .91 b) .42	.86 .36	.81 .24

Table 15: Comparison of correlation-coefficients  $r_{b_{ALO} r_{ALO}}$  for loss-rates  $r_{ALO}$  and their estimates  $b_{ALO}$  for different SSC sizes.

(Number of cases is the same as in Table 14.)

Obviously, the unexpected increase of  $r_{br}$  for SSC-sizes of less than five individuals was caused by the effect of the retirement rate  $r_{ALR}$ , which enters the model as rate and not as estimate. The resulting adjustment in the direction of the observed overall loss rate is relatively bigger for small SSC sizes, as was to be expected.

The extremely low value of  $r_{br}$  in Table 15 for graduation year 1972 and SSC sizes less than ten is caused by the following facts: 19 out of the 41 SSC's did not show any losses due to other-than-retirement reasons (ALO) over the whole observed time up to year  $k = 8$  after graduation. This resulted in 162 cases with  $r_{ALO}$  equal to zero. Out of the 144 cases for SSC sizes of less than 10 individuals,<sup>44</sup> 138 are cases with  $r_{ALO}$  equal to zero.

Thirty two of the 138 cases which represent the most extreme differences between  $r_{ALO}$  and  $b_{ALO}$  and thus contribute heavily to the low  $r_{br}$  values are shown in Table 16. The SSC's shown are the only ones out of their respective SSC groups which have zero ALO losses over the whole observed period and they happen to be also the only ones out of their groups with SSC sizes less than ten. In addition group xx2x consists completely of SSC's with sizes less than ten. Its seven SSC's make up 56 cases out of which 50 show  $r_{ALO}$  values of zero. The six remaining cases have an  $r_{ALO}$  value of .333.

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<sup>44</sup>See Table 14: 18 SSC with less than 10 officers x 8 observed years k.

SSC	k =	1	2	3	4	5	6	7	8
xx56	$r_{ALO}$	.0	.0	.0	.0	.0	.0	.0	.0
	$b_{ALO}$	.0	.0	.0	.049	.106	.179	.228	.244
xx67	$r_{ALO}$	.0	.0	.0	.0	.0	.0	.0	.0
	$b_{ALO}$	.0	.0	.0	.093	.111	.148	.185	.204
xx72	$r_{ALO}$	.0	.0	.0	.0	.0	.0	.0	.0
	$b_{ALO}$	.0	.0	.0	.020	.078	.098	.157	.235
1102	$r_{ALO}$	.0	.0	.0	.0	.0	.0	.0	.0
	$b_{ALO}$	.0	.0	.0	.021	.063	.104	.125	.167

Table 16: Comparison rates of  $r_{ALO}$  and their estimates for four selected SSC of the group of 18 SSC of graduation year 1972 which consist of less than ten individuals each.

The respective  $b_{ALO}$  values range from .0 to .03. Group xx3x--consisting of five SSC's--contributes 3 SSC's to the group of SSC's with less than ten individuals. Each of the three does not show any ALO losses within the eight observed years, whereas the two remaining SSC's show  $r_{ALO}$  values of up to .113 for  $k = 8$  and SSC xx31. This results in differences between the  $r_{ALO}$  values and their respective estimates of up to .083. The same reasoning applies to SSC group 13xx, which consists of seven SSC's. Three of them belong to the class of SSC's with less than ten individuals, and all three do not show ALO losses. The losses encountered for the rest of the SSC's in group 13xx result in differences between  $r_{ALO}$  values and their estimates of up to .71.

However, it has to be stated that the class of SSC's having sizes of less than ten individuals each represents only 67 out of the 768 officers who graduated in 1972. Up to the eighth year after graduation, which was the last year observed, only one of the 67 officers was lost due to other-than-retirement reasons. From the remaining 701 officers belonging to SSC's with sizes of ten and more individuals, 119 were lost due to the same reasons.

Except for group xx2x the trends in the SSC groups are results of losses occurring outside the class of SSC's with less than ten individuals. These group trends cover almost completely the fact that there occurred no loss in the small-size class of SSC's except the one in group xx2x.

Thus, the  $b_{ALO}$  values for those small SSC's are more a result of the respective SSC-group trends than of the actual losses within the SSC's. Correspondingly, for SSC's with sizes of less than ten individuals the actual group losses due to other-than-retirement reasons correlate with a coefficient of .9942 with the estimates  $b_{ALO}$  of the SSC losses whereas the  $r_{ALO}$  values for the actual individual losses correlate only with the coefficient of .06 with their estimates  $b_{ALO}$ . These extreme results shown above for graduation year 1972 could not be found for the rest of the graduation-years.

Thus, it cannot be assumed that smaller SSC sizes tend to have smaller loss rates or that for smaller SSC sizes the Bayesian estimator is less valid. Rather, it has to be assumed that the extreme low losses in small-size SSC for graduation year 1972 were a result of chance influences and that the Bayesian estimator achieved what it was supposed to achieve, namely to cope with non-system-inherent chance influences mainly on small-sized groups by adjusting the observed values in the direction of predominant trends.

## VI. SUMMARY OF CONCLUSIONS AND RECOMMENDATIONS

### A. CONCLUSIONS

The distinction between retirement losses and losses for other-than-retirement reasons that formed the basis for the construction of the estimation model was most helpful in analyzing underlying trends of the observable losses with respect to the graduation-year group an officer belonged to and with respect to the years after graduation. It led to the observation that the estimation and prediction effort can be restricted to a relatively short-time period of approximately ten years after graduation. The data showed that with the tenth year after graduation the relative share of observable retirement losses with respect to total losses has already reached 81%, with a strongly increasing trend. The data indicated also that losses due to other-than-retirement reasons do not vary linearly with subsequent years after graduation. They increase rapidly to a peak level between year five and year seven after graduation. Then they decrease steadily.

In addition it was found that relative to the year of graduation accumulated total losses, accumulated retirement losses, and accumulated losses due to other-than-retirement reasons varied with an increasing trend between graduation years 1970 and 1974 and--with the exception of the retirement losses--decreased again.

The ten SSC groups observed were found to differ significantly with respect to their losses due to other-than-retirement reasons. The significance of differences among the SSC's within their respective SSC groups was not testable because of limitations imposed by the available test procedures.

The extreme variations in the sample sizes among the 41 SSC's as well as for the single SSC's over time--ranging from zero to over 100 individuals--did not affect the applicability of the Bayesian estimation technique. The estimates showed the expected smoothing effect of the technique, which resulted from the fact that the value of the estimate is a compromise between the information gained from the behavior of the total SSC group and the special behavior of officers in the different SSC's within the groups. Thus the spoiling effect of chance influences especially on small-sized groups could be successfully limited.

Correlating the accumulated total loss rates with their estimates resulted in coefficients between .76 for the graduation year 1974 and .95 for graduation year 1971. As long as only SSC's with sizes bigger than ten individuals each were considered for calculating the coefficients, they ranged from .94 to .99. The magnitude of the correlation coefficient does not only indicate how close the estimates follow their respective rates. It was also found to be a useful indicator for the differences in the retention behavior of officers belonging

to different SSC's within the various SSC groups. A coefficient of 1.0 would be found only in cases where either all SSC groups consisted only of one SSC each or where the loss rates for all SSC's within a SSC group were identical for all SSC groups. With regard to this interpretation, the correlation coefficients were surprisingly high.

#### B. RECOMMENDATIONS FOR EXTENSIONS INTO PREDICTION MODELS

Using the estimates for the loss rates ( $b_{ALO}$ ) instead of the actual observed rates ( $r_{ALO}$ ) provides a more stable and reliable basis for carrying out predictions of future losses of officers who have already graduated and also of losses for future graduation-year groups.

Prediction could be done by using the binomial distribution in a way already proposed by McAfee,<sup>45</sup> however, without his stationarity assumption, as already discussed.

Let

$N_{i,j}$  be the number of officers belonging to SSC  $j$  who graduate in year  $i$ ;

$Y_{i,j,k}$  be the number of officers belonging to  $N_{i,j}$  who have been lost due to other-than-retirement reasons up to and including year  $k$  after graduation; and  
let

$P_{i,j,k}$  be the probability that an officer belonging to

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<sup>45</sup>McAfee, C. K., p. 12.

$N_{i,j}$  belongs also to  $Y_{i,j,k}$ , estimated by  
 $b_{ALO;i,j,k}$ .

Then

$$(19) P(Y_{i,j,k} = y | N_{i,j} = n) = \binom{y}{n} p_{i,j,k}^y (1 - p_{i,j,k})^{n-y}.$$

Because the stationarity assumption about  $p_{i,j,k}$  is no longer in existence, future values of  $p_{i,j,k}$  have to be predicted by projecting the corresponding  $b_{ALO;i,j,k}$  to the future. This can be done by exploiting known trends and establishing trend lines as proposed at the end of paragraph IV.B.3: For each year  $k$  after graduation and each SSC  $j$ , trends of the loss estimates  $b_{ALO}$  over graduation years  $i$  are observable. They can be translated into trend lines, for example, by using regression models. Thus, for all SSC  $j$  and years  $k$  after graduation of interest,  $b_{ALO}$  values for future graduation years  $i$  and future years  $k$  after graduation or  $b_{ALO}$  values for past graduation years  $i$  but future years  $k$  after graduation are obtainable.

Using these  $b_{ALO}$  values as entries for the respective  $p_{i,j,k}$  in above equation (19) yields probability statements about future losses due to other-than-retirement losses. Probability statements about total future losses are obtainable by using an approach similar to that one proposed in Chapter V. As long as only expected values of future loss rates are wanted, the approach represented by equation (19) is not needed: The predicted  $b_{ALO;i,j,k}$  values can be combined

with the known future  $r_{ALR;i,j,k}$  values as proposed in Chapter V. Thus, expected values of future loss rates will be obtained.

APPENDIX A-1  
SUBSPECIALTY CODES

Subspecialty-code	Subspecialty title
<u>Group xx2x NATIONAL SECURITY AFFAIRS</u>	
xx21	Mideast, Africa, or South Asia
xx22	Far East, Southeast Asia, or Pacific Ocean
xx23	Western Hemisphere
xx24	Europe, USSR
xx25	International Organizations and Negotiations
xx26	Strategic Planning
xx27	Nuclear Planning
<u>Group xx3x ADMINISTRATIVE SCIENCE</u>	
xx31	Financial Management
xx32	Material Management
xx33	Manpower/Personnel Analysis
xx34	Logistics Management
xx38	Human Resource Management
<u>Group xx4x, APPLIED LOGIC AND OPERATIONS SYSTEMS TECHNOLOGY</u>	
xx42	Operations Research/Systems Analysis
xx44	Anti-Submarine Warfare Systems Technology
xx48	Meteorology
xx49	Oceanography

Subspecialty-code                      Subspecialty Title

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Group xx5x    NAVAL SYSTEMS ENGINEERING

xx51                      Naval Construction and Engineering  
xx52                      Nuclear Engineering  
xx54                      Naval/Mechanical Engineering  
xx55                      Engineering Electronics  
xx56                      Underwater Acoustics

Group xx6x    WEAPON SYSTEMS ENGINEERING

xx61                      Weapons Systems Technology  
xx62                      Chemistry  
xx63                      Weapons Systems Science (Physics)  
xx67                      Nuclear Effects (Physics)

Group xx7x    AERONAUTICAL ENGINEERING

xx71                      Aeronautical Engineering  
xx72                      Aeronautical Engineering with Avionics

Group xx8x    COMMUNICATIONS

xx81                      Communications Engineering  
xx82                      Telecommunications Systems Management

Group xx9x    COMPUTER TECHNOLOGY

xx91                      Computer Science  
xx95                      Computer Systems Management

Subspecialty-code                      Subspecialty Title  
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Group 11xx CIVIL ENGINEERING

1101                      Facilities Engineering  
1102                      Petroleum Engineering  
1103                      Ocean Engineering

Group 13xx SUPPLY

1301                      Supply Acquisitions/Distribution Management  
1302                      Systems Inventory Management  
1304                      Material Movement  
1305                      Retailing  
1306                      Acquisition Contract Management  
1307                      Petroleum Management  
1308                      Subsistence Technology

APPENDIX A-2

ORIGINAL RECORD STRUCTURE

SSN	Month/Year of birth	Month/Year of entry	Year of graduation	Month/Year of Minimum Service Requirement	Month/Year of loss	Year eligible to retire	SSC
(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)

APPENDIX B

SAMPLE SIZES FOR SUBSPECIALTY CODES

SSC	1970	1971	1972	1973	1974	1975
xx21	6	7	3	0	3	0
xx22	4	8	9	11	10	4
xx23	4	1	3	4	3	1
xx24	7	16	5	11	15	7
xx25	2	4	6	2	1	0
xx26	8	2	5	4	7	1
xx27	2	1	1	2	1	0
xx31	66	66	62	59	56	72
xx32	32	37	34	24	33	25
xx33	3	6	5	7	2	2
xx34	3	3	4	6	3	4
xx38	2	3	2	4	7	11
xx42	45	69	85	96	54	57
xx44	0	1	0	0	0	20
xx48	4	18	34	23	25	30
xx49	43	51	44	36	28	24
xx51	2	8	11	11	9	12
xx52	14	6	19	15	18	17
xx54	44	50	54	38	28	33
xx55	27	18	27	26	31	18
xx56	6	7	6	6	11	7
xx61	9	10	18	9	6	17
xx62	10	11	14	12	12	10
xx63	17	19	20	19	26	22
xx67	2	3	1	2	2	0
xx71	25	41	47	54	44	39
xx72	4	1	2	3	5	5

SSC	1970	1971	1972	1973	1974	1975
xx81	6	8	15	6	8	9
xx82	0	31	48	30	33	25
xx91	21	31	26	22	12	22
xx95	72	88	57	56	63	34
1101	31	42	46	34	52	45
1102	0	1	1	1	0	2
1103	1	0	0	1	5	4
1301	14	23	23	13	13	13
1302	0	4	3	2	4	2
1304	5	2	6	3	6	5
1305	3	3	2	3	3	1
1307	5	4	3	3	0	3
1308	1	1	0	3	1	1
1306	8	10	17	8	9	14

## APPENDIX C

### Alternative Derivation of Equation (14) of Paragraph V.A

The following derivation omits any subscripts. The events and probabilities refer to one SSC  $j$  of one graduation-year and are derived for one year  $k$  after graduation. The results apply to any  $i$ ,  $j$ , and  $k$  observable as stated in paragraph V.A.

Let  $P(NLO)$  be the probability that an officer still serves in year  $k$  after graduation. Let  $P(R)$  be the probability that an officer reaches his retirement age between graduation and year  $k$  and actually retires for that reason.

Let

$$P(ALO) = 1 - P(NLO).$$

Then the total probability

$$P(R) = P(R|NLO)P(NLO) + P(R|ALO)P(ALO)$$

However  $P(R|ALO) = 0$  because an officer who left the forces due to other-than-retirement reasons prior to reaching his retirement age can obviously not retire. Thus

$$P(R) = P(R|NLO)P(NLO)$$

Now

$$P(R|NLO) = \frac{P(R \cap NLO)}{P(NLO)}$$

and thus

$$P(R) = \frac{P(R \cap NLO)}{P(NLO)} P(NLO)$$

From the data observed

$$P(R \cap NLO) = \frac{ALRo}{N} \quad \text{and}$$

$$P(NLO) = \frac{N - ALO}{N}$$

Thus

$$P(R) = \frac{ALRo/N}{(N - ALO)/N} (1 - P(ALO))$$

In order to obtain the probability that an officer will be lost between his graduation year and the year k after graduation, denoted by P(AL) only P(ALO) has to be added, that is

$$P(AL) = P(ALO) + P(R)$$

This statement corresponds with equation (8) in paragraph V.A.

Now

$$P(ALO) = \frac{ALO}{N}$$

and thus

$$P(AL) = \frac{ALO}{N} + \frac{ALRo}{(N - ALO)} \left(1 - \frac{ALO}{N}\right)$$

In terms of loss rates this equation is equivalent to equation (14) of paragraph V.A.

APPENDIX D  
TABLES OF LOSS RATES AND THEIR ESTIMATES

Appendix D shows in six sections (D-1 to D-6)--one section for each graduation year--loss rates and the estimates of the rates conditioned on the graduation year and conditioned on the Sub-specialty code an officer has obtained.

- Section D-1: Graduation year 1970 and  $k = 1, 2, \dots, 10$ ;  
(Tables 1 to 3)
- Section D-2: Graduation year 1971 and  $k = 1, 2, \dots, 9$ ;  
(Tables 1 to 3)
- Section D-3: Graduation year 1972 and  $k = 1, 2, \dots, 8$ ;  
(Tables 1 to 3)
- Section D-4: Graduation year 1973 and  $k = 1, 2, \dots, 7$ ;  
(Tables 1 to 3)
- Section D-5: Graduation year 1974 and  $k = 1, 2, \dots, 6$ ;  
(Tables 1 to 3)
- Section D-6: Graduation year 1975 and  $k = 1, 2, \dots, 5$ ;  
(Tables 1 to 3)

The content of the tables was explained in chapter V.C.1 and is explained in the respective headings of the tables within the six sections. An entry of 99.0 within the tables indicates that a rate could not be calculated because nobody having the SSC under consideration graduated in the respective graduation year.

TABLE 1 FCF GRADUATION-YEAR 70  
ACCUMULATED ACTUAL LOSSES

ROW 1 : ACTUAL LCSS RATE  
ROW 2 : ESTIMATE

SSC	1	2	3	4	5	6	7	8	9	10
21	C.C 0.0	0.167 0.051	C.333 0.103	0.500 0.154	0.833 0.256	0.833 0.256	0.833 0.256	0.833 0.256	0.833 0.256	C.833 C.256
22	C.C 0.0	0.0 0.027	0.0 0.054	0.0 0.081	0.0 0.135	0.0 0.135	0.0 0.135	0.0 0.135	0.0 0.135	C.0 C.135
23	C.C 0.0	0.0 0.027	0.0 0.054	0.0 0.081	0.0 0.135	0.0 0.135	0.0 0.135	0.0 0.135	0.0 0.135	C.0 C.135
24	C.C 0.0	0.0 0.025	0.0 0.050	0.0 0.075	0.0 0.125	0.0 0.125	0.0 0.125	0.0 0.125	0.0 0.125	C.286 C.375
25	C.C 0.0	0.0 0.029	0.0 0.057	0.0 0.086	0.0 0.143	0.0 0.143	0.0 0.143	0.0 0.143	0.0 0.143	1.000 1.000
26	C.C 0.0	0.0 0.024	0.0 0.049	0.0 0.073	0.0 0.122	0.0 0.451	0.0 0.451	0.0 0.451	0.0 0.451	C.375 C.451
27	C.C 0.0	0.0 0.029	0.0 0.057	0.0 0.086	0.0 0.143	0.0 0.500	0.0 0.500	0.0 0.500	0.0 0.500	C.500 C.571
31	C.C 0.0	0.0 0.0	0.0 0.0	C.015 0.012	0.030 0.023	0.076 C.077	0.091 0.057	0.121 0.128	0.152 0.158	C.212 C.218
32	C.C 0.0	0.0 0.0	0.0 0.0	0.0 0.007	0.0 0.014	0.063 C.051	0.125 0.057	0.219 0.154	0.313 0.291	C.344 C.323
33	C.C 0.0	0.0 0.0	0.0 0.0	0.0 0.009	0.0 0.018	0.0 C.046	0.0 0.055	0.333 0.370	0.333 0.370	C.333 C.370
34	C.C 0.0	0.0 0.0	0.0 0.0	0.0 0.009	0.0 0.018	0.0 C.046	0.0 0.055	0.0 0.055	0.0 0.055	C.0 C.055
38	C.C 0.0	0.0 0.0	0.0 0.0	0.0 0.009	0.0 0.019	0.0 C.046	0.0 0.500	0.0 0.528	0.0 0.528	C.500 C.528





TABLE 2 FOR GRADUATION-YEAR 70  
ACCUMULATED LCSES (EXCL. RETIREMENT)

ROW 1 : ACTUAL LCSS RATE  
ROW 2 : ESTIMATE

SSC	1	2	3	4	5	6	7	8	9	10
21	C.C C.C	0.167 0.051	0.333 0.103	0.500 0.154	0.833 0.256	0.833 0.256	0.833 0.256	0.833 0.256	0.833 0.256	C.833 C.256
22	C.C C.C	0.0 0.027	0.0 0.054	0.0 0.081	0.0 0.135	0.0 0.135	0.0 0.135	0.0 0.135	0.0 0.135	C.0 C.135
23	C.C C.C	0.0 0.027	0.0 0.054	0.0 0.081	0.0 0.135	0.0 0.135	0.0 0.135	0.0 0.135	0.0 0.135	C.0 C.135
24	C.C C.C	0.0 0.025	0.0 0.050	0.0 0.075	0.0 0.125	0.0 0.125	0.0 0.125	0.0 0.125	0.0 0.125	C.0 C.125
25	C.C C.C	0.0 0.029	0.0 0.057	0.0 0.086	0.0 0.143	0.0 0.143	0.0 0.143	0.0 0.143	0.0 0.143	C.0 C.143
26	C.C C.C	0.0 0.024	0.0 0.049	0.0 0.073	0.0 0.122	0.0 0.122	0.0 0.122	0.0 0.122	0.0 0.122	C.0 C.122
27	C.C C.C	0.0 0.029	0.0 0.057	0.0 0.086	0.0 0.143	0.0 0.143	0.0 0.143	0.0 0.143	0.0 0.143	C.0 C.143
31	C.C C.C	0.0 0.0	0.0 0.0	0.015 0.012	0.030 0.023	0.045 0.047	0.045 0.052	0.045 0.052	0.045 0.052	C.045 C.052
32	C.C C.C	0.0 0.0	0.0 0.0	0.0 0.007	0.0 0.014	0.062 0.051	0.094 0.065	0.054 0.065	0.054 0.065	C.054 C.065
33	C.C C.C	0.0 0.0	0.0 0.0	0.0 0.009	0.0 0.018	0.0 0.046	0.0 0.055	0.0 0.055	0.0 0.055	C.0 C.055
34	C.C C.C	0.0 0.0	0.0 0.0	0.0 0.009	0.0 0.018	0.0 0.046	0.0 0.055	0.0 0.055	0.0 0.055	C.0 C.055
38	C.C C.C	0.0 0.0	0.0 0.0	0.0 0.009	0.0 0.019	0.0 0.046	0.0 0.056	0.0 0.056	0.0 0.056	C.0 C.056

42	C.C	0.0	0.0	0.022	0.022	0.015	0.022	0.044	0.044	0.044	0.067	0.067	C.085
	C.C	0.0	0.0	0.015	0.015	0.010	0.015	0.025	0.029	0.044	0.051	0.051	C.066
44	SS.CCC	99.000	99.000	99.000	99.000	99.000	99.000	99.000	99.000	99.000	99.000	99.000	99.000
	C.C	0.0	0.0	0.011	0.011	0.010	0.011	0.022	0.022	0.033	0.043	0.043	C.054
4E	C.C	0.0	0.0	0.010	0.010	0.010	0.021	0.021	0.021	0.031	0.042	0.042	C.052
45	C.C	0.0	0.0	0.007	0.007	0.007	0.015	0.015	0.022	0.022	0.037	0.037	C.023
	C.C	0.0	0.0	0.007	0.007	0.007	0.015	0.015	0.022	0.022	0.037	0.037	C.044
51	C.C	0.0	0.0	0.0	0.0	0.011	0.095	0.500	0.500	0.500	0.500	0.500	C.500
	C.C	0.0	0.0	0.0	0.0	0.011	0.095	0.221	0.221	0.253	0.284	0.284	C.305
52	C.C	0.0	0.0	0.0	0.0	0.009	0.214	0.214	0.214	0.252	0.286	0.286	C.429
	C.C	0.0	0.0	0.0	0.0	0.009	0.112	0.215	0.215	0.252	0.280	0.280	C.318
54	C.C	0.0	0.0	0.0	0.0	0.007	0.114	0.227	0.227	0.248	0.255	0.255	C.255
	C.C	0.0	0.0	0.0	0.0	0.007	0.102	0.219	0.219	0.248	0.285	0.285	C.299
55	C.C	0.0	0.0	0.0	0.0	0.037	0.037	0.222	0.222	0.242	0.255	0.255	C.255
	C.C	0.0	0.0	0.0	0.0	0.017	0.083	0.217	0.217	0.242	0.275	0.275	C.292
56	C.C	0.0	0.0	0.0	0.0	0.010	0.0	0.0	0.0	0.167	0.167	0.167	C.167
	C.C	0.0	0.0	0.0	0.0	0.010	0.091	0.202	0.202	0.242	0.273	0.273	C.293
61	C.C	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.111	0.111	C.111
	C.C	0.0	0.0	0.0	0.0	0.0	0.085	0.128	0.128	0.170	0.255	0.255	C.255
62	C.C	0.0	0.0	0.0	0.0	0.0	0.100	0.300	0.300	0.300	0.300	0.300	C.300
	C.C	0.0	0.0	0.0	0.0	0.0	0.104	0.188	0.188	0.229	0.292	0.292	C.292
63	C.C	0.0	0.0	0.0	0.0	0.0	0.176	0.176	0.176	0.236	0.327	0.327	C.412
	C.C	0.0	0.0	0.0	0.0	0.0	0.127	0.164	0.164	0.236	0.327	0.327	C.327
67	C.C	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	C.0
	C.C	0.0	0.0	0.0	0.0	0.0	0.100	0.150	0.150	0.200	0.275	0.275	C.275
71	C.C	0.0	0.0	0.0	0.0	0.0	0.040	0.160	0.160	0.200	0.280	0.280	C.360
	C.C	0.0	0.0	0.0	0.0	0.0	0.037	0.148	0.148	0.185	0.255	0.255	C.333
72	C.C	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	C.0
	C.C	0.0	0.0	0.0	0.0	0.0	0.030	0.121	0.121	0.152	0.212	0.212	C.273



TABLE 3 FOR GRADUATION-YEAR 70

ACCUMULATED LOSSES

ROW 1 : ACC-LCNS RATE OF SSC-(RP INDICATED)  
 ROW 2 : ACC-LCNS RATE (DIFF) OF SSC-GRP  
 ROW 3 : ACC-LCNS RATE (RETIRE) OF SSC-GRP  
 LAST 3 ROWS REFER TO TOTAL GRAD-YEAR GROUP

SSC	1	2	3	4	5	6	7	8	9	10
20	C.C 0.0	0.030 0.030	0.061 0.061	0.091 0.091	0.152 0.152	0.273 0.143	0.273 0.143	0.273 0.143	0.303 0.179	C.394 C.152 C.286
30	C.C 0.0	0.0 0.0	0.0 0.0	0.009 0.009	0.019 0.019	0.066 0.041	0.104 0.097	0.160 0.097	0.208 0.097	C.235 C.057 C.210
40	0.0 0.0	0.0 0.0	0.011 0.011	0.011 0.011	0.011 0.011	0.022 0.022	0.033 0.022	0.076 0.033	0.155 0.043	C.207 C.054 C.161
50	C.C 0.0	0.0 0.0	0.0 0.0	0.0 0.0	0.011 0.011	0.057 0.041	0.247 0.041	0.280 0.043	0.333 0.075	C.355 C.301 C.077
60	C.C 0.0	0.0 0.0	C.C 0.0	C.C 0.0	0.0 0.0	0.103 0.103	0.158 0.0	0.211 0.0	0.316 0.037	C.395 C.289 C.148
70	0.0 0.0	0.0 0.0	0.0 0.0	0.0 0.0	0.0 0.0	0.065 0.034	0.172 0.138	0.267 0.042	0.316 0.091	C.379 C.310 C.108
80	C.C 0.0	0.0 0.0	C.C 0.0	0.0 0.0	0.0 0.0	0.0 0.0	0.0 0.0	0.0 0.0	0.0 0.0	C.0 C.0
90	C.C 0.0	0.0 0.0	C.C 0.0	C.C 0.0	0.0 0.0	0.054 0.022	0.129 0.036	0.154 0.085	0.237 0.123	C.258 C.140 C.137
1100	0.0 0.0	0.0 0.0	0.0 0.0	0.0 0.0	0.0 0.0	0.063 0.063	0.054 0.094	0.054 0.094	0.125 0.034	C.156 C.094 C.069
1300	C.C 0.0	0.0 0.0	C.C 0.0	C.C 0.0	0.0 0.0	0.0 0.0	0.050 0.028	0.075 0.041	0.150 0.125	C.275 C.025 C.256
70	0.0 0.0	0.002 0.002	0.005 0.005	0.009 0.009	0.016 0.016	0.071 0.055	0.112 0.100	0.173 0.116	0.224 0.133	C.281 C.144 C.160

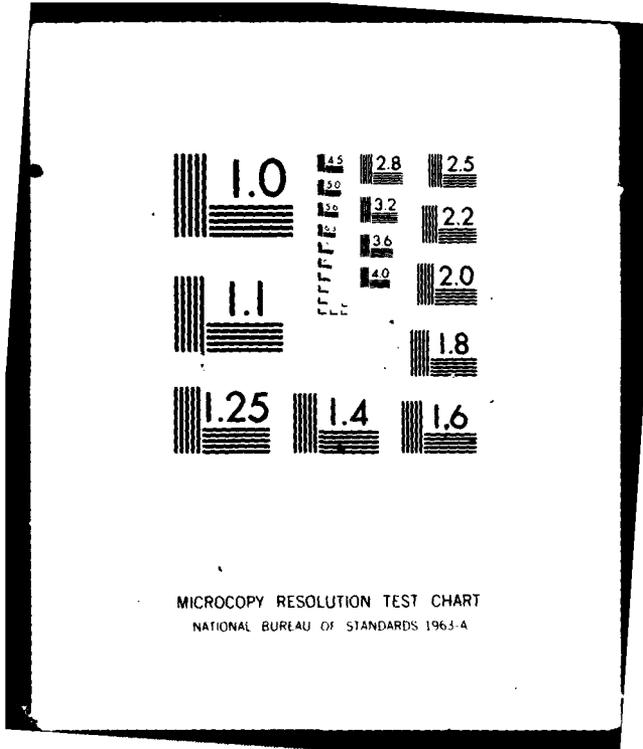
TABLE 1 FCR GRADUATION-YEAR 71  
ACCUMULATED ACTUAL LOSSES

ROW 1 : ACTUAL LOSS RATE  
ROW 2 : ESTIMATE

SSC	1	2	3	4	5	6	7	8	9
21	C:C 0:0	0:0 0:0	0:143 0:043	0:429 0:130	0:857 0:732	0:857 0:732	0:857 0:732	0:857 0:732	0:857 0:732
22	C:C 0:0	C:C 0:0	0:C 0:021	0:C 0:064	0:0 0:106	0:0 0:106	0:0 0:106	0:0 0:106	0:0 0:106
23	C:C 0:0	C:C 0:0	0:C 0:025	0:0 0:075	0:0 0:125	0:0 0:125	0:0 0:125	0:0 0:125	0:0 0:125
24	C:C 0:0	C:C 0:0	0:C 0:018	0:C 0:055	0:0 0:091	0:0 0:091	0:0 0:091	0:0 0:091	0:0 0:091
25	C:C 0:0	0:C 0:0	C:C 0:023	0:0 0:070	0:250 0:140	0:250 0:140	0:250 0:140	0:250 0:140	0:250 0:140
26	C:C 0:0	0:C 0:0	0:C 0:024	0:C 0:073	1:000 1:000	1:000 1:000	1:000 1:000	1:000 1:000	1:000 1:000
27	C:C 0:0	0:C 0:0	C:C 0:025	0:C 0:075	1:000 1:000	1:000 1:000	1:000 1:000	1:000 1:000	1:000 1:000
31	C:C 0:0	C:C 0:0	C:C 0:061	0:C 0:083	0:136 0:137	0:157 0:151	0:242 0:238	0:303 0:300	0:348 0:351
32	C:C 0:0	C:C 0:0	0:C 0:013	C:0 0:013	0:081 0:066	0:108 0:055	0:135 0:118	0:324 0:305	0:432 0:415
33	C:C 0:0	C:C 0:0	0:C 0:017	0:C 0:017	0:0 0:058	0:0 0:051	0:333 0:405	0:333 0:416	0:333 0:433
34	C:C 0:0	0:0 0:0	C:C 0:017	0:C 0:017	0:0 0:059	0:0 0:053	0:333 0:407	0:333 0:418	0:667 0:718
35	C:C 0:0	C:0 0:0	0:C 0:017	0:0 0:017	0:0 0:059	0:0 0:053	0:0 0:110	0:0 0:127	0:333 0:161

42	C:C	C:0	0:C	0:0	0:0	0:029	0:058	0:087	0:145	0:174
	C:0	0:0	0:0	0:0	0:043	C:08E	C:08E	0:114	0:162	0:151
44	C:0	0:0	C:0	0:0	0:0	0:021	0:057	0:071	0:086	0:086
	C:0	0:0	0:0	0:0	0:056	0:074	C:111	0:167	0:167	0:167
48	C:0	0:0	0:0	0:0	0:074	0:113	C:113	0:179	0:151	0:191
	C:0	0:0	0:0	0:0	0:059	0:137	C:054	0:157	0:157	0:157
49	C:0	0:0	C:0	0:0	0:032	C:054	0:110	0:110	0:120	0:120
	0:0	0:0	0:0	0:0	0:250	C:250	0:250	0:250	0:250	0:500
51	C:0	0:0	C:0	0:0	0:113	C:144	0:268	0:320	0:320	0:552
	0:0	0:0	0:0	0:0	0:0	0:0	0:167	0:167	0:167	0:167
52	C:0	0:0	C:0	0:0	0:095	C:12E	0:263	0:316	0:316	0:326
	0:0	0:0	0:0	0:0	0:160	C:24C	0:400	0:500	0:500	0:520
54	C:0	0:0	C:0	0:0	0:136	C:20C	0:359	0:435	0:435	0:450
	0:0	0:0	0:0	0:0	0:0	0:056	0:278	0:278	0:278	0:332
55	C:0	0:0	C:0	0:0	0:084	C:161	0:314	0:358	0:358	0:415
	0:0	0:0	0:0	0:0	0:0	0:0	0:0	0:143	0:143	0:143
56	C:0	0:0	C:0	0:0	0:094	C:12E	0:250	0:402	0:411	0:411
	0:0	0:0	0:0	0:0	0:0	0:0	0:200	0:200	0:200	0:200
61	C:0	0:0	C:0	0:0	0:094	0:151	0:329	0:346	0:346	0:363
	0:0	0:0	0:0	0:0	0:364	C:455	0:545	0:545	0:545	0:545
62	C:0	0:0	C:0	0:0	0:167	C:241	0:333	0:352	0:352	0:370
	0:0	0:0	0:0	0:0	0:105	0:211	0:316	0:421	0:421	0:474
62	C:0	0:0	C:0	0:0	0:147	C:225	0:326	0:413	0:413	0:449
	0:0	0:0	0:0	0:0	0:0	0:0	0:0	0:667	0:667	0:667
67	C:0	0:0	C:0	0:0	0:109	0:174	0:261	0:761	0:761	0:768
	0:0	0:0	0:0	0:0	0:058	0:122	0:155	0:366	0:366	0:390
71	C:0	0:0	C:0	0:0	0:097	C:121	0:194	0:363	0:363	0:388
	0:0	0:0	0:0	0:0	0:0	0:0	0:0	0:0	0:0	1:000
72	C:0	0:0	C:0	0:0	0:047	C:070	0:116	0:233	0:233	1:000
	0:0	0:0	0:0	0:0						





MICROCOPY RESOLUTION TEST CHART  
NATIONAL BUREAU OF STANDARDS 1963-A



TABLE 2 FOR GRADUATION-YEAR 71  
 ACCUMULATED LOSSES (EXCL. RETIREMENT)

RCM 1 : ACTUAL LOSS RATE  
 RCM 2 : ESTIMATE

SSC	1	2	3	4	5	6	7	8	9
21	C:C C:C	0:0 0:0	0:143 0:043	0:429 0:130	0:571 0:196	0:571 0:196	0:571 0:196	0:571 0:196	0:571 0:196
22	C:C C:C	0:0 0:0	0:0 0:021	0:0 0:064	0:0 0:106	0:0 0:106	0:0 0:106	0:0 0:106	0:0 0:106
23	C:C C:C	0:0 0:0	0:0 0:025	0:0 0:075	0:0 0:125	0:0 0:125	0:0 0:125	0:0 0:125	0:0 0:125
24	C:C C:C	0:0 0:0	0:0 0:018	0:0 0:055	0:0 0:091	0:0 0:091	0:0 0:091	0:0 0:091	0:0 0:091
25	C:C C:C	0:0 0:0	0:0 0:023	0:0 0:070	0:250 0:140	0:250 0:140	0:250 0:140	0:250 0:140	0:250 0:140
26	C:C C:C	0:0 0:0	0:0 0:024	0:0 0:073	0:0 0:122	0:0 0:122	0:0 0:122	0:0 0:122	0:0 0:122
27	C:C C:C	0:0 0:0	0:0 0:025	0:0 0:075	0:0 0:125	0:0 0:125	0:0 0:125	0:0 0:125	0:0 0:125
31	C:C C:C	0:0 0:0	0:030 0:022	0:030 0:022	0:061 0:061	0:106 0:099	0:121 0:116	0:136 0:133	0:152 0:155
32	C:C C:C	0:0 0:0	0:0 0:013	0:0 0:013	0:081 0:066	0:108 0:099	0:135 0:118	0:162 0:138	0:185 0:164
33	C:C C:C	0:0 0:0	0:0 0:017	0:0 0:017	0:0 0:058	0:0 0:091	0:0 0:107	0:0 0:124	0:0 0:145
34	C:C C:C	0:0 0:0	0:0 0:017	0:0 0:017	0:0 0:059	0:0 0:093	0:0 0:110	0:0 0:127	0:0 0:153
38	C:C C:C	0:0 0:0	0:0 0:017	0:0 0:017	0:059 0:059	0:093 0:093	0:110 0:110	0:127 0:127	0:161 0:161

42	C:C	0:0	0:0	0:0	0:0	0:0	0:0	0:0	0:014	0:014	0:029	0:058	0:058	0:077	0:058	0:077
44	C:C	0:0	0:0	0:0	0:0	0:0	0:0	0:0	0:021	0:057	0:071	0:086	0:086	0:086	0:086	0:086
48	C:C	0:0	0:0	0:0	0:0	0:0	0:0	0:0	0:019	0:056	0:056	0:056	0:056	0:056	0:056	0:056
49	C:C	0:0	0:0	0:0	0:0	0:0	0:0	0:0	0:059	0:118	0:137	0:137	0:137	0:137	0:137	0:137
51	C:C	0:0	0:0	0:0	0:0	0:0	0:0	0:0	0:250	0:250	0:250	0:250	0:250	0:250	0:250	0:250
52	C:C	0:0	0:0	0:0	0:0	0:0	0:0	0:0	0:095	0:126	0:167	0:167	0:167	0:167	0:167	0:167
54	C:C	0:0	0:0	0:0	0:0	0:0	0:0	0:0	0:140	0:200	0:340	0:340	0:340	0:340	0:340	0:340
55	C:C	0:0	0:0	0:0	0:0	0:0	0:0	0:0	0:084	0:112	0:222	0:222	0:222	0:222	0:222	0:222
56	C:C	0:0	0:0	0:0	0:0	0:0	0:0	0:0	0:094	0:125	0:250	0:250	0:250	0:250	0:250	0:250
61	C:C	0:0	0:0	0:0	0:0	0:0	0:0	0:0	0:094	0:151	0:100	0:100	0:100	0:100	0:100	0:100
62	C:C	0:0	0:0	0:0	0:0	0:0	0:0	0:0	0:364	0:455	0:545	0:545	0:545	0:545	0:545	0:545
63	C:C	0:0	0:0	0:0	0:0	0:0	0:0	0:0	0:097	0:158	0:263	0:263	0:263	0:263	0:263	0:263
67	C:C	0:0	0:0	0:0	0:0	0:0	0:0	0:0	0:109	0:174	0:261	0:261	0:261	0:261	0:261	0:261
71	C:C	0:0	0:0	0:0	0:0	0:0	0:0	0:0	0:049	0:072	0:122	0:122	0:122	0:122	0:122	0:122
72	C:C	0:0	0:0	0:0	0:0	0:0	0:0	0:0	0:047	0:070	0:116	0:116	0:116	0:116	0:116	0:116



TABLE 3 FCS GRADUATION-YEAR 71

ACCUMULATED LOSSES

ROW 1 : ACC-LESS RATE OF SSC-GRP INDICATED  
 ROW 2 : ACC-LESS RATE (OT-INT) OF SSC-GRP  
 ROW 3 : ACC-LESS RATE (INT-INT) OF SSC-GRP  
 LAST 3 ROWS REFER TO TOTAL GRAD-YEAR GROUP

SSC	1	2	3	4	5	6	7	8	9
20	0.0 0.0 0.0	0.0 0.0 0.0	0.026 0.026 0.026	0.077 0.077 0.077	0.256 0.256 0.256	0.256 0.256 0.256	0.256 0.256 0.256	0.282 0.282 0.282	0.333 0.333 0.333
30	0.0 0.0 0.0	0.0 0.0 0.0	0.033 0.033 0.033	0.052 0.052 0.052	0.104 0.104 0.104	0.148 0.148 0.148	0.209 0.209 0.209	0.304 0.304 0.304	0.383 0.383 0.383
40	0.0 0.0 0.0	0.0 0.0 0.0	0.0 0.0 0.0	0.0 0.0 0.0	0.093 0.093 0.093	0.093 0.093 0.093	0.127 0.127 0.127	0.151 0.151 0.151	0.161 0.161 0.161
50	0.0 0.0 0.0	0.0 0.0 0.0	0.0 0.0 0.0	0.0 0.0 0.0	0.117 0.117 0.117	0.117 0.117 0.117	0.153 0.153 0.153	0.182 0.182 0.182	0.227 0.227 0.227
60	0.0 0.0 0.0	0.0 0.0 0.0	0.0 0.0 0.0	0.0 0.0 0.0	0.120 0.120 0.120	0.120 0.120 0.120	0.159 0.159 0.159	0.199 0.199 0.199	0.242 0.242 0.242
70	0.0 0.0 0.0	0.0 0.0 0.0	0.0 0.0 0.0	0.0 0.0 0.0	0.095 0.095 0.095	0.115 0.115 0.115	0.150 0.150 0.150	0.187 0.187 0.187	0.225 0.225 0.225
80	0.0 0.0 0.0	0.0 0.0 0.0	0.0 0.0 0.0	0.0 0.0 0.0	0.091 0.091 0.091	0.077 0.077 0.077	0.103 0.103 0.103	0.103 0.103 0.103	0.203 0.203 0.203
90	0.0 0.0 0.0	0.0 0.0 0.0	0.0 0.0 0.0	0.0 0.0 0.0	0.087 0.087 0.087	0.087 0.087 0.087	0.124 0.124 0.124	0.135 0.135 0.135	0.228 0.228 0.228
1100	0.0 0.0 0.0	0.0 0.0 0.0	0.0 0.0 0.0	0.0 0.0 0.0	0.023 0.023 0.023	0.116 0.116 0.116	0.163 0.163 0.163	0.209 0.209 0.209	0.233 0.233 0.233
1300	0.0 0.0 0.0	0.0 0.0 0.0	0.0 0.0 0.0	0.0 0.0 0.0	0.128 0.128 0.128	0.170 0.170 0.170	0.170 0.170 0.170	0.225 0.225 0.225	0.253 0.253 0.253
71	0.0 0.0 0.0	0.0 0.0 0.0	0.007 0.007 0.007	0.013 0.013 0.013	0.091 0.091 0.091	0.137 0.137 0.137	0.157 0.157 0.157	0.226 0.226 0.226	0.312 0.312 0.312

TABLE 1 FOR GRADUATION-YEAR 72  
ACCUMULATED TOTAL LOSSES

ROW 1 : ACTUAL LOSS RATE  
ROW 2 : ESTIMATE

SSC	1	2	3	4	5	6	7	8
21	0.0 0.0	0.0 0.0	0.333 0.057	0.333 0.057	0.333 0.057	0.333 0.057	0.333 0.057	0.333 0.057
22	0.0 0.0	0.0 0.0	0.0 0.024	0.0 0.024	0.0 0.024	0.0 0.024	0.111 0.133	0.111 0.133
23	0.0 0.0	0.0 0.0	0.0 0.029	0.0 0.029	0.0 0.029	0.0 0.025	0.0 0.029	0.0 0.029
24	0.0 0.0	0.0 0.0	0.0 0.027	0.0 0.027	0.0 0.027	0.0 0.027	0.200 0.222	0.200 0.222
25	0.0 0.0	0.0 0.0	0.0 0.026	0.0 0.026	0.0 0.026	0.0 0.026	0.167 0.189	0.167 0.189
26	0.0 0.0	0.0 0.0	0.0 0.027	0.0 0.027	0.200 0.222	0.200 0.222	0.200 0.222	0.200 0.222
27	0.0 0.0	0.0 0.0	0.0 0.030	0.0 0.030	0.0 0.030	0.0 0.030	0.0 0.030	0.0 0.030
31	0.0 0.0	0.0 0.0	0.048 0.044	0.129 0.113	0.177 0.154	0.177 0.165	0.210 0.193	0.226 0.210
32	0.0 0.0	0.0 0.0	0.0 0.007	0.029 0.057	0.029 0.071	0.088 0.100	0.176 0.193	0.254 0.309
33	0.0 0.0	0.0 0.0	0.0 0.009	0.200 0.229	0.200 0.243	0.200 0.257	0.200 0.264	0.200 0.264
34	0.0 0.0	0.0 0.0	0.0 0.009	0.250 0.277	0.250 0.291	0.500 0.536	0.500 0.541	0.500 0.541
38	0.0 0.0	0.0 0.0	0.0 0.009	0.0 0.037	0.0 0.055	0.0 0.073	0.0 0.083	0.0 0.083

42	0.0	0.0	0.0	0.0	0.012	0.094	0.141	0.188	0.212
	0.0	0.0	0.0	0.0	0.032	0.112	0.159	0.214	0.233
44	55.000	99.000	99.000	99.000	59.000	99.000	99.000	99.000	99.000
	0.0	0.0	0.0	0.0	0.049	0.123	0.160	0.215	0.252
48	0.0	0.0	0.0	0.0	0.088	0.206	0.265	0.324	0.382
	0.0	0.0	0.0	0.0	0.066	0.143	0.175	0.243	0.314
49	0.0	0.0	0.0	0.0	0.091	0.114	0.114	0.182	0.227
	0.0	0.0	0.0	0.0	0.062	0.119	0.138	0.182	0.232
51	0.0	0.0	0.0	0.0	0.0	0.0	0.091	0.091	0.091
	0.0	0.0	0.0	0.0	0.047	0.102	0.180	0.227	0.242
52	0.0	0.0	0.0	0.0	0.0	0.105	0.211	0.263	0.316
	0.0	0.0	0.0	0.0	0.044	0.110	0.191	0.243	0.265
54	0.0	0.0	0.0	0.0	0.111	0.222	0.315	0.407	0.444
	0.0	0.0	0.0	0.0	0.070	0.160	0.257	0.337	0.376
55	0.0	0.0	0.0	0.0	0.0	0.037	0.222	0.333	0.333
	0.0	0.0	0.0	0.0	0.042	0.124	0.300	0.411	0.422
56	0.0	0.0	0.0	0.0	0.333	0.333	0.333	0.500	0.500
	0.0	0.0	0.0	0.0	0.366	0.404	0.453	0.614	0.622
61	0.0	0.0	0.0	0.0	0.0	0.056	0.111	0.222	0.278
	0.0	0.0	0.0	0.0	0.070	0.099	0.141	0.238	0.304
62	0.0	0.0	0.0	0.0	0.214	0.286	0.286	0.286	0.286
	0.0	0.0	0.0	0.0	0.119	0.213	0.240	0.267	0.281
63	0.0	0.0	0.0	0.0	0.100	0.100	0.200	0.300	0.350
	0.0	0.0	0.0	0.0	0.096	0.110	0.201	0.293	0.323
67	0.0	0.0	0.0	0.0	0.0	0.0	0.0	1.000	1.000
	0.0	0.0	0.0	0.0	0.093	0.111	0.148	1.000	1.000
71	0.0	0.0	0.0	0.0	0.064	0.149	0.181	0.258	0.404
	0.0	0.0	0.0	0.0	0.063	0.147	0.189	0.295	0.400
72	0.0	0.0	0.0	0.0	0.0	0.078	0.058	0.0	0.500
	0.0	0.0	0.0	0.0	0.020	0.078	0.058	0.157	0.618



TABLE 2 FCF GRADUATION-YEAR 72  
ACCUMULATED LCSSSES (EXCL. RETIREMENT)

ROW 1 : ACTUAL LCSS RATE  
ROW 2 : ESTIMATE

SSC	1	2	3	4	5	6	7	8
21	0.0 0.0	0.0 0.0	0.333 0.057	0.333 0.057	0.333 0.057	0.333 0.057	0.333 0.057	0.333 0.057
22	0.0 0.0	0.0 0.0	0.0 0.024	0.0 0.024	0.0 0.024	0.0 0.024	0.0 0.024	0.0 0.024
23	0.0 0.0	0.0 0.0	0.0 0.029	0.0 0.029	0.0 0.029	0.0 0.029	0.0 0.029	0.0 0.029
24	0.0 0.0	0.0 0.0	0.0 0.027	0.0 0.027	0.0 0.027	0.0 0.027	0.0 0.027	0.0 0.027
25	0.0 0.0	0.0 0.0	0.0 0.026	0.0 0.026	0.0 0.026	0.0 0.026	0.0 0.026	0.0 0.026
26	0.0 0.0	0.0 0.0	0.0 0.027	0.0 0.027	0.0 0.027	0.0 0.027	0.0 0.027	0.0 0.027
27	0.0 0.0	0.0 0.0	0.0 0.030	0.0 0.030	0.0 0.030	0.0 0.030	0.0 0.030	0.0 0.030
31	0.0 0.0	0.0 0.0	0.0 0.016	0.0 0.065	0.0 0.097	0.0 0.057	0.0 0.113	0.0 0.113
32	0.0 0.0	0.0 0.0	0.0 0.007	0.0 0.028	0.0 0.043	0.0 0.071	0.0 0.078	0.0 0.078
33	0.0 0.0	0.0 0.0	0.0 0.009	0.0 0.036	0.0 0.054	0.0 0.071	0.0 0.080	0.0 0.080
34	0.0 0.0	0.0 0.0	0.0 0.009	0.0 0.036	0.0 0.054	0.0 0.072	0.0 0.081	0.0 0.081
36	0.0 0.0	0.0 0.0	0.0 0.009	0.0 0.037	0.0 0.055	0.0 0.073	0.0 0.083	0.0 0.083

42	C:C 0:0	0:0 0:0	0:0 0:0	0:0 0:020	0:071 0:089	0:094 C:113	0:118 0:145	0:129 0:153
44	55:CCC 0:0	59:000 0:0	55:CCC 0:0	59:000 0:031	59:000 0:098	59:000 C:123	59:000 0:160	59:000 0:166
48	C:C 0:0	0:0 0:0	0:0 0:0	0:059 0:036	0:176 0:112	0:235 0:142	0:265 0:178	0:265 0:183
49	C:C 0:0	0:0 0:0	0:0 0:0	0:068 0:039	0:091 0:097	0:091 0:116	0:159 0:159	0:159 0:164
51	C:C 0:0	0:0 0:0	0:0 0:0	0:0 0:047	0:0 0:102	0:091 0:180	0:091 0:227	0:091 0:242
52	C:C 0:0	0:0 0:0	0:0 0:0	0:0 0:044	0:105 0:110	0:211 0:151	0:263 0:243	0:316 0:265
54	C:C 0:0	0:0 0:0	0:0 0:0	0:111 0:070	0:204 0:140	0:278 0:216	0:352 0:275	0:370 0:292
55	C:C 0:0	0:0 0:0	0:0 0:0	0:0 0:042	0:0 0:090	0:074 0:167	0:111 0:215	0:111 0:229
56	C:C 0:0	0:0 0:0	0:0 0:0	0:0 0:049	0:0 0:106	0:0 0:175	0:0 0:228	0:0 0:244
61	0:0 0:0	0:0 0:0	0:0 0:0	0:0 0:070	0:056 0:099	0:111 0:141	0:167 0:183	0:167 0:197
62	C:C 0:0	0:0 0:0	0:0 0:0	0:214 0:119	0:214 0:134	0:214 0:164	0:214 0:154	0:214 0:209
63	C:C 0:0	0:0 0:0	0:0 0:0	0:100 0:056	0:100 0:110	0:150 0:151	0:200 0:152	0:250 0:219
67	C:C 0:0	0:0 0:0	0:0 0:0	0:0 0:093	0:0 0:111	0:0 0:146	0:0 0:185	0:0 0:204
71	C:C 0:0	0:0 0:0	0:0 0:0	0:021 0:021	0:085 0:083	0:106 C:104	0:170 0:167	0:255 0:250
72	C:C 0:0	0:0 0:0	0:0 0:0	0:0 0:020	0:0 0:078	0:0 C:058	0:0 0:157	0:0 0:235



TABLE 3 FOR GRADUATION-YEAR 72

ACCUMULATED LOSSES

ROW 1 : ACC. LOSS RATE OF SSC-GRP INDICATED  
 ROW 2 : ACC. LOSS RATE (OTHER) OF SSC-GRP  
 ROW 3 : ACC. LOSS RATE (RETIRED) OF SSC-GRP  
 LAST 3 ROWS REFER TO TOTAL GRAD-YEAR GROUP

SSC	1	2	3	4	5	6	7	8
20	C.C C.C C.C	C.C 0.0 0.0	0.031 0.031 0.0	0.031 0.031 0.0	0.062 0.031 0.032	0.062 0.031 0.032	0.156 0.031 0.129	0.187 0.031 0.161
30	C.C C.C 0.0	0.0 0.0 0.0	0.028 0.039 0.019	0.103 0.037 0.068	0.131 0.056 0.079	0.135 0.079 0.041	0.206 0.086 0.133	0.232 0.086 0.174
40	C.C C.C 0.0	0.0 0.0 0.0	0.0 0.0 0.0	0.049 0.031 0.019	0.123 0.079 0.027	0.160 0.123 0.042	0.215 0.160 0.066	0.252 0.160 0.103
50	C.C C.C 0.0	0.0 0.0 0.0	0.0 0.0 C.C	0.048 0.031 0.018	0.145 0.111 0.038	0.256 0.188 0.064	0.342 0.239 0.135	0.368 0.256 0.149
60	C.C C.C 0.0	0.0 0.0 0.0	0.0 0.0 0.0	0.094 0.044 0.0	0.132 0.113 0.021	C.189 0.151 0.044	0.283 0.189 0.116	0.321 0.208 0.143
70	0.0 C.C C.C	0.0 0.0 0.0	C.C C.C C.C	0.081 0.020 0.042	0.143 0.087 0.067	0.184 C.102 0.091	0.286 0.163 0.146	0.408 0.245 0.216
80	C.C C.C 0.0	0.0 0.0 0.0	0.0 C.C C.C	0.032 0.032 0.0	0.063 0.063 0.0	0.075 0.063 0.017	0.159 0.063 0.102	0.206 0.063 0.133
90	C.C C.C C.C	0.0 0.0 0.0	C.C C.C 0.0	0.060 0.012 0.048	0.145 0.072 0.078	0.193 0.108 0.085	0.313 0.157 0.186	0.325 0.166 0.188
1100	0.0 C.C C.C	0.0 0.0 0.0	C.C C.C C.C	0.021 0.021 0.0	0.064 0.064 0.0	0.128 0.106 0.024	0.149 0.128 0.024	0.151 0.170 0.026
1300	C.C C.C C.C	0.0 0.0 0.0	0.0 C.C C.C	0.036 0.037 0.019	0.074 0.036 0.020	0.111 0.074 0.040	0.148 0.074 0.080	C.204 0.074 0.140
72	C.C C.C 0.0	0.0 0.0 0.0	0.005 0.003 0.003	0.061 0.036 0.026	0.117 0.081 0.040	0.165 0.113 0.060	0.237 0.142 0.111	0.279 0.156 0.143

TABLE 1 FCF GRADUATION-YEAR 73  
ACCUMULATED TOTAL LOSSES

ROW 1 : ACTUAL LCSS RATE  
ROW 2 : ESTIMATE

SSC	1	2	3	4	5	6	7
21	55.000 0.0	55.000 0.0	59.000 0.059	59.000 0.059	99.000 0.059	99.000 0.088	99.000 0.118
22	0.0 0.0	0.0 0.0	0.0 0.0	0.0 0.0	0.0 0.0	0.091 0.091	0.091 0.091
23	0.0 0.0	0.0 0.0	0.250 0.250	0.250 0.250	0.250 0.250	0.250 0.250	0.500 0.500
24	0.0 0.0	0.0 0.0	0.0 0.0	0.0 0.0	0.0 0.0	0.0 0.0	0.0 0.0
25	0.0 0.0	0.0 0.0	0.0 0.0	0.0 0.0	0.0 0.0	0.0 0.0	0.0 0.0
26	0.0 0.0	0.0 0.0	0.0 0.0	0.0 0.0	0.0 0.0	0.0 0.0	0.0 0.0
27	0.0 0.0	0.0 0.0	0.500 0.500	0.500 0.500	0.500 0.500	0.500 0.500	0.500 0.500
31	0.0 0.0	0.017 0.017	0.017 0.023	0.068 0.076	0.119 0.134	0.136 0.152	0.136 0.158
32	0.0 0.0	0.0 0.0	0.042 0.049	0.042 0.065	0.042 0.088	0.042 0.103	0.208 0.241
33	0.0 0.0	0.0 0.0	0.143 0.019	0.286 0.047	0.429 0.084	0.571 0.327	0.714 0.556
34	0.0 0.0	0.0 0.0	0.0 0.009	0.0 0.028	0.333 0.253	0.333 0.268	0.333 0.275
38	0.0 0.0	0.0 0.0	0.0 0.010	0.250 0.272	0.250 0.293	0.500 0.351	0.500 0.357





TABLE 2 FOR GRADUATION-YEAR 73  
 ACCUMULATED LCSS (EXCL. RETIREMENT)

ROW 1 : ACTUAL LCSS RATE  
 ROW 2 : ESTIMATE

SSC	1	2	3	4	5	6	7
21	99.000 0.0	99.000 0.0	99.000 0.0	99.000 0.0	99.000 0.0	99.000 0.0	99.000 0.0
22	0.0 0.0	0.0 0.0	0.0 0.0	0.0 0.0	0.0 0.0	0.0 0.0	0.0 0.0
23	0.0 0.0	0.0 0.0	0.0 0.0	0.0 0.0	0.0 0.0	0.0 0.0	0.0 0.0
24	0.0 0.0	0.0 0.0	0.0 0.0	0.0 0.0	0.0 0.0	0.0 0.0	0.0 0.0
25	0.0 0.0	0.0 0.0	0.0 0.0	0.0 0.0	0.0 0.0	0.0 0.0	0.0 0.0
26	0.0 0.0	0.0 0.0	0.0 0.0	0.0 0.0	0.0 0.0	0.0 0.0	0.0 0.0
27	0.0 0.0	0.0 0.0	0.0 0.0	0.0 0.0	0.0 0.0	0.0 0.0	0.0 0.0
31	0.0 0.0	0.0 0.0	0.0 0.006	0.017 0.025	0.034 0.050	0.051 0.066	0.051 0.075
32	0.0 0.0	0.0 0.0	0.0 0.008	0.0 0.024	0.0 0.048	0.0 0.065	0.0 0.081
33	0.0 0.0	0.0 0.0	0.143 0.015	0.286 0.047	0.429 0.084	0.425 0.103	0.429 0.112
34	0.0 0.0	0.0 0.0	0.0 0.009	0.0 0.028	0.167 0.066	0.167 0.085	0.167 0.094
38	0.0 0.0	0.0 0.0	0.0 0.010	0.0 0.029	0.0 0.058	0.250 0.087	0.250 0.096

42	C:C	0:0	0:0	0:0	0:0	0:042	0:063	0:125	0:146
	C:C	0:0	0:0	C:C	C:C	C:040	0:064	0:11E	0:135
44	55:000	99:000	99:000	99:000	99:000	99:000	99:000	99:000	99:000
	0:0	0:0	0:0	0:0	0:0	0:039	0:065	0:11C	0:129
4E	C:C	0:0	0:0	0:0	0:0	0:043	0:043	0:087	0:087
	C:C	0:0	0:0	0:0	0:0	0:039	0:062	0:1C7	0:124
45	C:C	0:0	0:0	0:0	0:0	0:028	0:083	0:083	0:111
	0:0	0:0	0:0	0:0	0:0	0:037	0:068	0:105	0:126
51	C:C	0:0	0:0	0:0	0:0	0:0	0:0	0:091	0:182
	0:0	0:0	0:0	0:0	0:0	0:037	0:112	0:224	0:308
52	C:C	0:0	0:0	0:0	0:0	0:0	0:267	0:400	0:400
	0:0	0:0	0:0	0:0	0:0	0:036	0:144	0:261	0:333
54	C:C	0:0	0:0	0:0	0:0	0:079	0:184	0:237	0:342
	0:0	0:0	0:0	0:0	0:0	0:052	0:142	0:239	0:328
55	C:C	0:0	0:0	0:0	0:0	0:038	0:038	0:269	0:385
	0:0	0:0	0:0	0:0	0:0	0:041	0:107	0:246	0:336
56	C:C	0:0	0:0	0:0	0:0	0:0	0:0	0:0	0:0
	0:0	0:0	0:0	0:0	0:0	0:039	0:118	0:225	0:304
61	C:C	0:0	0:0	0:0	0:0	0:0	0:0	0:0	0:0
	0:0	0:0	0:0	0:0	0:0	0:039	0:118	0:137	0:196
62	C:C	0:0	0:0	0:0	0:0	0:083	0:250	0:333	0:500
	0:0	0:0	0:0	0:0	0:0	0:056	0:167	0:204	0:256
63	C:C	0:0	0:0	0:0	0:0	0:053	0:158	0:158	0:211
	0:0	0:0	0:0	0:0	0:0	0:049	0:148	0:164	0:230
67	C:C	0:0	0:0	0:0	0:0	0:0	0:0	0:0	0:0
	0:0	0:0	0:0	0:0	0:0	0:045	0:136	0:155	0:227
71	C:0	0:0	0:0	0:0	0:0	0:037	0:093	0:185	0:255
	0:0	0:0	0:0	0:0	0:0	0:036	0:090	0:180	0:252
72	0:0	0:0	0:0	0:0	0:0	0:0	0:0	0:0	0:0
	0:0	0:0	0:0	0:0	0:0	0:033	0:083	0:167	0:233



TABLE 3 FOR GRADUATION-YEAR 73

ACCUMULATED LESSES

ROW 1 : ACCUMULATED RATE OF SSC-GRP INDICATED  
 ROW 2 : ACCUMULATED RATE (TOTAL) OF SSC-GRP  
 ROW 3 : ACCUMULATED RATE (NET) OF SSC-GRP  
 LAST 3 ROWS REFER TO TOTAL GRAD-YEAR GROUP

SSC	1	2	3	4	5	6	7
20	C.C 0.0 0.0	0.0 0.0 0.0	0.039 0.039 0.039	0.039 0.039 0.039	0.039 0.039 0.039	0.088 0.088 0.088	0.118 0.0 0.118
30	0.0 0.0 0.0	0.010 0.010 0.010	0.030 0.030 0.030	0.030 0.030 0.030	0.140 0.080 0.080	C.17C 0.080 0.080	0.23C 0.040 0.140
40	C.C 0.0 0.0	0.0 0.0 0.0	0.013 0.013 0.013	0.017 0.017 0.017	0.129 0.025 0.025	0.194 0.044 0.044	0.232 0.044 0.140
50	0.0 0.0 0.0	0.0 0.0 0.0	0.031 0.031 0.031	0.042 0.042 0.042	0.187 0.079 0.079	0.333 0.125 0.125	0.417 0.323 0.138
60	C.C 0.0 0.0	0.0 0.0 0.0	0.0 0.0 0.0	0.033 0.048 0.038	0.238 0.143 0.111	0.31C 0.167 0.171	0.429 0.238 0.238
70	C.C 0.0 0.0	0.0 0.0 0.0	C.033 0.033 0.033	0.038 0.038 0.038	0.158 0.078 0.078	0.242 0.118 0.118	0.321 0.242 0.140
80	0.0 0.0 0.0	0.0 0.0 0.0	0.034 0.034 0.034	0.111 0.044 0.039	0.194 0.134 0.063	0.278 0.144 0.063	0.366 0.339 0.169
90	C.C 0.0 0.0	0.0 0.0 0.0	0.013 0.013 0.013	0.044 0.014 0.014	0.218 0.167 0.082	0.282 0.172 0.111	0.323 0.172 0.179
1100	C.C 0.0 0.0	0.0 0.0 0.0	C.033 0.033 0.033	0.038 0.038 0.038	0.083 0.083 0.083	0.083 0.083 0.083	0.083 0.083 0.083
1300	0.0 0.0 0.0	0.0 0.0 0.0	0.0 0.0 0.0	0.0 0.0 0.0	0.057 0.057 0.057	C.037 0.037 0.037	0.114 0.057 0.057
73	0.0 0.0 0.0	0.001 0.001 0.001	0.024 0.017 0.017	0.072 0.037 0.036	0.152 0.043 0.068	0.220 0.138 0.068	0.273 0.147 0.129

TABLE 1 FOR GRADUATION-YEAR 74  
 ACCUMULATED TOTAL LCSESSE

ROW 1 : ACTUAL LCSS RATE  
 ROW 2 : ESTIMATE

SSC	1	2	3	4	5	6
21	C:C	0:0	0:0	0:0	0:0	0:0
	C:C	0:0	0:0	0:0	0:0	0:0
22	0:C	0:0	0:0	0:100	0:200	0:200
	0:C	0:0	0:C	0:040	0:147	0:147
23	C:C	0:0	0:0	0:0	0:0	0:0
	C:C	0:0	0:C	0:023	0:023	0:023
24	C:C	0:0	0:0	0:0	0:0	0:067
	C:C	0:0	0:C	0:018	0:018	0:084
25	0:C	0:0	0:C	0:0	1:000	1:000
	C:C	0:0	0:C	0:024	1:000	1:000
26	0:0	0:0	0:0	0:0	0:0	0:0
	C:C	0:0	0:C	0:021	0:021	0:021
27	C:C	0:0	0:0	0:0	0:0	0:0
	C:C	0:0	0:C	0:024	0:024	0:024
31	C:C	0:0	0:0	0:054	0:143	0:175
	C:C	0:006	0:C13	0:062	0:144	0:180
32	C:C	0:030	0:C30	0:030	0:182	0:272
	C:C	0:015	0:C22	0:045	0:194	0:284
33	C:C	0:0	0:0	0:500	0:500	0:500
	C:C	0:010	0:C15	0:058	0:117	0:117
34	C:0	0:0	0:0	0:333	0:333	0:333
	C:C	0:010	0:C19	0:365	0:404	0:404
38	C:C	0:0	0:286	0:429	0:429	0:571
	C:C	0:009	0:150	0:370	0:407	0:556

42	C:C	0:0	0:0	0:0	0:074	0:167	0:222
	C:C	0:0	C:C	0:0	0:068	0:137	0:194
44	S:C:C	S:0:000	S:0:000	S:0:000	S:0:000	S:0:000	S:0:000
	C:C	0:0	C:C	0:056	0:121	0:178	0:178
48	C:C	0:0	0:0	0:049	0:138	0:196	0:196
	C:C	0:0	0:0	0:036	0:089	0:107	0:107
49	C:C	0:0	0:0	0:044	0:089	0:160	0:160
	C:C	0:0	0:0	0:111	0:222	0:333	0:333
51	C:C	0:009	C:C19	0:038	0:142	0:245	0:245
	C:C	0:0	0:0	0:026	0:056	0:278	0:278
52	C:C	0:009	C:C17	0:026	0:122	0:285	0:285
	C:C	0:0	0:0	0:107	0:286	0:425	0:425
54	C:C	0:008	0:036	0:129	0:280	0:421	0:421
	C:C	0:0	0:0	0:065	0:290	0:419	0:419
55	C:C	0:016	0:031	0:039	0:251	0:386	0:386
	C:C	0:0	0:0	0:0	0:091	0:091	0:091
56	C:C	0:005	0:019	0:028	0:200	0:285	0:285
	C:C	0:0	0:0	0:0	0:0	0:0	0:0
61	C:C	0:0	0:0	0:019	0:115	0:192	0:192
	C:C	0:0	0:0	0:0	0:250	0:333	0:333
62	C:C	0:0	0:0	0:017	0:155	0:241	0:241
	C:C	0:038	0:038	0:077	0:269	0:423	0:423
63	C:C	0:038	0:038	0:067	0:277	0:417	0:417
	C:C	0:0	0:0	0:0	0:500	0:500	0:500
67	C:C	0:0	0:0	0:021	0:563	0:604	0:604
	C:C	0:045	C:068	0:136	0:273	0:341	0:341
71	C:C	0:043	0:066	0:130	0:264	0:332	0:332
	C:C	0:0	0:0	0:0	0:200	0:200	0:200
72	C:C	0:037	C:037	0:053	0:319	0:333	0:333



TABLE 2 FCR GRADUATION-YEAR 74  
 ACCUMULATED LCSES (EXCL. RETIREMENT)

ROW 1 : ACTUAL LOSS RATE  
 ROW 2 : ESTIMATE

SSC	1	2	3	4	5	6
21	C:C C:C	0:0 0:0	0:0 0:0	0:0 0:023	0:0 0:023	0:0 0:023
22	C:C C:C	0:0 0:0	0:0 0:0	C:100 0:040	0:100 0:040	0:100 0:040
23	C:C C:C	0:0 0:0	0:0 0:0	C:C 0:023	0:0 0:023	0:0 0:023
24	C:C C:C	0:0 0:0	C:C 0:0	C:C 0:018	0:0 0:018	C:0 0:018
25	C:C 0:0	0:0 0:0	0:0 0:0	0:0 0:024	0:0 0:024	0:0 0:024
26	C:0 C:0	0:0 0:0	0:0 C:0	0:0 0:021	0:0 0:021	0:0 0:021
27	C:C 0:0	0:0 0:0	C:C 0:0	0:0 0:024	0:0 0:024	0:0 0:024
31	0:C 0:0	0:0 0:006	0:C 0:013	0:C36 0:045	0:107 0:108	0:107 0:108
32	C:C C:C	C:030 0:015	0:C30 C:022	0:030 0:045	0:091 0:104	0:091 0:104
33	C:0 0:0	0:0 0:010	0:C 0:019	0:500 0:058	0:500 0:117	C:500 0:117
34	C:C 0:0	C:C 0:010	0:C 0:019	0:C 0:048	0:0 0:106	0:0 0:106
38	C:C 0:0	0:0 0:009	C:143 0:028	0:143 0:056	0:143 0:111	0:143 0:111

42	C-C C-C	0-C 0-C	0-C 0-C	0-C 0-C	0.056 0.050	0.148 0.118	0.185 0.155
44	SS-CCC C-C	SS-000 0-C	SS-CCC 0-C	SS-000 0.047	SS-000 0.103	SS-000 0.140	SS-000 0.140
48	0-C 0-C	0-C 0-C	0-C 0-C	0-C 0.045	0.080 0.098	0.120 0.136	
49	C-C 0-C	0-C 0-C	0-C 0-C	0.036 0.044	0.036 0.089	0.071 0.126	
51	0-C 0-C	0-C 0.009	0-C 0.019	0.111 0.038	0.222 0.142	0.333 0.245	
52	C-C C-C	0-C 0.009	C-C 0.017	0-C 0.026	0.056 0.122	0.222 0.235	
54	0-C C-C	0-C 0.008	0-C 0.016	0-C 0.024	0.143 0.136	0.250 0.240	
55	C-C 0-C	0-C 0.016	C-C 0.031	0.065 0.039	0.194 0.148	0.250 0.250	
56	C-C 0-C	0-C 0.009	0-C 0.019	0-C 0.028	0-C 0.120	0-C 0.213	
61	0-C C-C	0-C 0-C	C-C 0-C	0-C 0.019	0-C 0.115	0-C 0.192	
62	C-C C-C	0-C 0-C	C-C 0-C	0-C 0.017	0.250 0.155	0.333 0.241	
63	C-C 0-C	0-C 0-C	0-C 0-C	0.038 0.028	0.115 0.125	0.231 0.222	
67	C-C 0-C	0-C 0-C	C-C 0-C	0-C 0.021	0-C 0.125	0-C 0.208	
71	0-C 0-C	0-C 0.043	0-C 0.045	0.114 0.108	0.182 0.172	0.205 0.194	
72	C-C 0-C	0-C 0.037	C-C 0.037	0-C 0.093	0-C 0.148	0-C 0.167	



TABLE 3 FOR GRADUATION-YEAR 74

ACCUMULATED LOSSES

ROW 1 : ACC. LOSS RATE OF SSC-GRP INDICATED  
 ROW 2 : ACC. LOSS RATE (OTHER) OF SSC-GRP  
 ROW 3 : ACC. LOSS RATE (REF) OF SSC-GRP  
 LAST 3 ROWS REFER TO TOTAL GRAD-YEAR GROUP

SSC	1	2	3	4	5	6
20	0.0 0.0 0.0	0.0 0.0 0.0	0.0 0.0 0.0	0.023 0.023 0.0	0.073 0.023 0.051	0.100 0.023 0.077
30	0.0 0.0 0.0	0.010 0.010 0.0	0.030 0.020 0.010	0.089 0.030 0.042	0.188 0.109 0.089	0.248 0.175 0.136
40	0.0 0.0 0.0	0.0 0.0 0.0	0.0 0.0 0.0	0.056 0.047 0.010	0.121 0.103 0.021	0.178 0.146 0.043
50	0.0 0.0 0.0	0.010 0.010 0.0	0.031 0.021 0.011	0.062 0.031 0.032	0.216 0.134 0.095	0.351 0.237 0.145
60	0.0 0.0 0.0	0.022 0.0 0.022	0.022 0.0 0.022	0.043 0.022 0.022	0.239 0.130 0.123	0.348 0.217 0.167
70	0.0 0.0 0.0	0.041 0.041 0.0	0.061 0.041 0.021	0.122 0.103 0.023	0.265 0.192 0.122	0.327 0.192 0.172
80	0.0 0.0 0.0	0.024 0.0 0.024	0.024 0.0 0.024	0.024 0.0 0.024	0.098 0.049 0.051	0.144 0.073 0.075
90	0.0 0.0 0.0	0.040 0.013 0.027	0.033 0.027 0.027	0.107 0.033 0.056	0.160 0.093 0.113	0.293 0.107 0.209
1100	0.0 0.0 0.0	0.018 0.018 0.0	0.033 0.033 0.0	0.105 0.105 0.0	0.175 0.175 0.0	0.222 0.222 0.0
1300	0.0 0.0 0.0	0.0 0.0 0.0	0.0 0.0 0.0	0.028 0.028 0.0	0.056 0.056 0.0	0.111 0.111 0.0
74	0.0 0.0 0.0	0.015 0.009 0.006	0.026 0.013 0.011	0.071 0.048 0.024	0.166 0.105 0.069	0.243 0.149 0.112

TABLE 1 FCR GRADUATION-YEAR 75  
ACCUMULATED ACTUAL LOSSES

ROW 1 : ACTUAL LOSS RATE  
ROW 2 : ESTIMATE

SSC	1	2	3	4	5
21	55.CCC 0.0	55.000 0.0	55.CCC 0.C77	55.000 0.154	55.000 0.308
22	C.C 0.0	0.0 0.0	0.C 0.C59	0.C 0.059	0.0 0.118
23	C.C 0.0	0.0 0.0	0.C 0.C71	0.C 0.071	0.0 0.143
24	C.C 0.0	0.0 0.0	0.143 0.100	C.2E6 0.250	0.571 0.520
25	55.CCC 0.0	55.000 0.0	55.CCC 0.C77	55.000 0.154	55.000 0.308
26	C.C 0.0	0.0 0.0	0.C 0.071	0.C 0.071	0.0 0.143
27	55.CCC 0.0	55.000 0.0	55.CCC 0.C77	55.000 0.154	55.000 0.308
31	C.C14 0.014	C.042 0.049	C.C42 0.C49	0.C57 0.104	0.181 0.181
32	C.C 0.0	C.040 0.029	0.C4C 0.C29	C.C80 0.090	0.120 0.143
33	C.C 0.0	C.500 0.034	C.500 0.034	1.000 1.000	1.000 1.000
34	C.C 0.0	0.0 0.025	C.C C.C25	C.250 0.288	0.250 0.301
3E	C.C 0.0	0.0 0.024	C.C 0.024	0.C51 0.056	0.182 0.165

42	C.C18	C.C35	C.C7C	0.123	0.228
	C.C11	0.028	0.050	0.102	0.193
44	0.C	0.C	C.C	0.0	0.0
	C.CC7	0.0C7	0.020	0.033	0.073
48	0.C	0.0	0.C33	0.067	0.100
	0.0C6	0.006	0.051	0.070	0.114
49	0.C	0.C	C.C	0.0	0.083
	0.0C6	0.006	0.019	0.032	0.118
51	0.C	0.C	0.C	0.083	0.083
	0.C	0.010	0.010	0.061	0.091
52	0.C	0.C	C.C	0.059	0.118
	0.0	0.010	0.010	0.058	0.096
54	0.C	0.C61	0.121	0.182	0.303
	0.C	0.047	0.109	0.180	0.286
55	0.C	0.C56	0.C56	0.111	0.222
	0.0	0.065	0.065	0.113	0.247
56	0.C	0.C	0.C	0.143	0.143
	0.0	0.011	0.011	0.188	0.216
61	0.C	0.C59	0.C55	0.176	0.235
	0.C	0.073	0.087	0.214	0.305
62	0.C	0.0	C.1CC	C.200	0.200
	0.C	0.017	0.051	0.085	0.136
63	0.C	0.045	C.C45	0.045	0.227
	0.C	0.028	0.042	0.056	0.189
67	55.CCC	55.000	59.CCC	55.C00	59.000
	0.C	0.041	0.061	0.122	0.224
71	0.C	0.0	0.C26	0.077	0.077
	0.C	0.012	0.037	0.086	0.086
72	0.C	0.200	C.2CC	C.200	0.200
	0.C	0.041	0.041	0.082	0.082

81	C:C	0:0	0:111	0:111	0:111	0:222
	C:C	0:0	0:132	0:132	0:132	0:206
82	0:0	0:0	0:040	0:040	0:040	0:120
	C:0	0:0	0:034	0:034	0:034	0:125
51	C:C	0:0	0:0	0:0	0:0	0:091
	C:0	0:0	0:0	0:013	0:013	0:128
95	C:C	0:0	0:029	0:059	0:059	0:235
	C:C	0:0	0:029	0:052	0:052	0:216
1101	0:0	0:0	0:0	0:089	0:089	0:133
	C:C	0:0	0:0	0:085	0:085	0:138
1102	0:0	0:0	0:0	0:0	0:0	0:500
	0:0	0:0	0:0	0:057	0:057	0:113
1103	0:0	0:0	0:0	0:0	0:0	0:091
	C:0	0:0	0:0	0:055	0:055	0:091
1301	C:0	0:0	0:0	0:0	0:0	0:077
	C:0	0:019	0:019	0:019	0:019	0:058
1302	C:0	0:0	0:0	0:0	0:0	0:0
	C:C	0:024	0:024	0:024	0:024	0:049
1304	C:C	0:0	0:0	0:0	0:0	0:0
	C:C	0:023	0:023	0:023	0:023	0:045
1305	C:0	0:0	0:0	0:0	0:0	0:0
	C:C	0:025	0:025	0:025	0:025	0:050
1306	C:0	0:071	0:071	0:071	0:071	0:071
	C:0	0:038	0:038	0:038	0:038	0:057
1307	C:0	0:0	0:0	0:0	0:0	0:0
	C:C	0:024	0:024	0:024	0:024	0:048
130E	C:0	0:0	0:0	0:0	0:0	0:0
	C:C	0:025	0:025	0:025	0:025	0:050

TABLE 2 FCR GRADUATION-YEAR 75  
 ACCUMULATED LCSSSES (EXCL. RETIREMENT)

ROW 1 : ACTUAL LCSS RATE  
 ROW 2 : ESTIMATE

SSC	1	2	3	4	5
21	55.CCC C.C	99.000 0.0	99.000 0.C77	99.000 0.C77	99.000 0.154
22	C.C	0.0	0.0	0.0	0.0
23	C.C	0.0	0.C59	0.C59	0.118
24	C.C	0.0	0.0	0.0	0.0
25	C.C	0.0	0.C71	0.C71	0.143
26	C.C	0.0	0.143	0.143	0.286
27	C.C	0.0	0.100	0.100	0.200
31	55.CCC C.C	99.000 0.0	99.000 0.C77	99.000 0.C77	99.000 0.154
32	C.C	0.0	0.0	0.0	0.0
33	C.C	0.0	0.C71	0.C71	0.143
34	55.CCC C.C	99.000 0.0	99.000 0.C77	99.000 0.C77	99.000 0.154
38	C.C	0.0	0.014	0.014	0.069
	C.C	0.022	0.022	0.042	0.070
	C.C	0.040	0.040	0.040	0.040
	C.C	0.029	0.C29	0.C50	0.065
	C.C	0.500	0.500	0.500	0.500
	C.C	0.034	0.C34	0.C60	0.078
	C.C	0.0	0.0	0.0	0.0
	C.C	0.025	0.C25	0.C51	0.068
	C.C	0.0	0.0	0.091	0.091
	C.C	0.024	0.C24	0.C56	0.072

42	C:C18 0:C11	0:018 0:011	0:C53 C:C32	C:070 C:C48	0:140 0:101
44	C:C C:C07	0:0 C:C07	0:0 C:C20	0:0 0:C33	0:0 0:073
48	0:C C:C06	0:0 0:006	0:0 C:C19	0:033 0:037	0:067 0:081
49	C:C C:C06	0:0 C:006	0:0 C:C15	0:0 0:C32	0:042 0:077
51	C:C C:C	0:0 C:010	0:0 C:C10	0:083 0:C61	0:083 0:091
52	C:C C:C	0:0 C:010	0:0 C:C10	0:059 C:C58	0:118 0:096
54	C:C C:C	0:030 C:C17	C:030 0:C17	0:061 0:C58	0:121 0:100
55	0:C C:C	0:0 0:C10	0:0 C:C10	0:056 C:C57	0:056 0:086
56	C:C C:C	0:0 0:011	0:0 0:C11	0:0 C:C53	0:0 0:085
61	C:C C:C	0:0 C:C15	0:0 C:C30	0:0 0:C45	0:0 0:091
62	C:C C:C	0:0 C:017	0:100 0:C51	0:200 C:C85	0:200 0:136
63	C:C C:C	0:045 C:028	0:045 C:C42	0:045 0:056	0:182 0:141
67	S:C00 C:C	S:000 0:C20	S:C00 0:C41	S:000 C:C61	S:000 0:122
71	C:C C:C	0:0 C:C12	C:0 C:C12	0:051 C:C60	0:051 0:060
72	C:C 0:C	0:200 0:041	C:200 C:C41	0:200 C:C82	0:200 0:082

81	C:C C:C	C:0 C:0	0:C 0:C23	0:0 0:C23	0:0 0:C23	0:111 0:093
82	C:C C:C	0:0 0:0	0:C4C 0:C34	0:040 0:C34	0:080 0:085	
81	0:C 0:C	0:0 0:0	0:C 0:C	0:0 0:C13	0:091 0:128	
85	C:C 0:C	0:0 0:0	0:C 0:C	0:029 0:022	0:176 0:156	
11C1	C:C C:C	0:0 0:0	0:C 0:C	0:067 0:C63	0:089 0:094	
11C2	C:C C:C	0:0 0:C	C:C C:C	0:0 0:057	0:500 0:113	
11C3	0:C C:0	0:0 0:0	0:C 0:C	0:0 0:055	0:0 0:091	
1301	C:C C:C	0:0 0:C19	0:C 0:C19	0:0 0:019	0:077 0:058	
13C2	C:C C:C	0:0 0:024	0:C 0:C24	0:0 0:024	0:0 0:049	
13C4	C:C C:0	0:0 0:C23	C:C C:C23	0:0 0:C23	0:0 0:045	
13C5	C:C C:0	0:0 0:025	C:C C:C25	0:0 0:025	0:0 0:050	
13C6	C:C C:C	0:071 0:038	C:C71 C:C38	C:C71 0:038	0:071 0:057	
13C7	C:C C:0	0:0 0:024	C:C 0:C24	0:0 0:C24	0:0 0:048	
13C8	C:C C:C	0:0 0:025	C:C 0:C25	0:0 0:C25	0:0 0:050	

TABLE 3 FCR GRADUATION-YEAR 75

ACCLMULATED LCSSSES

ROW 1 : ACC-LCSS RATE (CF SSC-GRP INDICATED)  
 ROW 2 : ACC-LCSS RATE (CIVIL) CF SSC-GRP  
 ROW 3 : ACC-LCSS RATE (RETIRED) CF SSC-GRP  
 LAST 3 ROWS REFER TO TOTAL GRAD-YEAR GROUP

SSC	1	2	3	4	5
20	C-C C-C C-C	0-0 0-0 0-0	0-C77 0-C77 0-0	0-154 0-177 0-063	0-308 0-154 0-182
30	0-C49 C-C C-C	0-C46 0-026 0-018	0-C44 0-026 0-C18	0-114 0-053 0-065	0-184 0-070 0-123
40	C-C C-C 0-0	0-012 0-008 0-008	0-C38 0-023 0-014	0-069 0-038 0-032	0-137 0-084 0-098
50	C-C C-C 0-0	0-C34 0-011 0-023	0-C37 0-011 0-C47	0-126 0-027 0-071	0-207 0-072 0-127
60	0-0 C-C C-C	0-041 0-020 0-021	0-C61 0-041 C-C21	0-122 0-061 C-063	0-224 0-122 0-116
70	0-0 C-C 0-0	0-023 0-023 0-023	0-C45 0-C23 0-023	0-091 0-058 0-024	0-091 0-068 0-024
80	C-C C-C 0-0	0-0 0-0 0-0	0-C39 0-029 0-030	0-139 0-029 0-030	0-147 0-068 0-065
90	0-0 C-C C-C	0-0 0-0 0-0	0-C18 0-018 0-018	0-036 0-018 0-018	0-178 0-143 0-042
1100	C-C C-C 0-0	0-0 0-0 0-0	0-C 0-C 0-0	0-078 0-C39 0-021	0-137 0-098 0-043
1300	C-C C-C 0-0	0-026 0-026 0-0	0-C26 0-026 0-C	0-026 0-026 0-0	0-031 0-031 0-0
75	0-003 C-C C-C	0-023 0-013 0-010	0-040 0-C21 0-C30	0-C27 0-047 0-042	0-162 0-061 0-078

**APPENDIX E  
SCATTERPLOTS OF LOSS RATES VERSUS THEIR ESTIMATES**

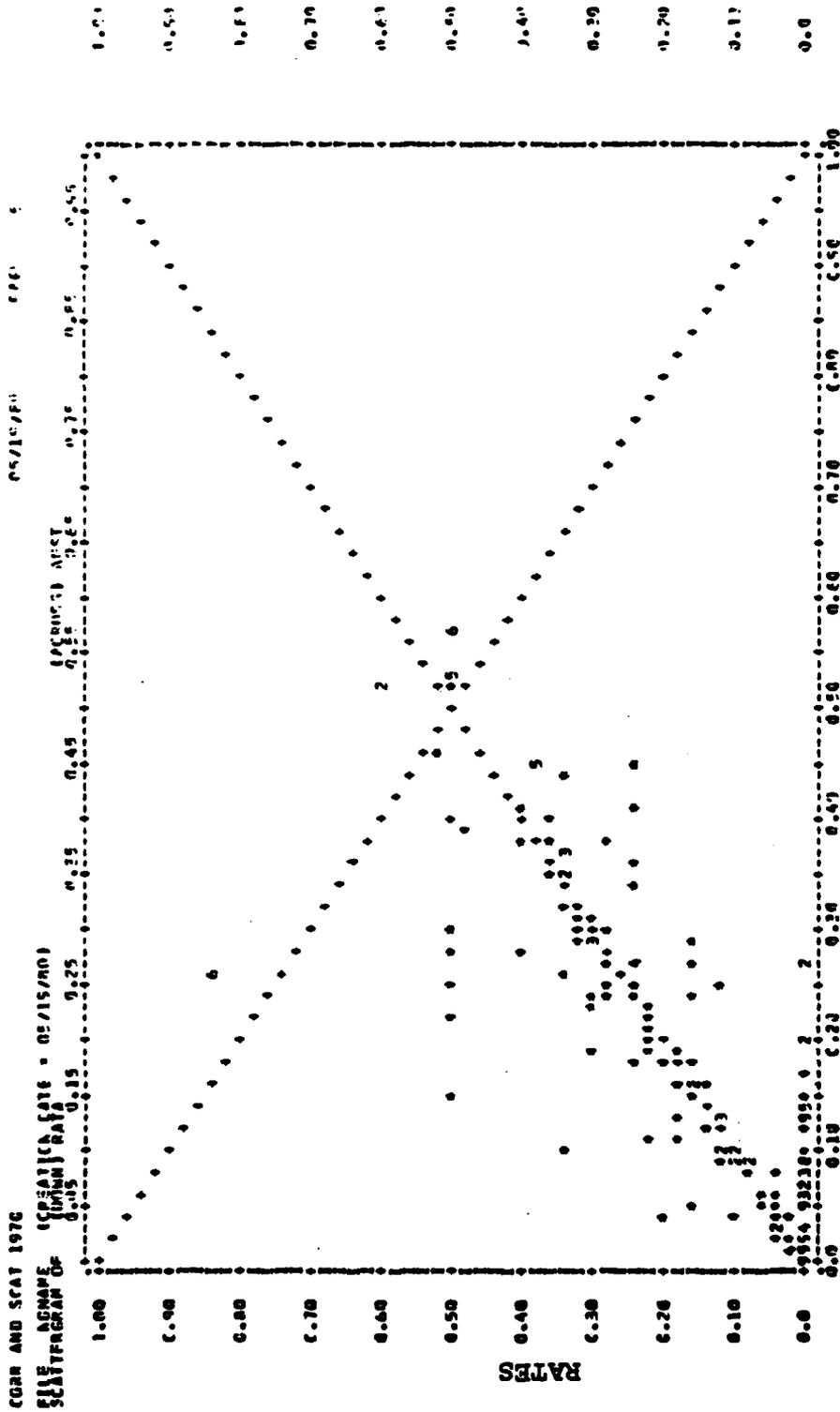


Figure E-I: Graduation-year group 1970:  
Scatterplot of accumulated rates ( $r_{70}^j$ ) vs. their  
estimates for all years  $k$  after graduation and for  
all SSC  $j$ .

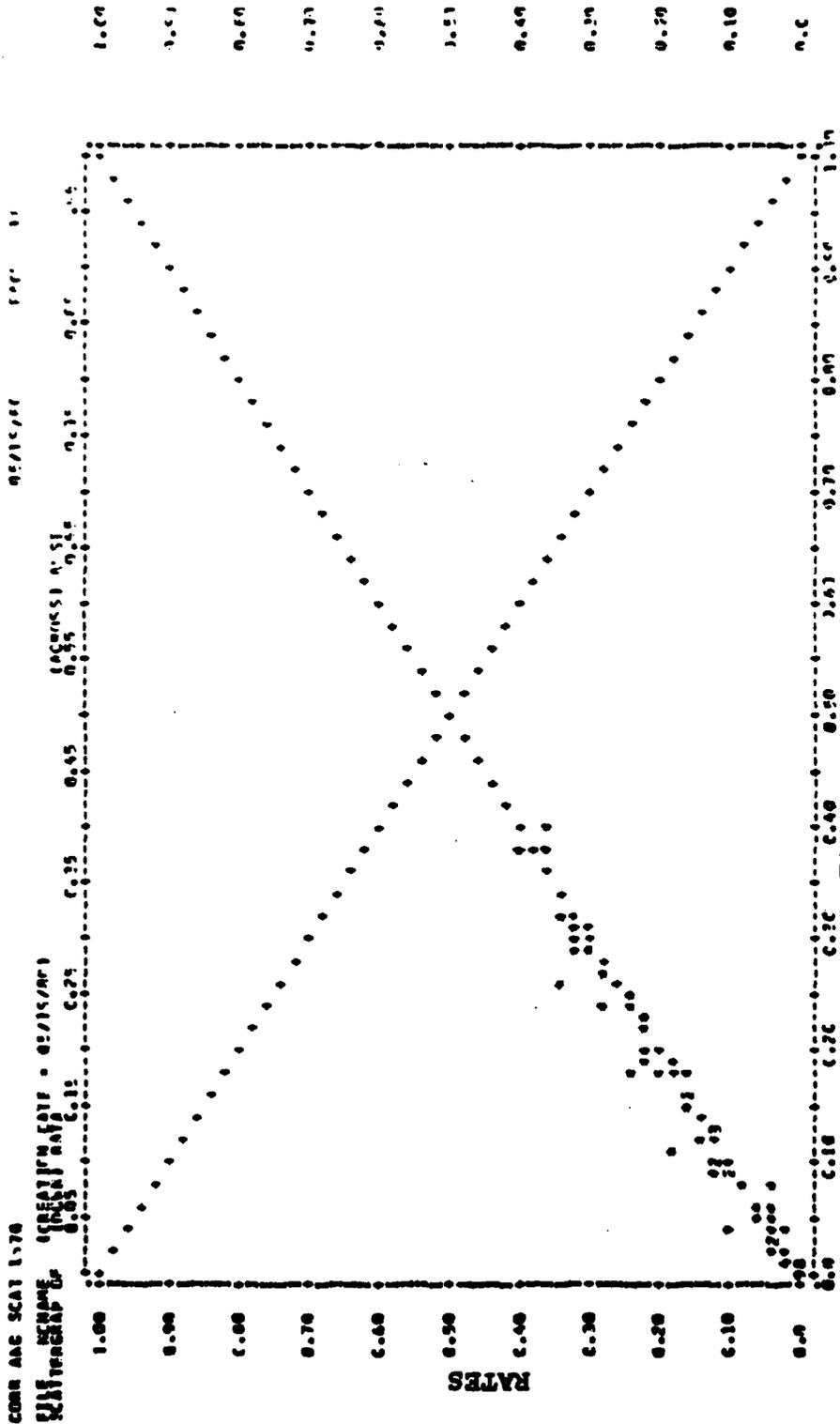


Figure E-II: Graduation-year group 1970:  
 Scatterplot of accumulated loss rates ( $r_{70j}$ ) vs. their  
 estimates ( $b_{70jk}$ ) for all years  $k$  after graduation and  
 SSC  $j$  with sample sizes  $> 20$ .

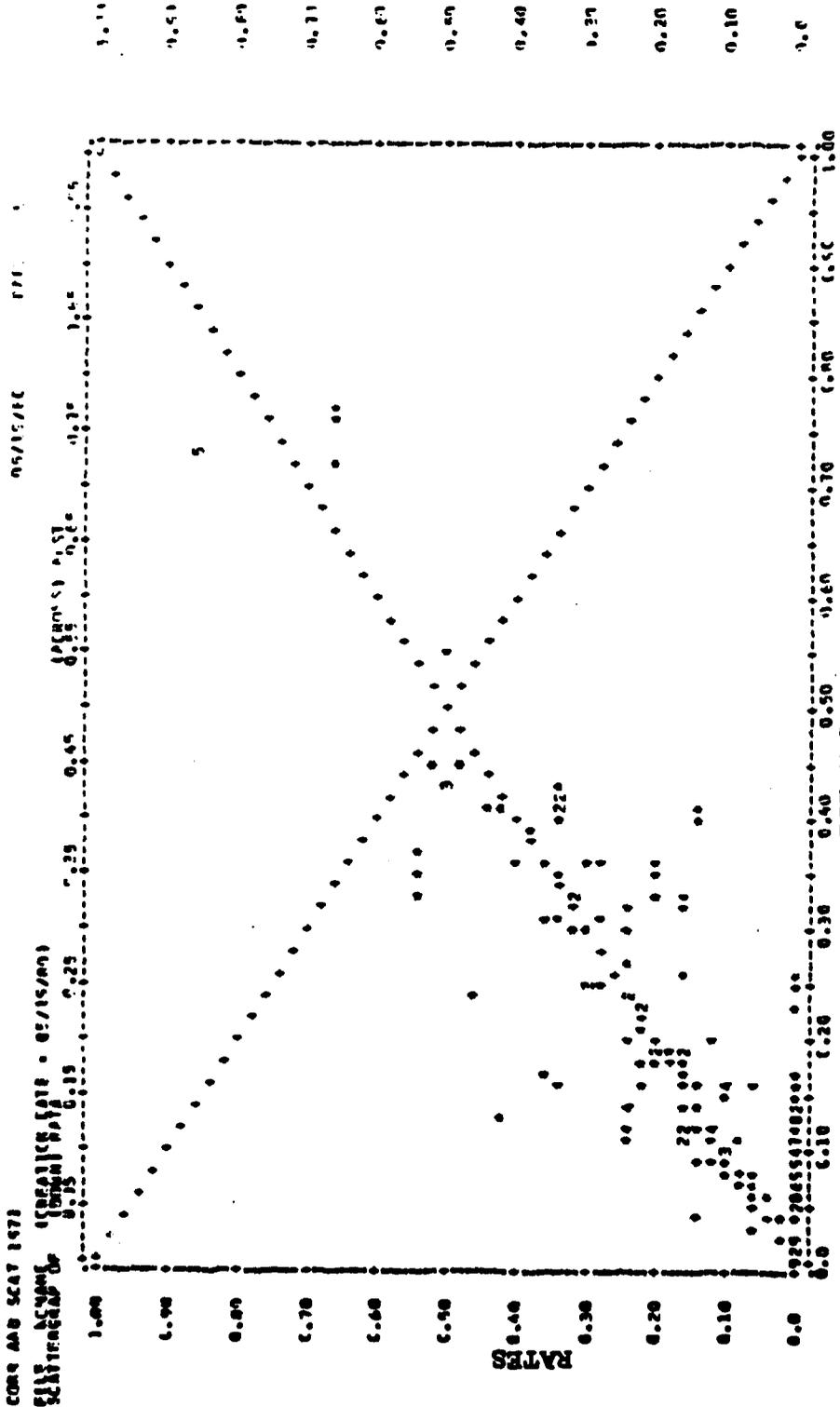


Figure E-III: Graduation-year group 1971: Scatterplot of accumulated loss rates ( $r_{71}^k$ ) vs. their estimates ( $b_{71}^k$ ) for all years  $k$  after graduation and for all  $j$ .

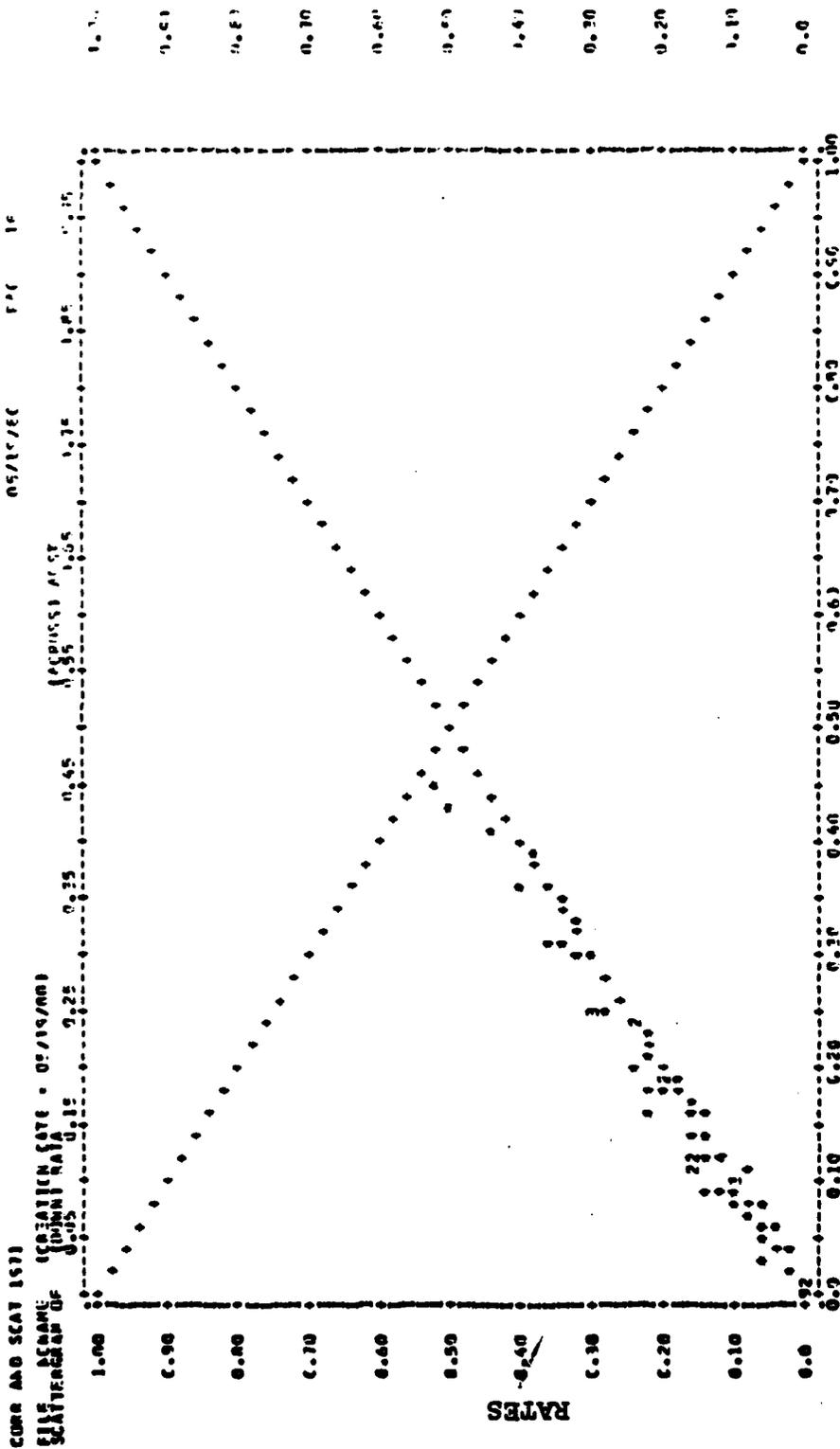


Figure E-IV: Graduation-year group 1971: Scatterplot of accumulated loss rates ( $r_{71jk}$ ) vs. their estimates ( $b_{71jk}$ ) for all years  $k$  after graduation and SSC  $j$  with sample sizes  $> 20$ .

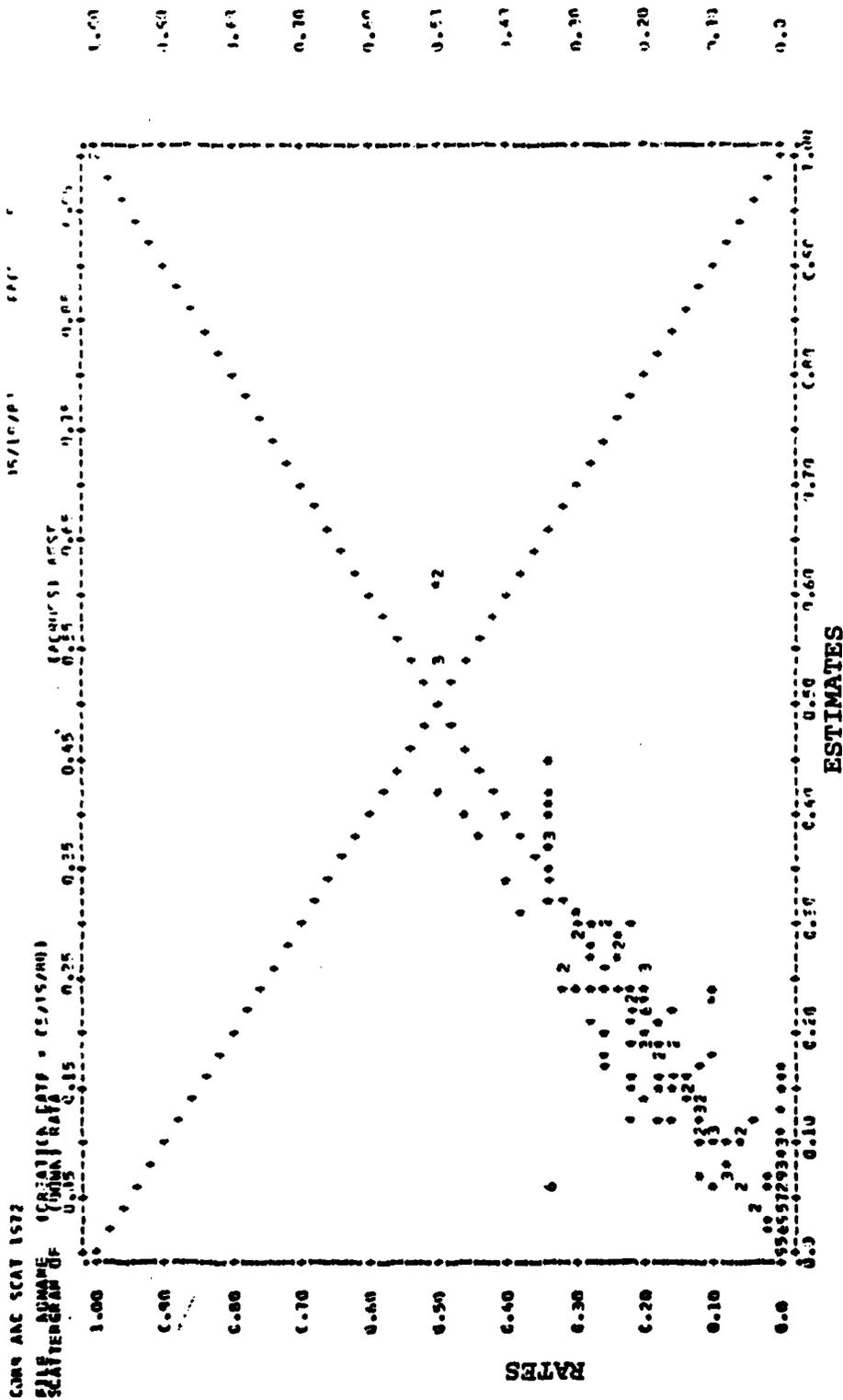


Figure E-V: Graduation-year group 1972: Scatterplot of accumulated loss rates ( $r_{72,j,k}$ ) vs. their estimates ( $b_{72,j,k}$ ) for all years  $k$  after graduation and for all SSC.

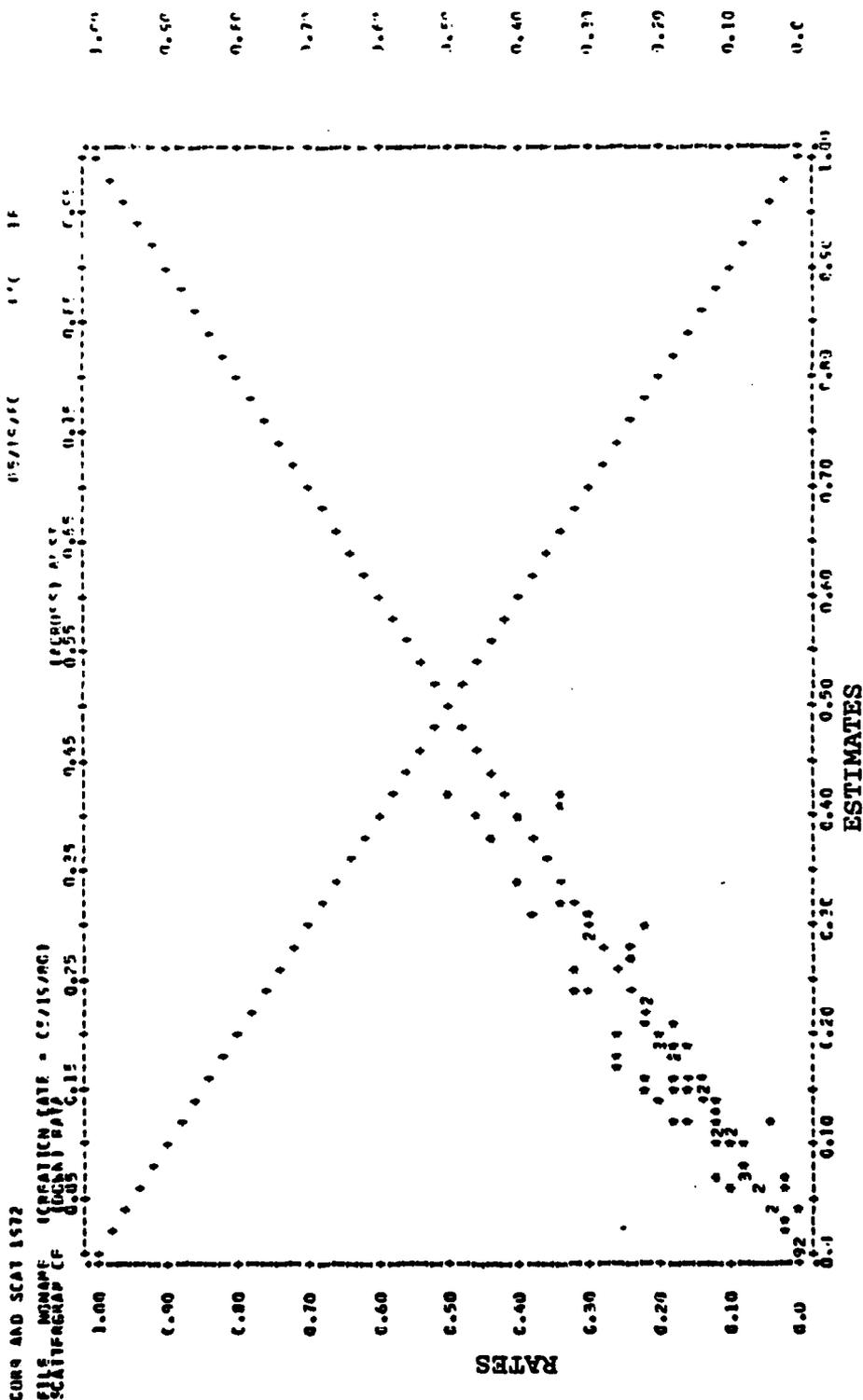


Figure E-VI: Graduation-year group 1972:  
 Scatterplot of accumulated loss rates ( $r_{72}^k$ ) vs. their  
 estimates ( $b_{72}^k$ ) for all years  $k$  after graduation and  
 SSC  $j$  with sample sizes  $> 20$ .

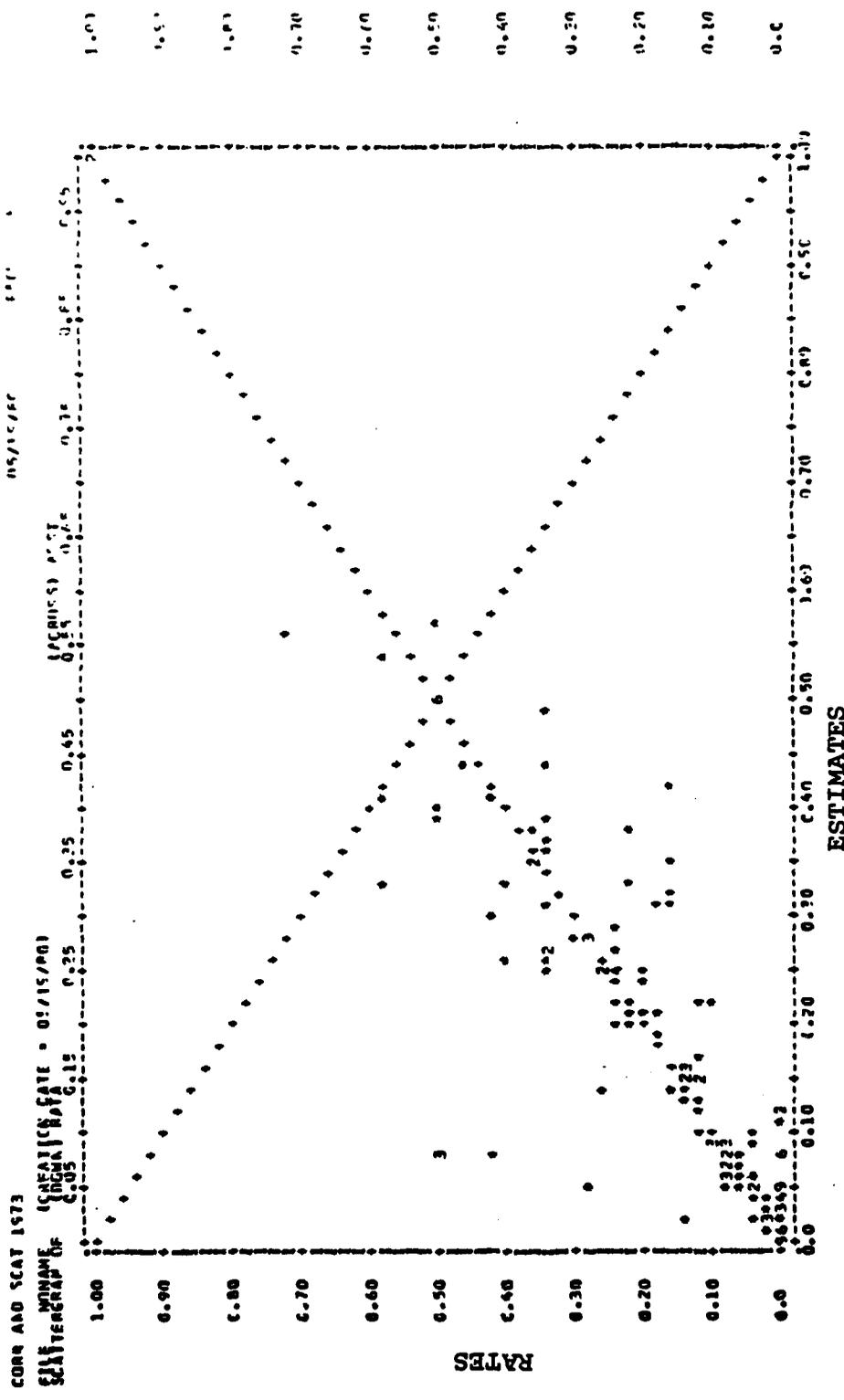


Figure E-VII: Graduation-year group 1973  
 Scatterplot of accumulated loss rates ( $r_{3,j,k}$ ) vs. their  
 estimates ( $b_{3,j,k}$ ) for all years  $k$  after graduation and  
 for all SSC  $j$ .

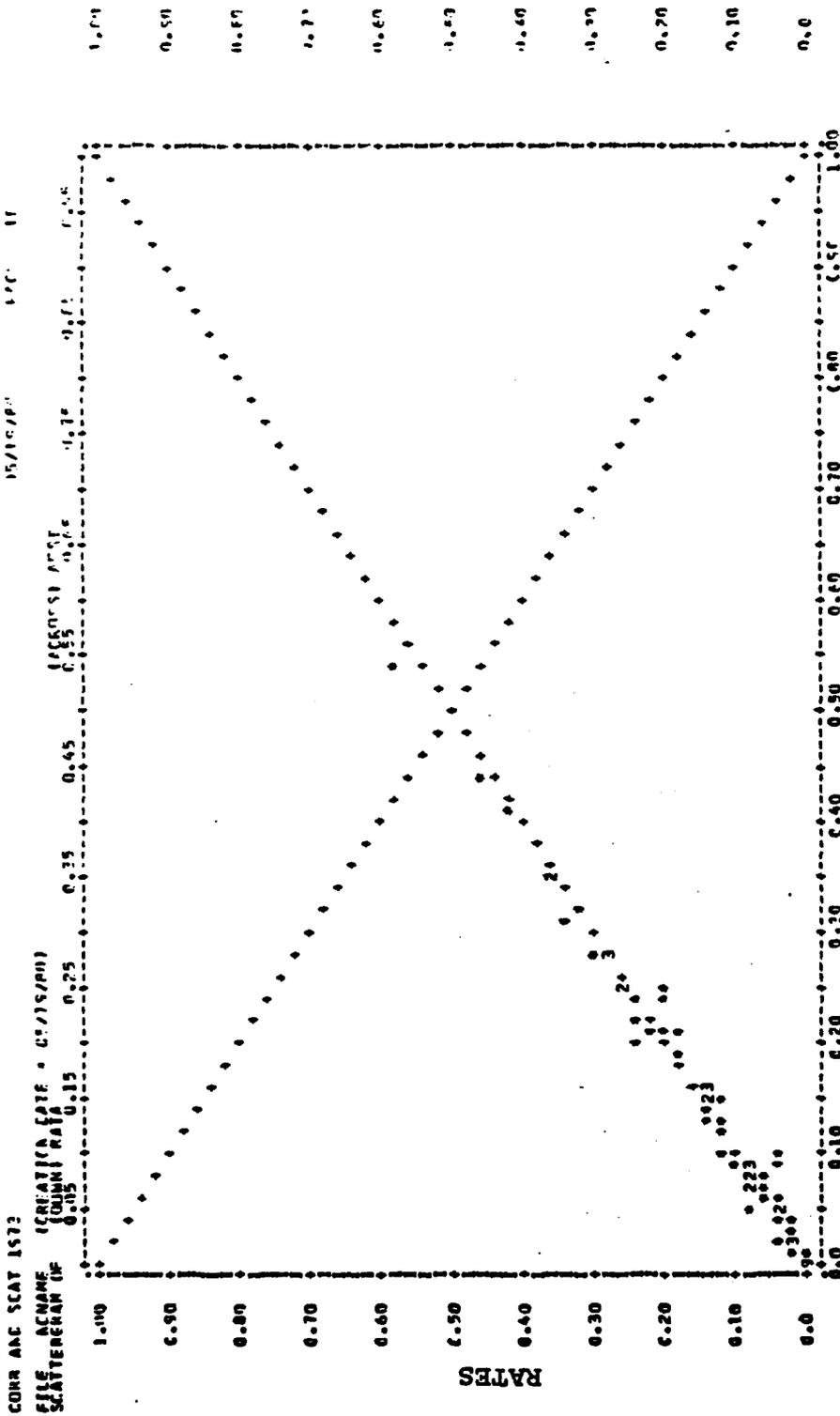


Figure E-VIII: Graduation-year group 1973:  
 Scatterplot of accumulated loss rates ( $r_{73, k}$ ) vs. their  
 estimates ( $b_{73, j, k}$ ) for all years  $k$  after graduation and  
 SSC  $j$  with sample sizes  $> 20$ .

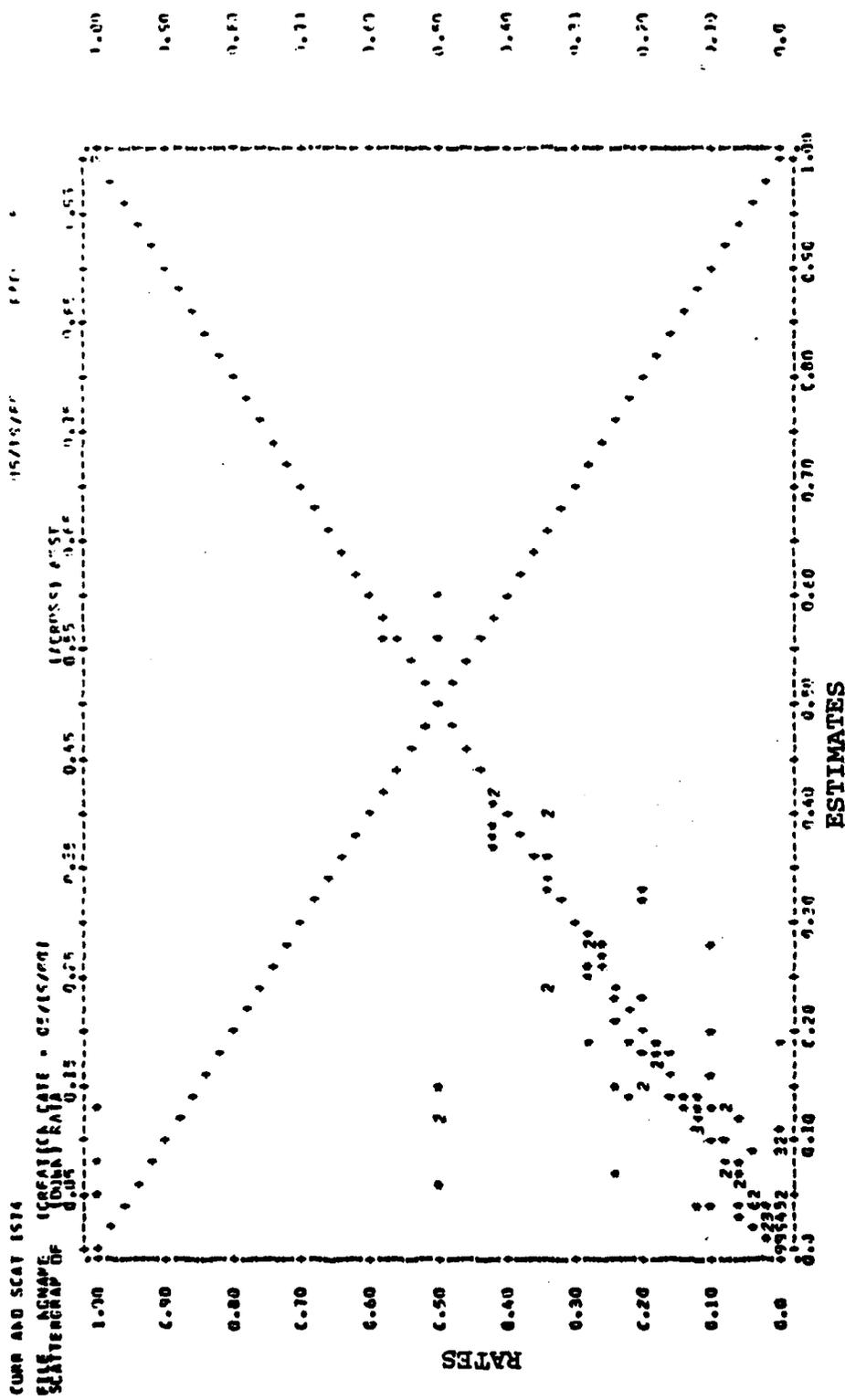


Figure E-IX: Graduation-year group 1974:  
 Scatterplot of accumulated loss rates ( $r_{74,j,k}$ ) vs. their  
 estimates ( $b_{74,j,k}$ ) for all years  $k$  after graduation and  
 for all SSC  $j$ .

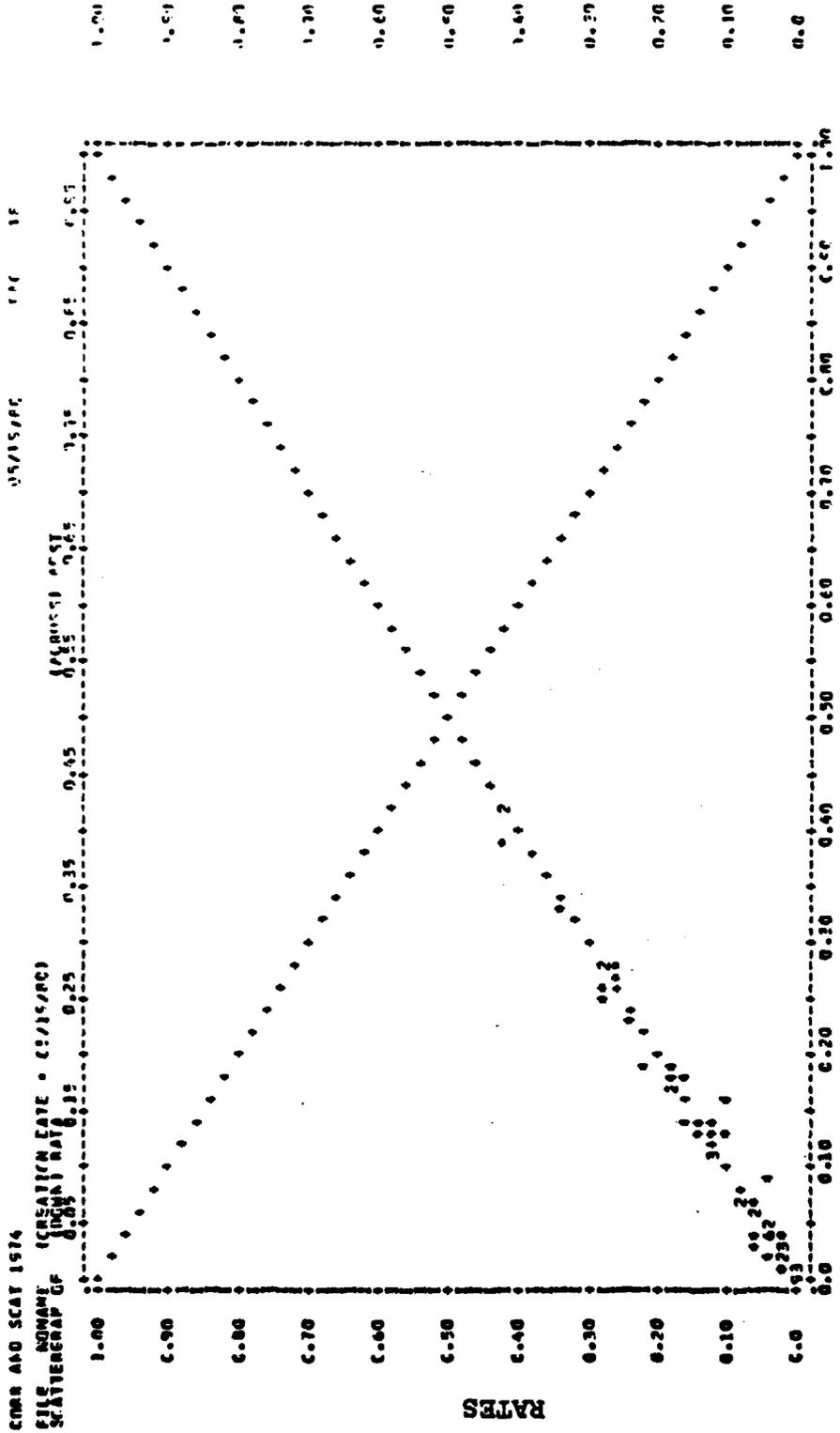


Figure E-X: Graduation-year group 1974:  
 Scatterplot of accumulated loss rates ( $r_j^k$ ) vs. their  
 estimates ( $b7A_j^k$ ) for all years  $k$  after graduation and  
 SSC  $j$  with sample sizes  $> 20$ .

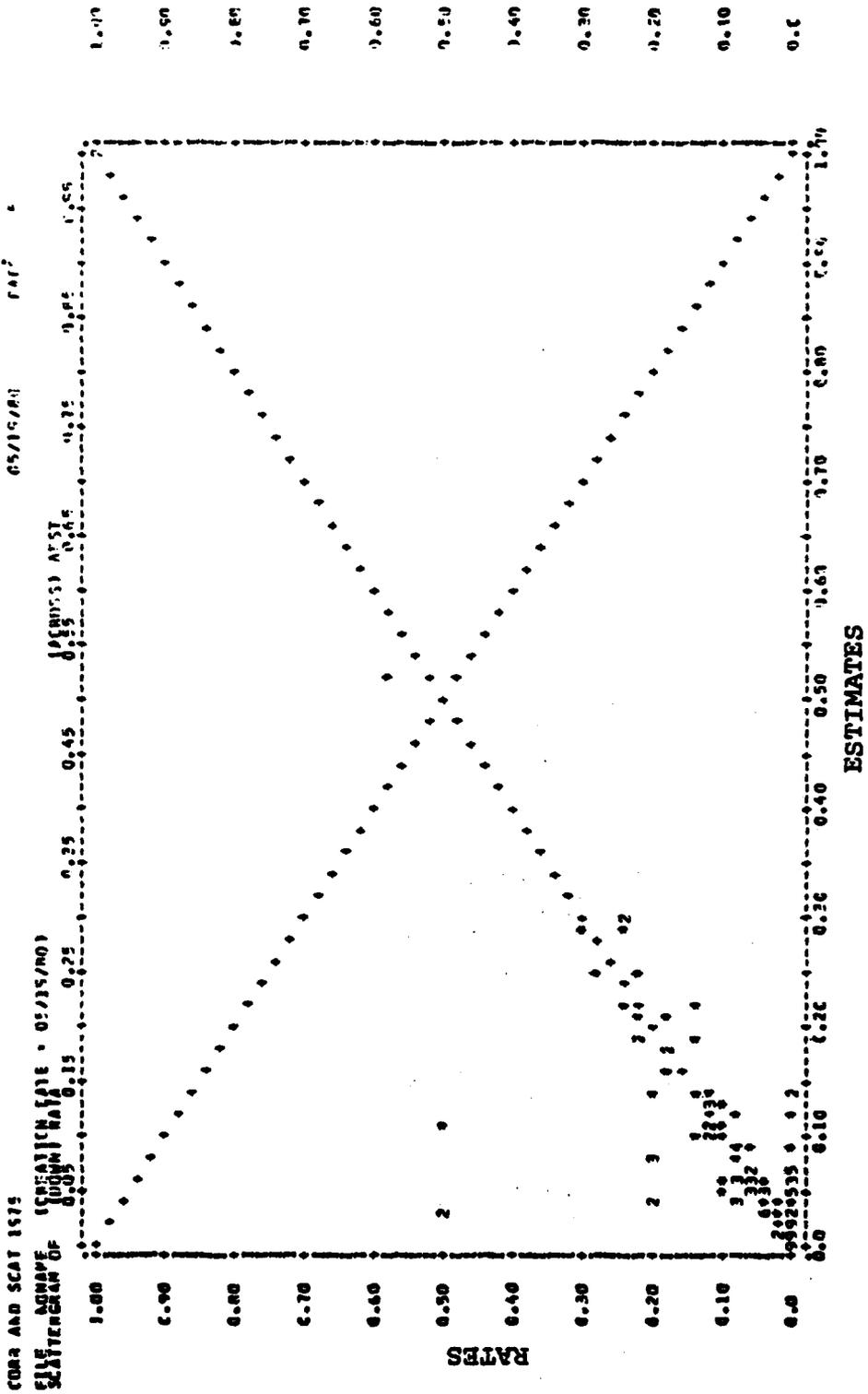


Figure E-XI: Graduation-year group 1975:  
 Scatterplot of accumulated loss rates ( $r_{75,j,k}$ ) vs. their  
 estimates ( $b_{75,j,k}$ ) for all SSC  $j,k$   
 for all SSC  $j,k$ .

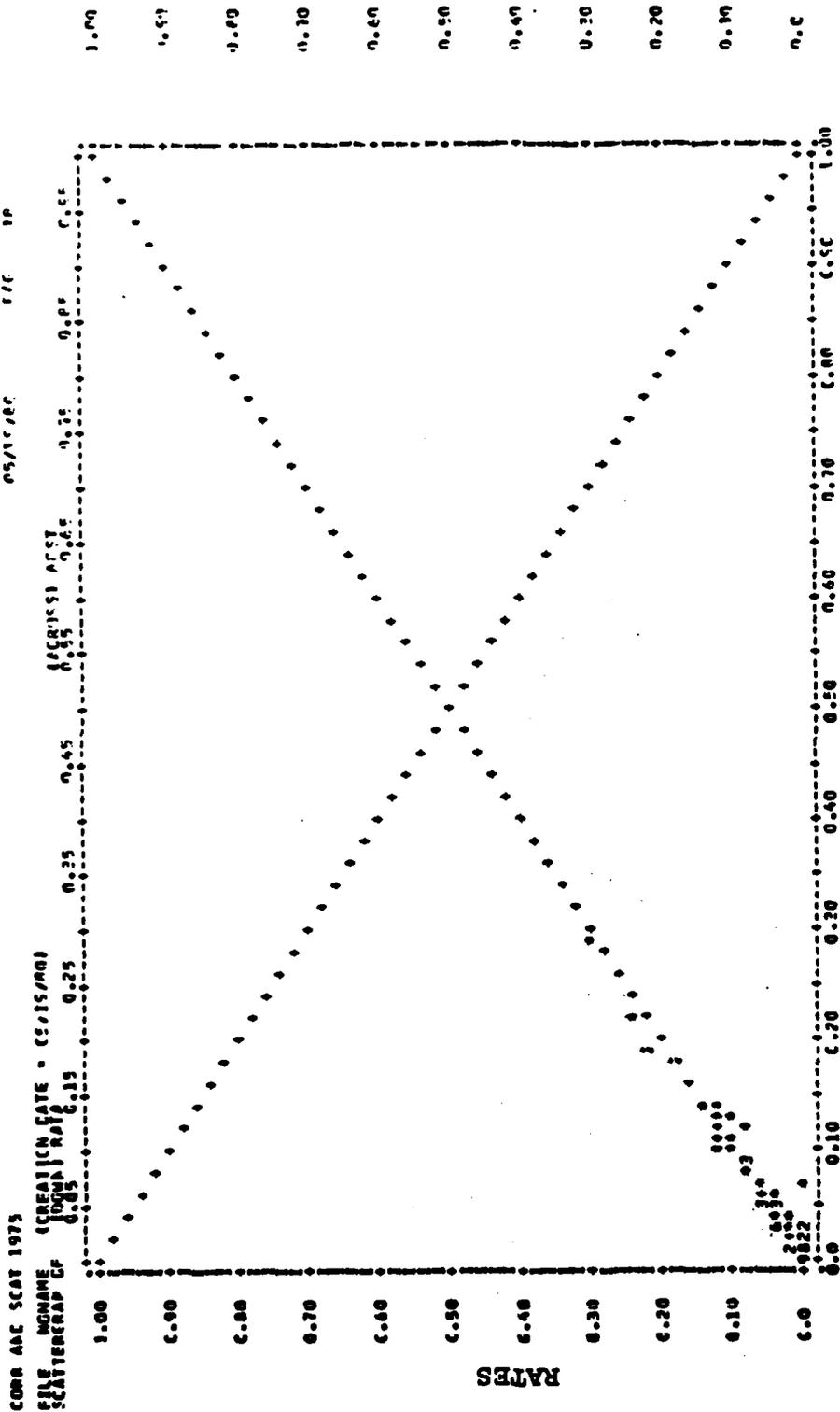


Figure E-XII: Graduation-year group 1975:  
 Scatterplot of accumulated loss rates ( $r_{75}^k$ ) vs. their  
 estimates ( $b_{75}^k$ ) for all years  $k$  after graduation and  
 SSC  $j$  with sample sizes  $> 20$ .

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